

**A large genome-wide association study of age-related macular degeneration highlights contributions of rare and common variants.**

Lars G. Fritsche<sup>1†</sup>, Wilmar Igl<sup>2†</sup>, Jessica N. Cooke Bailey<sup>3†</sup>, Felix Grassmann<sup>4†</sup>, Sebanti Sengupta<sup>1†</sup>, Jennifer L. Bragg-Gresham<sup>1,5</sup>, Kathryn P. Burdon<sup>6</sup>, Scott J. Hebring<sup>7</sup>, Cindy Wen<sup>8</sup>, Mathias Gorski<sup>2</sup>, Ivana K. Kim<sup>9</sup>, David Cho<sup>10</sup>, Donald Zack<sup>11,12,13,14,15</sup>, Eric Souied<sup>16</sup>, Hendrik P. N. Scholl<sup>11,17</sup>, Elisa Bala<sup>18</sup>, Kristine E. Lee<sup>19</sup>, David J. Hunter<sup>20,21</sup>, Rebecca J. Sardell<sup>22</sup>, Paul Mitchell<sup>23</sup>, Joanna E. Merriam<sup>24</sup>, Valentina Cipriani<sup>25,26</sup>, Joshua D. Hoffman<sup>27</sup>, Tina Schick<sup>28</sup>, Yara T. E. Lechanteur<sup>29</sup>, Robyn H. Guymer<sup>30</sup>, Matthew P. Johnson<sup>31</sup>, Yingda Jiang<sup>32</sup>, Chloe M. Stanton<sup>33</sup>, Gabriëlle H. S. Buitendijk<sup>34,35</sup>, Xiaowei Zhan<sup>1,36,37</sup>, Alan M. Kwong<sup>1</sup>, Alexis Boleda<sup>38</sup>, Matthew Brooks<sup>39</sup>, Linn Gieser<sup>38</sup>, Rinki Ratnapriya<sup>38</sup>, Kari E. Branham<sup>39</sup>, Johanna R. Foerster<sup>1</sup>, John R. Heckenlively<sup>39</sup>, Mohammad I. Othman<sup>39</sup>, Brendan J. Vote<sup>6</sup>, Helena Hai Liang<sup>30</sup>, Emmanuelle Souzeau<sup>40</sup>, Ian L. McAllister<sup>41</sup>, Timothy Isaacs<sup>41</sup>, Janette Hall<sup>40</sup>, Stewart Lake<sup>40</sup>, David A. Mackey<sup>6,30,41</sup>, Ian J. Constable<sup>41</sup>, Jamie E. Craig<sup>40</sup>, Terrie E. Kitchner<sup>7</sup>, Zhenglin Yang<sup>42,43</sup>, Zhiguang Su<sup>44</sup>, Hongrong Luo<sup>8,44</sup>, Daniel Chen<sup>8</sup>, Hong Ouyang<sup>8</sup>, Ken Flagg<sup>8</sup>, Danni Lin<sup>8</sup>, Guanping Mao<sup>8</sup>, Henry Ferreyra<sup>8</sup>, Klaus Stark<sup>2</sup>, Claudia N. von Strachwitz<sup>45</sup>, Armin Wolf<sup>46</sup>, Caroline Brandl<sup>2,4,47</sup>, Guenther Rudolph<sup>46</sup>, Matthias Olden<sup>2</sup>, Margaux A. Morrison<sup>48</sup>, Denise J. Morgan<sup>48</sup>, Matthew Schu<sup>49,50,51,52,53</sup>, Jeeyun Ahn<sup>54</sup>, Giuliana Silvestri<sup>55</sup>, Evangelia E. Tsironi<sup>56</sup>, Kyu Hyung Park<sup>57</sup>, Lindsay A. Farrer<sup>49,50,51,52,53</sup>, Anton Orlin<sup>58</sup>, Alexander Brucker<sup>59</sup>, Mingyao Li<sup>60</sup>, Christine Curcio<sup>61</sup>, Saddek Mohand-Saïd<sup>62,63,64,65</sup>, José-Alain Sahel<sup>62,63,64,65,66,67,68</sup>, Isabelle Audo<sup>62,63,64,69</sup>, Mustapha Benchaboune<sup>65</sup>, Angela J. Cree<sup>70</sup>, Christina A. Rennie<sup>71</sup>, Srinivas V. Goverdhan<sup>70</sup>, Michelle Grunin<sup>72</sup>, Shira Hagbi-Levi<sup>72</sup>, Peter Campochiaro<sup>11,13</sup>, Nicholas Katsanis<sup>73,74,75</sup>, Frank G. Holz<sup>17</sup>, Frédéric Blond<sup>62,63,64</sup>, Hélène Blanché<sup>76</sup>, Jean-François Deleuze<sup>76,77</sup>, Robert P. Igo Jr.<sup>3</sup>, Barbara Truitt<sup>3</sup>, Neal S. Peachey<sup>18,78</sup>, Stacy M. Meuer<sup>19</sup>, Chelsea E. Myers<sup>19</sup>, Emily L. Moore<sup>19</sup>, Ronald Klein<sup>19</sup>, Michael A. Hauser<sup>79,80,81</sup>, Eric A. Postel<sup>79</sup>, Monique D. Courtenay<sup>22</sup>, Stephen G. Schwartz<sup>82</sup>, Jaclyn L. Kovach<sup>82</sup>, William K. Scott<sup>22</sup>, Gerald Liew<sup>23</sup>, Ava G. Tfan<sup>23</sup>, Bamini Gopinath<sup>23</sup>, John C. Merriam<sup>24</sup>, R. Theodore Smith<sup>24,84</sup>, Jane C. Khan<sup>41,83,85</sup>, Humma Shahid<sup>85,86</sup>, Anthony T. Moore<sup>25,26,87</sup>, J. Allie McGrath<sup>27</sup>, René Laux<sup>3</sup>, Milam A. Brantley Jr.<sup>88</sup>, Anita Agarwal<sup>88</sup>, Lebriz Ersoy<sup>28</sup>, Albert Caramoy<sup>28</sup>, Thomas Langmann<sup>28</sup>, Nicole T. M. Saksens<sup>29</sup>, Eiko K. de Jong<sup>29</sup>, Carel B. Hoyng<sup>29</sup>, Melinda S. Cain<sup>30</sup>, Andrea J. Richardson<sup>30</sup>, Tammy M. Martin<sup>89</sup>, John Blangero<sup>31</sup>, Daniel E. Weeks<sup>32,90</sup>, Bal Dhillon<sup>91</sup>, Cornelia M. van Duijn<sup>35</sup>, Kimberly F. Doheny<sup>92</sup>, Jane Romm<sup>92</sup>, Caroline C. W. Klaver<sup>34,35</sup>, Caroline Hayward<sup>33</sup>, Michael B. Gorin<sup>93,94</sup>, Michael L. Klein<sup>89</sup>, Paul N. Baird<sup>30</sup>, Anneke I. den Hollander<sup>29,95</sup>, Sascha Fauser<sup>28</sup>, John R. W. Yates<sup>25,26,85</sup>, Rando Allikmets<sup>24,96</sup>, Jie Jin Wang<sup>23</sup>, Debra A. Schaumberg<sup>20,97,98</sup>, Barbara E. K. Klein<sup>19</sup>, Stephanie A. Hagstrom<sup>78</sup>, Itay Chowers<sup>72</sup>, Andrew J. Lotery<sup>70</sup>, Thierry Lèveillard<sup>62,63,64</sup>, Kang Zhang<sup>8,44</sup>, Murray H. Brilliant<sup>7</sup>, Alex W. Hewitt<sup>6,30,41</sup>, Anand Swaroop<sup>38</sup>,

Emily Y. Chew<sup>99</sup>, Margaret A. Pericak-Vance<sup>22‡</sup>, Margaret DeAngelis<sup>48‡</sup>, Dwight Stambolian<sup>10‡</sup>, Jonathan L. Haines<sup>3,100‡</sup>, Sudha K. Iyengar<sup>3‡\*</sup>, Bernhard H. F. Weber<sup>4‡</sup>, Gonçalo R. Abecasis<sup>1‡\*</sup>, Iris M. Heid<sup>2‡\*</sup>

<sup>1</sup>Center for Statistical Genetics, Department of Biostatistics, University of Michigan, Ann Arbor, MI 48109, USA

<sup>2</sup>Department of Genetic Epidemiology, University of Regensburg, Germany

<sup>3</sup>Department of Epidemiology and Biostatistics, Case Western Reserve University School of Medicine, 2103 Cornell Rd Cleveland, OH 44106

<sup>4</sup>Institute of Human Genetics, University of Regensburg, Germany

<sup>5</sup>Kidney Epidemiology and Cost Center, Department of Biostatistics, Department of Internal Medicine - Nephrology, University of Michigan, Ann Arbor, MI 48109, USA

<sup>6</sup>School of Medicine, Menzies Research Institute Tasmania, University of Tasmania, Hobart, Tasmania, Australia

<sup>7</sup>Center for Human Genetics, Marshfield Clinic Research Foundation, 1000 N. Oak Ave. Marshfield, WI 54449, USA

<sup>8</sup>Department of Ophthalmology, University of California San Diego and VA San Diego Health System, La Jolla, California 92093, USA

<sup>9</sup>Retina Service, Massachusetts Eye and Ear, Department of Ophthalmology Harvard Medical School, Boston, MA, USA

<sup>10</sup>Department of Ophthalmology, Perelman School of Medicine, University of Pennsylvania

<sup>11</sup>Department of Ophthalmology, Wilmer Eye Institute - Johns Hopkins University School of Medicine - 400 North Broadway, Smith Building - Baltimore, Maryland 21287, USA

<sup>12</sup>Department of Molecular Biology and Genetics - Johns Hopkins University School of Medicine - 400 North Broadway, Smith Building - Baltimore, Maryland 21287, USA

<sup>13</sup>Department of Neuroscience - Johns Hopkins University School of Medicine - 400 North Broadway, Smith Building - Baltimore, Maryland 21287, USA

<sup>14</sup>Institute of Genetic Medicine - Johns Hopkins University School of Medicine - 400 North Broadway, Smith Building - Baltimore, Maryland 21287, USA

<sup>15</sup>Institut de la Vision, Université Pierre et Marie Curie, 17 rue Moreau, Paris, France

<sup>16</sup>Hôpital Intercommunal de Créteil, Hôpital Henri Mondor - Université Paris Est Créteil, France

<sup>17</sup>University of Bonn - Department of Ophthalmology - Ernst-Abbe-Str. 2 - D-53127 Bonn – Germany

<sup>18</sup>Louis Stokes Cleveland VA Medical Center, 10701 East Boulevard, Cleveland OH 44106, USA

<sup>19</sup>Department of Ophthalmology and Visual Sciences, University of Wisconsin, Madison, WI

- <sup>20</sup>Department of Epidemiology, Harvard School of Public Health
- <sup>21</sup>Department of Nutrition, Harvard School of Public Health
- <sup>22</sup>John P. Hussman Institute for Human Genomics, Miller School of Medicine, University of Miami, Miami, Florida, United States
- <sup>23</sup>Centre for Vision Research, Department of Ophthalmology and Westmead Millennium Institute for Medical Research, University of Sydney, Sydney, Australia
- <sup>24</sup>Department of Ophthalmology Columbia University, New York, NY
- <sup>25</sup>UCL Institute of Ophthalmology, University College London, London, EC1V 9EL UK
- <sup>26</sup>Moorfields Eye Hospital, London, EC1V 2PD UK
- <sup>27</sup>Center for Human Genetics Research, Vanderbilt University Medical Center, Nashville, Tennessee, United States
- <sup>28</sup>University Hospital of Cologne, Department of Ophthalmology, Kerpener Str. 62, 50924 Cologne, Germany
- <sup>29</sup>Department of Ophthalmology, Radboud University Medical Centre, Nijmegen, the Netherlands
- <sup>30</sup>Centre for Eye Research Australia, University of Melbourne, Royal Victorian Eye and Ear Hospital, East Melbourne, Victoria, 3000, Australia
- <sup>31</sup>South Texas Diabetes and Obesity Institute, University of Texas Rio Grande Valley School of Medicine, Brownsville, TX 78520
- <sup>32</sup>Department of Biostatistics, Graduate School of Public Health, University of Pittsburgh, Pittsburgh, PA 15261, USA
- <sup>33</sup>MRC Human Genetics Unit, Institute of Genetics and Molecular Medicine, University of Edinburgh, EH4 2XU Scotland, UK
- <sup>34</sup>Department of Ophthalmology, Erasmus Medical Center, Rotterdam, the Netherlands PO box 2040, 3000CA
- <sup>35</sup>Department of Epidemiology, Erasmus Medical Center, Rotterdam, the Netherlands PO box 2040, 3000CA
- <sup>36</sup>Quantitative Biomedical Research Center, Department of Clinical Science, University of Texas Southwestern Medical Center, 5323 Harry Hines Boulevard, Dallas, TX 75390, USA
- <sup>37</sup>Center for the Genetics of Host Defense, University of Texas Southwestern Medical Center, 5323 Harry Hines Boulevard, Dallas, TX 75390-8505, USA
- <sup>38</sup>Neurobiology Neurodegeneration & Repair Laboratory (N-NRL), National Eye Institute, National Institutes of Health, Bethesda, MD 20892, USA
- <sup>39</sup>Department of Ophthalmology and Visual Sciences, University of Michigan, Kellogg Eye Center, Ann Arbor, MI 48105, USA
- <sup>40</sup>Department of Ophthalmology, Flinders Medical Centre, Flinders University, Adelaide, South Australia, Australia

- <sup>41</sup>Centre for Ophthalmology and Visual Science, Lions Eye Institute, University of Western Australia, Perth, Western Australia, Australia
- <sup>42</sup>Sichuan Provincial Key Laboratory for Human Disease Gene Study, Hospital of the University of Electronic Science and Technology of China and Sichuan Provincial People's Hospital, Chengdu, China
- <sup>43</sup>Sichuan Translational Medicine Hospital, Chinese Academy of Sciences, Chengdu, China
- <sup>44</sup>Molecular Medicine Research Center, State Key Laboratory of Biotherapy, West China Hospital, Sichuan University, Sichuan 610041, China
- <sup>45</sup>EyeCentre Southwest Stuttgart
- <sup>46</sup>University Eye Clinic, Ludwig-Maximilians-University Munich
- <sup>47</sup>Department of Ophthalmology, University Hospital Regensburg, Regensburg, Germany
- <sup>48</sup>Department of Ophthalmology and Visual Sciences, University of Utah, Salt Lake City, UT, USA
- <sup>49</sup>Department of Medicine (Biomedical Genetics), Boston University Schools of Medicine and Public Health, Boston, MA, USA
- <sup>50</sup>Department of Ophthalmology, Boston University Schools of Medicine and Public Health, Boston, MA, USA
- <sup>51</sup>Department of Neurology, Boston University Schools of Medicine and Public Health, Boston, MA, USA
- <sup>52</sup>Department of Epidemiology, Boston University Schools of Medicine and Public Health, Boston, MA, USA
- <sup>53</sup>Department of Biostatistics, Boston University Schools of Medicine and Public Health, Boston, MA, USA
- <sup>54</sup>Department of Ophthalmology, Seoul Metropolitan Government Seoul National University Boramae Medical Center, Seoul, Republic of Korea
- <sup>55</sup>Centre for Experimental Medicine, Queen's University, Belfast, UK
- <sup>56</sup>Department of Ophthalmology, University of Thessaly, School of Medicine, Larissa, Greece
- <sup>57</sup>Department of Ophthalmology, Seoul National University Bundang Hospital, Seongnam, Republic of Korea
- <sup>58</sup>Department of Ophthalmology, Weill Cornell Medical College, New York, NY, USA
- <sup>59</sup>Scheie Eye Institute, Department of Ophthalmology, University of Pennsylvania Perelman School of Medicine, Philadelphia, PA 19104
- <sup>60</sup>Department of Biostatistics and Epidemiology University of Pennsylvania Perelman School of Medicine, Philadelphia, PA 19104
- <sup>61</sup>Department of Ophthalmology, The University of Alabama at Birmingham 1670 University Boulevard VH 360 Birmingham, AL 35294-0019
- <sup>62</sup>INSERM, U968, Paris, F-75012, France

- <sup>63</sup>UPMC Univ Paris 06, UMR\_S 968, Institut de la Vision, Department of Genetics, Paris, F-75012, France
- <sup>64</sup>CNRS, UMR\_7210, Paris, F-75012, France
- <sup>65</sup>Centre Hospitalier National d'Ophtalmologie des Quinze-Vingts, INSERM-DHOS CIC 503, Paris, F-75012, France
- <sup>66</sup>Fondation Ophtalmologique Adolphe de Rothschild, Paris, F-75019, France
- <sup>67</sup>Institute of Ophthalmology, University College of London, London, WC1E 6BT, UK
- <sup>68</sup>Académie des Sciences–Institut de France, Paris, F-75006 France
- <sup>69</sup>Department of Molecular Genetics, Institute of Ophthalmology, London, UK
- <sup>70</sup>Clinical and Experimental Sciences, Faculty of Medicine, University of Southampton, UK
- <sup>71</sup>University Hospital Southampton, Southampton, UK
- <sup>72</sup>Department of Ophthalmology, Hadassah Hebrew University Medical Center, Jerusalem. Israel
- <sup>73</sup>Center for Human Disease Modeling, Duke University USA
- <sup>74</sup>Department of Cell Biology, Duke University USA
- <sup>75</sup>Department of Pediatrics, Duke University USA
- <sup>76</sup>CEPH Fondation Jean Dausset 27 rue Juliette Dodu 75010 Paris, France
- <sup>77</sup>CEA – IG – Centre National de Génotypage 2 rue Gaston Crémieux 91057 Evry Cédex, France
- <sup>78</sup>Cole Eye Institute, Cleveland Clinic, 9500 Euclid Avenue, Cleveland OH 44195
- <sup>79</sup>Department of Ophthalmology, Duke University Medical Center, Durham NC
- <sup>80</sup>Department of Medicine, Duke University Medical Center, Durham NC
- <sup>81</sup>Duke Molecular Physiology Institute, Duke University Medical Center, Durham NC
- <sup>82</sup>Bascom Palmer Eye Institute, University of Miami Miller School of Medicine, 3880 Tamiami Trail North, Naples, FL 34103
- <sup>83</sup>Department of Ophthalmology, Royal Perth Hospital, Perth, Western Australia 6001, Australia
- <sup>84</sup>Department of Ophthalmology, NYU School of Medicine, New York, NY
- <sup>85</sup>Department of Medical Genetics, Cambridge Institute for Medical Research, University of Cambridge, Cambridge, CB2 0QQ, UK
- <sup>86</sup>Department of Ophthalmology, Cambridge University Hospitals NHS Foundation Trust, Cambridge, CB2 0QQ, UK
- <sup>87</sup>Department of Ophthalmology UCSF Medical School, USA
- <sup>88</sup>Department of Ophthalmology and Visual Sciences, Vanderbilt University, Nashville, Tennessee, United States
- <sup>89</sup>Casey Eye Institute, Oregon Health & Science University, Portland OR 97239, USA

<sup>90</sup>Department of Human Genetics, Graduate School of Public Health, University of Pittsburgh, Pittsburgh, PA 15261, USA

<sup>91</sup>School of Clinical Sciences University of Edinburgh, EH16 4SB, Scotland, UK

<sup>92</sup>Center for Inherited Disease Research (CIDR) Institute of Genetic Medicine Johns Hopkins University School of Medicine Baltimore, MD

<sup>93</sup>Department of Ophthalmology, David Geffen School of Medicine—UCLA, Stein Eye Institute, Los Angeles, CA 90095, USA

<sup>94</sup>Department of Human Genetics, David Geffen School of Medicine—UCLA, Los Angeles, CA 90095, USA

<sup>95</sup>Department of Human Genetics, Radboud University Medical Centre, Nijmegen, the Netherlands

<sup>96</sup>Department of Pathology & Cell Biology, Columbia University, New York, NY

<sup>97</sup>Center for Translational Medicine, Moran Eye Center, University of Utah School of Medicine, Salt Lake City UT

<sup>98</sup>Division of Preventive Medicine, Brigham & Women's Hospital, Harvard Medical School

<sup>99</sup>Division of Epidemiology and Clinical Applications, Clinical Trials Branch, National Eye Institute, National Institutes of Health, Bethesda, MD 20892, USA

<sup>100</sup>Institute for Computational Biology, Case Western Reserve University School of Medicine, 2103 Cornell Rd Cleveland, OH 44106

†First authors, ‡Senior authors

\*Correspondence to: [iris.heid@klinik.uni-regensburg.de](mailto:iris.heid@klinik.uni-regensburg.de) (I.M.H.), [goncalo@umich.edu](mailto:goncalo@umich.edu) (G.R.A.), [ski@case.edu](mailto:ski@case.edu) (S.K.I.),

## ABSTRACT

Advanced age-related macular degeneration (AMD) is the leading cause of blindness in the elderly with limited therapeutic options. Here, we report on a study of >12 million variants including 163,714 directly genotyped, most rare, protein-altering variant. Analyzing 16,144 patients and 17,832 controls, we identify 52 independently associated common and rare variants ( $P < 5 \times 10^{-8}$ ) distributed across 34 loci. While wet and dry AMD subtypes exhibit predominantly shared genetics, we identify the first signal specific to wet AMD, near *MMP9* (difference- $P = 4.1 \times 10^{-10}$ ). Very rare coding variants (frequency < 0.1%) in *CFH*, *CFI*, and *TIMP3* suggest causal roles for these genes, as does a splice variant in *SLC16A8*. Our results support the hypothesis that rare coding variants can pinpoint causal genes within known genetic loci and illustrate that applying the approach systematically to detect new loci requires extremely large sample sizes.

Advanced age-related macular degeneration (AMD) is an ocular neurodegenerative disease and the leading cause of vision loss among the elderly with prevalence estimated at 5% for those above 75 years of age<sup>1,2</sup>. The disease is characterized by reduced function of the retinal pigment epithelium (RPE) and loss of photoreceptors in the macula. Advanced AMD is classified as wet (choroidal neovascularization, CNV, when accompanied by angiogenesis) or dry AMD (geographic atrophy, GA, when angiogenesis is absent). These advanced stages of disease are typically preceded by a gradual accumulation of acellular debris in the form of drusen and by pigmentary abnormalities in the macula<sup>3</sup>. Advanced AMD is estimated to affect 9.6 million patients currently worldwide and early AMD stages more than 154.6 million<sup>4</sup>. At present, our understanding of disease biology and therapies remains limited<sup>5</sup>.

Genetic variants, whether associated with small or large changes in disease risk, can help uncover disease mechanisms and provide entry points into therapy. Analysis of common variation have uncovered numerous risk loci for a multitude of complex diseases (see Web Resources) including 21 loci for AMD<sup>6-12</sup>. However, for most disease loci, translation into biological insights remains a major challenge, since the functional consequences of associated common variants are typically subtle<sup>13</sup> and therefore open to inconsistent interpretations.

With advances in sequencing technology, it is expected that genetic analyses will gradually extend to rare variation, which often has more obvious functional consequences<sup>14,15</sup> and thus can accelerate translation of genetic findings into biological understanding<sup>14,16</sup>. For example, identifying multiple disease-associated coding variants in the same gene would provide strong evidence that disrupting gene function leads to disease<sup>17</sup> particularly when these are naturally occurring knock-out alleles. Studies that implicate specific rare variants in complex diseases are few and limited in their generalizability, as they either rely on special populations<sup>8,18,19</sup>, on targeted examinations of a few genes<sup>7,9-11,20,21</sup>, or on genome-wide assessments of relatively modest numbers of individuals<sup>22-25</sup>. In contrast, systematic analyses of common variation are now available in hundreds of thousands of phenotyped individuals<sup>26,27</sup>. Thus, there remains considerable uncertainty about the relative role of rare variants in complex disease and the best strategies to identify highly informative rare variants. Importantly, the optimal sample sizes and study designs for such studies remain poorly understood<sup>16</sup>.

Here, we set out to systematically examine common and rare variation of AMD in the International AMD Genomics Consortium (IAMGCG) incorporating both a genome-wide approach as well as enrichment from a targeted approach. The preceding largest study of AMD examined ~2.4 million variants including ~18,000 imputed or genotyped protein-altering variants using meta-analysis<sup>6</sup>. Customizing a chip for *de novo* centralized genotyping, we analyze >12 million variants including 163,714 directly typed protein-altering variants in



43,566 unrelated subjects of predominantly European ancestry. Our study thus constitutes a detailed simultaneous assessment of common and rare variation in a complex disease and a large sample, setting expectations for other well-powered studies combining common and rare variant information.

## RESULTS

### The study data and genomic heritability

We gathered advanced AMD cases with GA and/or CNV, intermediate AMD cases, and control subjects across 26 studies (**Supplementary Table 1**). While recruitment and ascertainment strategies varied (**Supplementary Table 2**), DNA samples were collected and genotyped centrally. Making maximal use of genotyping technologies, we utilized a chip with (i) the usual genome-wide variant content, (ii) exome content comparable to the exome chip (adding protein-altering variants from across all exons), and a specific customization to add (iii) protein-altering variants detected by our prior sequencing of known AMD loci (see **Methods**) and (iv) previously observed and predicted variation in *TIMP3* and *ABCA4*, two genes implicated in monogenic retinal dystrophies. After quality control, we retained 439,350 directly typed variants including a grid of 264,655 common variants (frequency among controls >1%) distributed across autosomes, sex chromosomes, and mitochondria, primarily (93%) non-coding, and 163,714 directly genotyped protein-altering variants (including 8,290 from known AMD loci), mostly rare (88% with frequency among controls  $\leq$ 1%). Imputation to the 1000 Genomes reference panel enabled examining a total of 12,023,830 variants (**Supplementary Table 3A**). Our final data set included a total of 43,566 subjects consisting of 16,144 advanced AMD patients and 17,832 control subjects of European ancestry for our primary analysis, as well as 6,657 Europeans with intermediate disease and 2,933 subjects with Non-European ancestry (**Supplementary Table 3B, Supplementary Figure 1**).

Altogether, our genotyped markers accounted for 46.7%<sup>28</sup> of variability in advanced AMD risk in the European ancestry subjects (95% confidence interval [CI] 44.5% to 48.8%). Regarding AMD subtypes, estimates for CNV ( $h^2 = 44.3\%$ , CI 42.2% to 46.5%) and GA ( $h^2 = 52.3\%$ , CI 47.2% to 57.4%) were similar; a bivariate analysis<sup>29</sup> showed a high genetic correlation of 0.85 (CI 0.78 to 0.92) between disease subtypes.

### Thirty-Four Susceptibility Loci for AMD

We first conducted a genome-wide single variant analysis of the >12 million genotyped or imputed variants (applying genomic control correction,  $\lambda=1.13$ ) comparing the 16,144 advanced AMD patients and 17,832 control subjects of European ancestry (**full results online; see Web resources**). We obtained >7000 genome-wide significant variants ( $P \leq 5 \times 10^{-8}$ , **Supplementary Figure 2**). To identify independently associated variants, we

adopted sequential forward selection (**Supplementary Figure 3**), resulting in 52 independently associated variants that reach genome-wide significance (**Supplementary Table 4, Supplementary File 1**). These are distributed across 34 locus regions (**Figure 1A**), each extending across the identified and correlated variants,  $r^2 \geq 0.5$ ,  $\pm 500\text{kb}$  (**Supplementary Table 5**). While each of these 52 variants points to a genomic element contributing to AMD biology, variants in the 34 different loci reside relatively far from each other and likely contribute to disease by regulating or modifying the function of different genes. The 34 loci include 16 loci that reached genome-wide significance for the first time (novel loci, **Table 1**) and include genes with compelling biology like extra-cellular matrix genes (*COL4A3*, *MMP19*, *MMP9*), an ABC transporter linked to HDL cholesterol (*ABCA1*), and a key activator in immune function (*PILRB*). Also included are 18 of the 21 AMD loci that reached genome-wide significance previously<sup>6,9</sup> (known loci, **Table 1**), between-study heterogeneity was low, particularly for the new loci (**Supplementary Note 1**).

Most associated variants are common (45 out of 52) with fully conditioned odds ratios (OR) from 1.1 to 2.9 (**Figure 1B, Supplementary Table 4**) with two interacting variants (**Supplementary Note 2**). We also observed seven rare variants with frequencies between 0.01% and 1% and ORs between 1.5 and 47.6 (**Figure 1B, Supplementary Table 4**). All of these variants were also rare in Non-European ancestries (**Supplementary Table 6**, extended association results on Non-European in **Supplementary File 2**). All seven rare variants are located in/near complement genes: four non-synonymous (*CFH*:R1210C, *CFI*:G119R, *C9*:P167S, *C3*:K155Q) and previously found in targeted analyses of complement genes<sup>7-11</sup>; three others (*CFH*: rs148553336, rs191281603, rs35292876) described here for the first time, including two with the rare allele decreasing disease risk by ~2.5 to ~3.3-fold and one increasing risk 1.6 fold. To ensure validity of our results, we verified associations of lead variants in sensitivity analyses that relied on alternate association tests, adjusted for age, gender, or ten ancestry principal components, or were restricted to population-based controls or controls  $\geq 50$  years of age (data not shown). Altogether, our genome-wide single variant analysis nearly doubled the number of AMD loci and has identified several novel rare variants in *CFH*.

### **Prioritizing variants within 52 association signals**

It is often challenging to translate common variant association signals into mechanistic understanding of biology; two key challenges are (i) a large number of variants with similar signals because of linkage disequilibrium and (ii) their often subtle functional consequences. Without narrowing down the lists of candidate variants, follow-up functional experiments are complicated. In our large data set, we were able to prioritize among nearby variants: we computed each variant's ability to explain the observed signal and derived, for each of the 52

signals, the smallest set of variants that included the causal variant with 95% probability<sup>30,31</sup>. The 52 credible sets each included from 1 to >100 variants (total of 1,345 variants, **Supplementary File 3**). For 27 out of the 52 sets, the sets were small with  $\leq 10$  variants (19 signals with  $\leq 5$  variants, **Supplementary Table 7**); seven sets included only one variant -- demonstrating the potential for fine-mapping association signals when dense genotype data is systematically analyzed in large samples. Among the 205 variants with >5% probability of causing the statistical signal, we observe 11 protein-altering (all non-synonymous) variants (versus 2 expected assuming 1% protein-altering variants overall, P for enrichment =  $8.7 \times 10^{-6}$ , **Supplementary Table 8**). These variants provide a focused starting point for future functional analyses, although we recognize that the analysis has limitations [for example, when causal variants are not genotyped nor well-imputed, or when the signal is due to a combination of multiple variants, see **Supplementary Figure 4** for a counter example]. We also note that other variants in each locus (potentially including variants in linkage disequilibrium with lead variants and/or other variants nearby) could also contribute to disease risk.

### Rare Variant Association Signals

Analysis of rare variants that potentially alter peptide sequences (non-synonymous), truncate proteins (premature stop), or affect RNA splicing (splice site) can help to identify causal mechanisms – particularly when multiple such associated variants reside in the same gene<sup>16,32</sup>. We examined the cumulative effect of these protein-altering variants with a frequency  $\leq 1\%$  in each of our ancestry groups. Genome-wide, no signal was detected with  $P \leq 0.05/17,044 = 2.9 \times 10^{-6}$  outside the 34 AMD loci (**Figure 1C**). Within the 34 loci, we found 14 genes with significant disease burden ( $P < 0.05/703$  genes =  $7.1 \times 10^{-5}$ , **Supplementary Table 9**). To eliminate settings where a rare variant burden finding is a linkage disequilibrium shadow of a nearby stronger common variant, we evaluated each burden signal upon its independence from already identified variants in the locus (from **Supplementary Table 4**). Four of the 14 genes preserved a significant ( $P < 0.05/703 = 7.1 \times 10^{-5}$ ) rare variant burden when conditioning on already identified variants in the locus (*CFH*, *CFI*, *TIMP3*, *SLC16A8*; conditioned  $P = 1.2 \times 10^{-6}$ ,  $1.0 \times 10^{-8}$ ,  $9.0 \times 10^{-8}$ , or  $3.1 \times 10^{-6}$ , respectively, **Table 2**). Sensitivity analyses provide similar (excluding previously sequenced subjects) and extended results (prioritizing variants with high predicted functionality, **Supplementary Note 3**). Several interesting patterns emerge, many of which we owe to our chip design.

First, three of the four rare variant burden signals (*CFH*, *CFI*, *TIMP3*) are due to very rare variants, each with frequency  $< 0.1\%$ , all genotyped (**Supplementary File 4**). Many human genetic studies have used frequency thresholds of 1% to 5% as a working definition of “rare”, but our data suggests that trait associated variants with clear function may often be

much rarer – likely necessitating very large sample sizes for analysis. In two genes (*CFH*, *CFI*), the rare burden was detected because we enriched arrays with variants from previous sequencing of known AMD loci in cases and controls<sup>10</sup> (54 of 80 variants). The burden findings in *CFH* (new, **Supplementary Note 4**) and *CFI*<sup>9</sup> together with variants *CFH*:R1210C and *CFI*:G119R<sup>7,9</sup>, corroborate a causal role for these genes in AMD etiology.

The third signal (*TIMP3*) was in a gene previously associated with Sorsby's fundus dystrophy, a rare disease with early onset at <45 years of age but with clinical presentation strikingly similar to AMD<sup>33,34</sup>. Because the majority of Sorsby's alleles disrupt cysteine-cysteine bonds in *TIMP3*, we arrayed all possible cysteine disrupting sites together with other previously described Sorsby's risk alleles<sup>33,34</sup>. The nine rarest *TIMP3* variants were cumulatively associated with >30-fold increased risk of disease. *TIMP3* resides in an established AMD locus<sup>5,35</sup> targeted in previous sequencing efforts<sup>32,35</sup>, that were too small to demonstrate an excess of rare variation on this scale (1 variant in 17,832 controls versus 29 variants in 16,144 cases). Interestingly, although Sorsby-associated *TIMP3* variants typically occur in exon 5, four of the unpaired cysteine residues we observed map to other exons – perhaps because unpaired cysteines in different locations impair protein folding in different ways, contributing to variation in disease severity or age of onset: disease onset for our 29 cases with *TIMP3* variants was  $\geq 50$  years of age (average 64.5 years). AMD cases with these rare *TIMP3* risk alleles still exhibited much higher counts of AMD risk alleles across the genome than controls, suggesting that *TIMP3* is not a monogenic cause of AMD but contributes to disease together with alleles at the other risk loci. Our finding illustrates a locus where complex and monogenic disorders arise from variation in the same gene, similar to *MC4R* and *POMC* in obesity<sup>36</sup> or *UMOD* in kidney function<sup>37</sup>. In a similar approach, we analyzed 146 rare protein-altering variants in *ABCA4*, a gene underlying Stargardt disease<sup>38</sup>, but found no association (P=0.97).

The rare variant burden signal in *SLC16A8* was primarily driven by a putative splice variant (c.214+1G>C, rs77968014, minor allele frequency among controls, CAF = 0.81%, OR = 1.5, imputed with R<sup>2</sup>=0.87, **Supplementary File 4**). This is thus not a true burden from multiple rare variants, but a single variant emerging as significant due to the reduced multiple testing from gene-wide testing (single variant association P =  $9.1 \times 10^{-6}$ , conditioned on rs8135665 P =  $1.3 \times 10^{-6}$ ). This variant is interesting as it is predicted to disrupt processing of the encoded transcript (as +1 G variant, Human Splicing Finder 3.0); however, functional analyses in relevant tissue would be required to substantiate the direct implication for gene function and AMD. *SLC16A8* encodes a cell membrane transporter, involved in transport of pyruvate, lactate and related compounds across cell membranes<sup>39</sup>. This class of proteins mediates the acidity level in the outer retinal segments, and *SLC16A8* gene knock-out animals have changes in visual function and scotopic electroretinograms, but not overt retinal

pathology<sup>40</sup>. Interestingly, a progressive loss of *SLC16A8* expression in eyes affected with GA was reported with increasing severity of disease<sup>41</sup>. In summary, our chip design and our large data set enabled us not only to detect interesting features of AMD genetics, but also to provide guidance for future investigations on rare variants.

### **From Disease Loci to Biological Insights**

In addition to fine-mapping and searches for protein-altering rare variants, many analyses can further narrow the list of candidate genes in our identified loci. We annotated the 368 genes closest to our 52 association signals (index variant and proxies,  $r^2 \geq 0.5$ ,  $\pm 100\text{kb}$ , **Supplementary File 5**), noting among these the genes those that contained associated credible set variants (**Supplementary File 3**) or a rare variant burden (**Table 2**) – these are the highest priority candidates, consistent with previous analysis of putative cis-regulatory variants<sup>42</sup>. We further checked whether genes were expressed in retina (82.6% of genes) or RPE/choroid (86.4%, **Supplementary File 6**). We sought relevant eye phenotypes in genetically modified mice (observed in 32 of the 368 queried genes, **Supplementary File 7**). We tagged genes in biological pathways enriched across loci, such as the alternative complement pathway, HDL transport, and extracellular matrix organization and assembly (**Supplementary Table 10**) – highlighting genes that connect multiple pathways (*COL4A3/COL4A4*, *ABCA1*, *MMP9*, and *VTN*). We also highlighted genes that were approved or experimental drug targets (31 of the 368 queried, **Supplementary File 8**). Finally, we prioritized genes where at least one of the credible set variants (**Supplementary File 3**) was protein-altering or located in a putative functional region (promoter, 3'/5' UTR).

All this information is summarized in the gene priority score table (**Supplementary File 9, Supplementary Note 5**), which uses a simple customizable scoring scheme to assign priority: the scheme using equal weights for each column assigns highest scores per novel locus (**Figure 2A, Supplementary Table 11**) to genes such as master regulators of immune function (*PILRB*), matrix metalloproteinase genes (*MMP9*, *MMP19*), a gene involved in lipid transport (*ABCA1*), a gene playing a role in lipid peroxidation and inflammation (*GPX4*), an inhibitor of the complement cascade (*VTN*), another collagen gene known to cause Alport's syndrome (*COL4A3*), a gene causing a developmental monogenic disorder, the Noonan syndrome (*PTPN11*), and a retinol dehydrogenase involved in the regeneration of cone and rod photoreceptor segments previously associated with autosomal recessive night-blindness (*RDH5*). All of these are expressed in relevant tissues, several of these show relevant mouse phenotypes (*MMP9*, *MMP19*, *COL4A3*, *PTPN11*, *GPX4*, and *RDH5*), and six of these are current drug targets (*ABCA1*, *MMP19*, *RDH5*, *PTPN11*, *VTN*, *GPX4*). In the known AMD loci, the highest scores per locus included the usual suspects (*CFH*, *CFI*, *CFB*, *C3*, and *APOE*) as well as *TIMP3* and *SLC16A8* (**Figure 2B**). This summary of evidence may help prioritize

genes for follow-up functional experiments. It should be noted that much of the information was collected specifically for the genes in the identified loci (for example, by reviewing literature for animal models for each respective gene), rather than systematically annotating all genes genome-wide, so that this summary of evidence is not amenable to formal statistical enrichment analysis.

### Commonalities and differences between advanced AMD subtypes

Previously identified risk variants all contribute to the two advanced AMD subtypes, CNV and GA. We compared association signals between our 10,749 cases with CNV and 3,235 cases with GA. Four of the 34 lead variants show significant difference ( $P_{\text{diff}} < 0.05/34 = 0.00147$ ) between disease subtypes (in the loci *ARMS2/HTRA1*, *CETP*, *MMP9*, *SYN3/TIMP3*, **Figure 3A, Supplementary Table 12**). Variant rs42450006 upstream of *MMP9* was the only one that was specific to one subtype, being exclusively associated with CNV (frequency in controls = 14.1%;  $OR_{\text{CNV}} = 0.78$  vs.  $OR_{\text{GA}} = 1.04$ ;  $P_{\text{diff}} = 4.1 \times 10^{-10}$ ), but not with GA ( $P_{\text{GA}}=0.39$ ). The signal was markedly stronger in an analysis restricted to CNV (**Supplementary Note 6**). The *MMP9* signal for neovascular disease fits well with prior evidence: upregulation of *MMP9* appears to induce neovascularization<sup>43</sup>; a feedback loop between *VEGF* signaling and *MMP9* has been proposed in the RPE<sup>44</sup>. VEGF currently provides an effective therapy for patients with neovascular AMD, but the struggle to keep vision continues. Beyond confirming a shared genetic predisposition of the two subtypes, our data identifies – for the first time – one variant that is specific to one subtype.

### Commonalities and differences between advanced AMD and earlier disease stages

We evaluated our association signals in 6,657 individuals with intermediate AMD, defined as having more than five macular drusen greater than 63 $\mu\text{m}$  and/or pigmentary changes in the RPE. Examining all genotyped variants<sup>28</sup>, we found a correlation of  $\rho = 0.78$ , indicating substantial overlap between genetic determinants of advanced effects and intermediate AMD (95% CI 0.69 to 0.87). Among our 34 index variants, 24 showed nominally significant association ( $P_{\text{intermediate}} \leq 0.05$ ) with intermediate AMD (2 expected,  $P_{\text{binomial}} = 4.8 \times 10^{-24}$ ); all had ORs in the same direction but smaller in magnitude (**Figure 3B, Supplementary Table 13**). The other 10 variants showed no association with intermediate AMD ( $P_{\text{intermediate}} > 0.05$ ), despite sufficient power (**Supplementary Table 14**). Interestingly, these 10 variants point to 7 extra-cellular matrix genes (*COL15A1*, *COL8A1*, *MMP9*, *PCOLCE*, *MMP19*, *CTRB1/2*, *ITGA7*, **Supplementary Table 15**), based on which one may hypothesize that the extra-cellular matrix points to a disease subtype without early stage manifestation or with extremely rapid progression. If confirmed, a group of rapidly progressing patients or without

early symptoms might eventually derive maximum benefit from genetic diagnosis and future preventive therapies.

### **An Accounting of AMD Genetics**

To account for progress made here in understanding AMD genetics, we estimated the proportion of disease risk explained by our 52 independent variants and compared it to our initial estimates of heritability obtained by examining all genotyped variants. We computed a weighted risk score of the 52 variants<sup>45</sup> and modeled a realistic genetic risk score distribution (see **Materials and Methods**). Individuals in the highest decile of genetic risk have a 44-fold increased risk of developing advanced AMD compared to the lowest decile; of these, 22.7% are predicted to have AMD in an elderly general population above 75 years of age with ~5% disease prevalence (**Figure 4A, Supplementary Table 16**). Altogether, the 52 variants explain 27.2% of disease variability (**Figure 4B**, also highlighting results based on other prevalence assumptions), including a 1.4% contribution from rare variants. The 52 identified variants thus explain more than half of the genomic heritability (estimated as 46.7%, see first results chapter). The balance might be attributed to additional variation not studied here, or to genetic interaction with environmental factors such as smoking, diet or sunlight exposure, or to chance.

## **DISCUSSION**

We set out to improve our understanding of rare and common genetic variation for macular degeneration biology, so as to guide the development of therapeutic interventions and facilitate early diagnosis, monitoring and prevention of disease. AMD is an ideal role model to study complex disease genetics: it was the focus of the first successful genome-wide study of common variants<sup>46</sup>, and a total of 21 disease susceptibility loci with a broad range of effect sizes have been identified altogether<sup>6-12</sup>. Here, we systematically examine rare variation (through direct genotyping) and common variation (through genotyping and imputation) for AMD in a study designed to discover >80% of associated protein-altering variants with an allele frequency of >0.1% and >3-fold increased disease risk (or >0.5% frequency and >1.8-fold increased disease risk). Our study provides a comprehensive simultaneous assessment of common and rare variation enabling us to understand the relative roles of rare and common variants and the scientific insights to be gained from rare variation.

Rare protein-altering variants are an especially attractive target for genetic studies because most of these variants are expected to damage gene function. Furthermore, observing that many rare variants in a gene are, together, associated with a change in disease risk strongly suggests that the gene is causally implicated in disease biology and –

further – suggests the consequences of mimicking or blocking gene action using a drug. Our study demonstrates that when rare variants are systematically assessed in genome-wide assessments of large numbers of cases and controls, significant signals can be assigned to single rare variants as well as to rare variant burden in specific genes.

Our study also demonstrates the challenges of these analyses. For three of the genes where we identified a rare variant burden, the accumulated evidence was spread across very rare variants with frequencies <0.1% in controls. Most of these variants derived from our enrichment of the chip with protein-altering variants in known AMD loci based on our own sequencing including AMD patients. This emphasizes the value of a hybrid approach with direct targeted sequencing in large sample sizes including patients to detect very rare variants and genotyping these variants in even larger sample sizes for association analysis. Another conclusion is about required sample sizes: although such rare variants are expected to exist in nearly all genes, no rare variant burden was observed in most of the 34 loci we studied. For these loci, identifying causal mechanisms through the study of rare protein-altering variants will require even larger sample sizes to identify variants missed by our customized exome arrays. While our findings of rare variant burden are predominantly from targeted enrichment, the knowledge about effect sizes and frequencies of contributing variants illustrates that applying the approach genome-wide to detect new loci requires extremely large sample sizes. In our view, a recent estimate that sequencing of 25,000 cases will be needed to identify genes where rare variants have a substantial impact on disease risk is likely to be a starting point for rare variant analysis, rather than an ultimate target, particularly given the fact that effect sizes for AMD risk alleles appear to be larger than for many other complex traits<sup>16</sup>.

In addition to corroborating previous reports of rare variants that disrupt genes in the complement pathway and lead to large increases in disease risk, our study also includes two unexpected rare variant findings. First, we show that a putative splice variant in *SLC16A8* can greatly increase the risk of age-related macular degeneration – providing strong evidence that the gene is directly involved in disease biology. *SLC16A8* is a lactate transporter expressed<sup>39</sup> specifically by the RPE, and a deficit of lactate transport toward the choroid vasculature results in acidification of the retina and photoreceptor dysfunction as reported for *SLC16A8* knock-out mice<sup>40</sup>. Second, we show a >30-fold excess of rare *TIMP3* mutations among putative cases of macular degeneration. *TIMP3* is an especially attractive candidate that has been the subject of previous, underpowered, genetic association studies.

While it has been hypothesized that studies of rare and low frequency genetic variants will greatly increase the proportion of genetic risk that can be explained, our results don't support this. Our study and others successfully identify many low frequency disease risk alleles, and these provide clues about disease biology, but our results also show that



common variants make a much larger contribution to variability in disease risk. Common variants suggest a large number of interesting leads and pathways for future analysis (**Supplementary Table 11, Figure 2A**), including attractive candidates such as immune regulators (*PILRB*), genes implicated in mouse ocular phenotypes (*MMP9*, *MMP19*, *COL4A3*, *PTPN11*, *GPX4*, and *RDH5*), and proven drug targets (*ABCA1*, *MMP19*, *RDH5*, *PTPN11*, *VTN*, *GPX4*). In a literature search, we identified no previous candidate gene association studies targeting our novel loci, although several model organism, cellular, and functional studies evaluated potential links between genes in these loci and AMD (highlights of this search in **Supplementary Table 11**) and a few loci were nominally associated and proposed as candidates in prior genome-wide searches<sup>47,48</sup>. As richer functional annotations of the genome<sup>49</sup> become available in diverse cell types, systematic assessment of overlap between these and our loci should clarify disease biology.

Our study also suggests additional important observations. While our results show that the majority of genetic risk is shared between GA and CNV, we also identify – for the first time – a variant that is specific to one advanced AMD subtype: a genetic variant near *MMP9* is specific to CNV, a candidate gene also supported by prior gene expression analyses in the Bruch’s membrane of patients with neovascular disease<sup>50</sup>. Future efforts extending to longitudinal data might help improve the dissection of pure CNV and pure GA and their genetic make-up even further, but longitudinal data has still to be extended to yield sufficient sample sizes. If substantiated, the fact that nearly all disease associated variants modulate risk of both CNV and GA has potentially significant therapeutic consequences. It implies that individuals at high risk of CNV are also at high risk of GA. This suggests that therapeutic strategies which mitigate CNV but not GA will only provide temporary relief to patients – who are likely to remain at high risk of developing GA and may still require future interventions to prevent it.

Therefore, our findings have several important implications for future studies of rare variation in human complex traits. First, they clearly emphasize the need for very large sample sizes in population studies: the functionally most interesting variants we identify have frequencies in the range of 0.01 – 1.0% and, despite their strong impact on disease risk, could only be implicated using 10,000s of individuals. Second, they illustrate the value of hybrid approaches, where sequencing is used to detect interesting variants and custom arrays and imputation are used to examine these variants in very large samples. Since all the large effect rare variants we identify reside in or near GWAS loci, as with most complex trait associated rare variants<sup>7-11,20,21,23,51</sup>, focused studies around GWAS loci may continue to be a cost-effective compromise. Third, our analysis of cysteine variants in *TIMP3* illustrates not only the potential for targeted variant discovery but the critical need to understand the consequences of rare variants when analyzing them together. While very large samples will

be needed, our results also show that the effort to extend genetic studies to rare variants is worthwhile as these variants can pinpoint causal genes and advance our understanding of disease biology.

## References

1. Smith, W. *et al.* Risk factors for age-related macular degeneration: Pooled findings from three continents. *Ophthalmology* **108**, 697-704 (2001).
2. Chakravarthy, U., Evans, J. & Rosenfeld, P.J. Age related macular degeneration. *BMJ* **340**, c981 (2010).
3. Ferris, F.L. *et al.* A simplified severity scale for age-related macular degeneration: AREDS Report No. 18. *Arch Ophthalmol* **123**, 1570-4 (2005).
4. Wong, W.L. *et al.* Global prevalence of age-related macular degeneration and disease burden projection for 2020 and 2040: a systematic review and meta-analysis. *Lancet Glob Health* **2**, e106-16 (2014).
5. Fritsche, L.G. *et al.* Age-related macular degeneration: genetics and biology coming together. *Annu Rev Genomics Hum Genet* **15**, 151-71 (2014).
6. Fritsche, L.G. *et al.* Seven new loci associated with age-related macular degeneration. *Nat Genet* **45**, 433-9, 439e1-2 (2013).
7. Raychaudhuri, S. *et al.* A rare penetrant mutation in CFH confers high risk of age-related macular degeneration. *Nat Genet* **43**, 1232-6 (2011).
8. Helgason, H. *et al.* A rare nonsynonymous sequence variant in C3 is associated with high risk of age-related macular degeneration. *Nat Genet* **45**, 1371-4 (2013).
9. Seddon, J.M. *et al.* Rare variants in CFI, C3 and C9 are associated with high risk of advanced age-related macular degeneration. *Nat Genet* **45**, 1366-70 (2013).
10. Zhan, X. *et al.* Identification of a rare coding variant in complement 3 associated with age-related macular degeneration. *Nat Genet* **45**, 1375-9 (2013).
11. van de Ven, J.P. *et al.* A functional variant in the CFI gene confers a high risk of age-related macular degeneration. *Nat Genet* **45**, 813-7 (2013).
12. Arakawa, S. *et al.* Genome-wide association study identifies two susceptibility loci for exudative age-related macular degeneration in the Japanese population. *Nat Genet* **43**, 1001-4 (2011).
13. Gibson, G. Rare and common variants: twenty arguments. *Nat Rev Genet* **13**, 135-45 (2011).
14. Do, R., Kathiresan, S. & Abecasis, G.R. Exome sequencing and complex disease: practical aspects of rare variant association studies. *Hum Mol Genet* **21**, R1-9 (2012).
15. Nelson, M.R. *et al.* An abundance of rare functional variants in 202 drug target genes sequenced in 14,002 people. *Science* **337**, 100-4 (2012).
16. Zuk, O. *et al.* Searching for missing heritability: designing rare variant association studies. *Proc Natl Acad Sci U S A* **111**, E455-64 (2014).
17. Vogelstein, B. *et al.* Cancer genome landscapes. *Science* **339**, 1546-58 (2013).
18. Styrkarsdottir, U. *et al.* Severe osteoarthritis of the hand associates with common variants within the ALDH1A2 gene and with rare variants at 1p31. *Nat Genet* **46**, 498-502 (2014).
19. Styrkarsdottir, U. *et al.* Nonsense mutation in the LGR4 gene is associated with several human diseases and other traits. *Nature* **497**, 517-20 (2013).
20. Rivas, M.A. *et al.* Deep resequencing of GWAS loci identifies independent rare variants associated with inflammatory bowel disease. *Nat Genet* **43**, 1066-73 (2011).
21. Flannick, J. *et al.* Loss-of-function mutations in SLC30A8 protect against type 2 diabetes. *Nat Genet* **46**, 357-63 (2014).
22. Cruchaga, C. *et al.* Rare coding variants in the phospholipase D3 gene confer risk for Alzheimer's disease. *Nature* **505**, 550-4 (2014).
23. Do, R. *et al.* Exome sequencing identifies rare LDLR and APOA5 alleles conferring risk for myocardial infarction. *Nature* **518**, 102-6 (2015).
24. Lange, L.A. *et al.* Whole-exome sequencing identifies rare and low-frequency coding variants associated with LDL cholesterol. *Am J Hum Genet* **94**, 233-45 (2014).
25. Walters, R.G. *et al.* A new highly penetrant form of obesity due to deletions on chromosome 16p11.2. *Nature* **463**, 671-5 (2010).
26. Locke, A.E. *et al.* Genetic studies of body mass index yield new insights for obesity biology. *Nature* **518**, 197-206 (2015).
27. Shungin, D. *et al.* New genetic loci link adipose and insulin biology to body fat distribution. *Nature* **518**, 187-96 (2015).
28. Yang, J., Lee, S.H., Goddard, M.E. & Visscher, P.M. GCTA: a tool for genome-wide complex trait analysis. *Am J Hum Genet* **88**, 76-82 (2011).
29. Lee, S.H., Yang, J., Goddard, M.E., Visscher, P.M. & Wray, N.R. Estimation of pleiotropy between complex diseases using single-nucleotide polymorphism-derived genomic relationships and restricted maximum likelihood. *Bioinformatics* **28**, 2540-2 (2012).

30. Wellcome Trust Case Control, C. *et al.* Bayesian refinement of association signals for 14 loci in 3 common diseases. *Nat Genet* **44**, 1294-301 (2012).
31. Wen, X. Bayesian model selection in complex linear systems, as illustrated in genetic association studies. *Biometrics* **70**, 73-83 (2014).
32. Nejentsev, S., Walker, N., Riches, D., Egholm, M. & Todd, J.A. Rare variants of IFIH1, a gene implicated in antiviral responses, protect against type 1 diabetes. *Science* **324**, 387-9 (2009).
33. Sorsby, A. & Mason, M.E. A fundus dystrophy with unusual features. *Br J Ophthalmol* **33**, 67-97 (1949).
34. Weber, B.H., Vogt, G., Wolz, W., Ives, E.J. & Ewing, C.C. Sorsby's fundus dystrophy is genetically linked to chromosome 22q13-qter. *Nat Genet* **7**, 158-61 (1994).
35. Abecasis, G.R. *et al.* Age-related macular degeneration: a high-resolution genome scan for susceptibility loci in a population enriched for late-stage disease. *Am J Hum Genet* **74**, 482-94 (2004).
36. Speliotes, E.K. *et al.* Association analyses of 249,796 individuals reveal 18 new loci associated with body mass index. *Nat Genet* **42**, 937-48 (2010).
37. Kottgen, A. *et al.* New loci associated with kidney function and chronic kidney disease. *Nat Genet* **42**, 376-84 (2010).
38. Allikmets, R. *et al.* Mutation of the Stargardt disease gene (ABCR) in age-related macular degeneration. *Science* **277**, 1805-7 (1997).
39. Halestrap, A.P. The SLC16 gene family - structure, role and regulation in health and disease. *Mol Aspects Med* **34**, 337-49 (2013).
40. Daniele, L.L., Sauer, B., Gallagher, S.M., Pugh, E.N., Jr. & Philp, N.J. Altered visual function in monocarboxylate transporter 3 (Slc16a8) knockout mice. *Am J Physiol Cell Physiol* **295**, C451-7 (2008).
41. Shoshan, V., MacLennan, D.H. & Wood, D.S. A proton gradient controls a calcium-release channel in sarcoplasmic reticulum. *Proc Natl Acad Sci U S A* **78**, 4828-32 (1981).
42. Stranger, B.E. *et al.* Patterns of cis regulatory variation in diverse human populations. *PLoS Genet* **8**, e1002639 (2012).
43. Lambert, C. *et al.* Gene expression pattern of cells from inflamed and normal areas of osteoarthritis synovial membrane. *Arthritis Rheumatol* **66**, 960-8 (2014).
44. Hollborn, M. *et al.* Positive feedback regulation between MMP-9 and VEGF in human RPE cells. *Invest Ophthalmol Vis Sci* **48**, 4360-7 (2007).
45. Rudnicka, A.R. *et al.* Age and gender variations in age-related macular degeneration prevalence in populations of European ancestry: a meta-analysis. *Ophthalmology* **119**, 571-80 (2012).
46. Klein, R.J. *et al.* Complement factor H polymorphism in age-related macular degeneration. *Science* **308**, 385-9 (2005).
47. Chen, W. *et al.* Genetic variants near TIMP3 and high-density lipoprotein-associated loci influence susceptibility to age-related macular degeneration. *Proc Natl Acad Sci U S A* **107**, 7401-6 (2010).
48. Logue, M.W. *et al.* A search for age-related macular degeneration risk variants in Alzheimer disease genes and pathways. *Neurobiol Aging* **35**, 1510.e7-18 (2014).
49. The Encode Project Consortium. An integrated encyclopedia of DNA elements in the human genome. *Nature* **489**, 57-74 (2012).
50. Hussain, A.A., Lee, Y., Zhang, J.J. & Marshall, J. Disturbed matrix metalloproteinase activity of Bruch's membrane in age-related macular degeneration. *Invest Ophthalmol Vis Sci* **52**, 4459-66 (2011).
51. Johansen, C.T. *et al.* Excess of rare variants in genes identified by genome-wide association study of hypertriglyceridemia. *Nat Genet* **42**, 684-7 (2010).
52. Price, A.L. *et al.* Pooled association tests for rare variants in exon-resequencing studies. *Am J Hum Genet* **86**, 832-8 (2010).

### Acknowledgments:

Data permitted for sharing by respective Institutional Review Boards, and/summary statistics reported in the paper will be archived in the database of Genotypes and Phenotypes (dbGaP; <http://www.ncbi.nlm.nih.gov/gap>).

The UWA, LEI & Flinders group acknowledges financial support for participant recruitment and sample processing provided by the National Health and Medical Research Council

(NHMRC) of Australia (#1023911), the Ophthalmic Research Institute of Australia, the BrightFocus Foundation and a Ramaciotti Establishment Grant. CERA receives Operational Infrastructure Support from the Victorian Government. KPB, JEC and AWH are supported by NHMRC Fellowships. The authors acknowledge the support of B. Usher-Ridge, L. Palmer, L. Ma and DL Lim in patient recruitment and data collection.

The Pittsburgh group acknowledges funding to MBG from NIH/NEI R01 EY09859, Research to Prevent Blindness (N.Y, N.Y.), Harold and Pauline Price Foundation.

The BDES (Beaver Dam Eye Study) was supported by grant EY06594 (to RK and BEKK) from the National Institutes of Health, as well as Senior Scientific Investigator Awards (to RK and BEKK) and an unrestricted grant (to the University of Wisconsin Department of Ophthalmology and Visual Sciences) from Research to Prevent Blindness.

The Cambridge group was supported by the Medical Research Council, UK (grant G0000067 to JRWY, ATM), the Macular Disease Society (JRWY, ATM); the Guide Dogs for the Blind Association (ATM, JRWY) and the Department of Health's NIHR Biomedical Research Centre for Ophthalmology at Moorfields Eye Hospital and UCL Institute of Ophthalmology. We thank the clinicians who helped with recruitment, the Reading Centre at Moorfields Eye Hospital, London for grading fundus photographs and the subjects who participated in the research.

The EUGENDA-Cologne group was supported by a grant from the Retinovit foundation.

The Vanderbilt group was supported by the National Institutes of Health Grants AG019085 (JLH), EY023164 (JLH), EY022310 (JLH), EY012118 (JLH, AA, MAB), AG044089 (JDH), T32 EY007157 (JNCB) and a PhRMA Informatics fellowship (JNCB).

The MMAP-Penn group was supported by the Arnold and Mabel Beckman Initiative for Macular Research (CAC), Research to Prevent Blindness Inc (CAC), EyeSight Foundation of Alabama (CAC), and NIH EY023164 (DS).

The Oregon group was supported by National Eye Institute grants EY021532, and EY0105712, and an unrestricted departmental grant from Research to Prevent Blindness.

The Edinburgh group using the Scottish AMD study was funded by the Chief scientists Office (Scotland) CSO reference number: CZB/4/79 and would like to thank Alan Wright, Ana Armbrecht and Fraser Imrie for collecting the samples and all of the individuals who participated in this study.

The EU/JHU study acknowledges the support of the CEPH Biological Resource Centre by the French Ministère de l'Enseignement Supérieur et de la Recherche, Foundation Fighting Blindness Clinical Research Institute (FFB, CRI), an unrestricted grant to the Wilmer Eye Institute from Research to Prevent Blindness, and Baylor-Johns Hopkins Center for Mendelian Genetics (National Human Genome Research Institute, NHGRI/NIH; 1U54HG006542-01).

The Jerusalem study was supported by grants from the Israel Science fund (ISF) and the Israeli Ministry of Health.

The Southampton study acknowledges Southampton Wellcome Trust Clinical Research Facility for research nurse support in collecting DNA samples, Helen Griffiths (Clinical and Experimental Sciences, University of Southampton) for technical support in processing DNA and all the patients who contributed to this work. AJL supported at the University of Southampton by funding from The Wellcome Trust (076169/A), American Health Assistance Foundation (M2007110), Macula Vision Research Foundation, TFC Frost Charitable Trust, Brian Mercer Charitable Trust, Macular Society, Hobart Trust and the Gift of Sight appeal.

The Marshfield group was supported by grants NIH NCATS: UL1TR000427, NIH NHGRI: 1U01HG006389, and support from the Marshfield Clinic Research Foundation.

The Melbourne study was supported by the National Health and Medical Research Council Australia, project grant 1008979, Centre for Clinical Research Excellence #529923 - Translational Clinical Research in Major Eye Diseases. NHMRC Research Fellowship (PNB, #1028444). CERA receives Operational Infrastructure Support from the Victorian Government.

The Miami group was supported by National Institutes of Health Grants R01 EY012118 (MAP-V, WKS, JLK, SGS, MDC), EY023164 (MAP-V, WKS, JLH, MAP-V), EY022310 (MAP-V) and T32 EY023194 (RJS) and P30-EY005722. All Bascom Palmer Eye Institute authors are partially supported by NIH Center Core Grant P30EY014801 and an unrestricted grant from Research to Prevent Blindness, New York, NY, USA.

The MMAP-Michigan and AREDS groups were supported by Intramural Research Program of the National Eye Institute (ZO1 EY000475); the AREDS study was supported by the National Eye Institute/National Institutes of Health, (contract no.: HHS-NOI-EY-0-2127), Bethesda Maryland; the AREDS2 study was supported by the intramural program funds and contracts from the National Eye Institute/National Institutes of Health (NEI/NIH), Department of Health and Human Services, Bethesda, MD. Contract No. HHS-N-260-2005-00007-C. ADB Contract No. N01-EY-5-0007. The Michigan study was supported by the National Eye Institute (EY0022005) and the National Human Genome Research Institute (HG006513 HG007022), Foundation Fighting Blindness and National Institutes of Health/National Eye Institute Grant-EY016862.

The NHS/HPF studies were supported by EY021900, EY017362, EY13824, EY009611, CA87969, CA49449, and HL35464.

The Regensburg group was supported by BMBF-01ER1206 (to IMH), EFKS 2012\_A147 (IMH), BMBF-01GP1308 (IMH), the Deutsche Forschungsgemeinschaft (grant WE 1259/19-1 and WE1259/19-2, BHFw), and the Alcon Research Institute (BHFw).

The Rotterdam-Clinic study was supported by ZoNMW project number: 170885606, MDfonds, Landelijke Stichting voor Blinden en Slechtzienenden (LSBS).

The UCSD study was supported by grants from NIH (grants EY014428, EY018660, P30EY022589), 863 Program (2014AA021604), and Research to Prevent Blindness. ZS is supported by 863 Program (2014AA021604), ZY is supported by National Natural Science Foundation of China (81170883 and 81430008), KZ is supported by NIH grants (1R01EY018660-01A10) and VA Merit Award.

The EUGENDA-Neijmegen study was supported by MD Fonds, Gelderse Blindenstichting, Algemene Vereniging ter Voorkoming van Blindheid, Stichting Nederlands Oogheelkundig Onderzoek, Oogfonds.

The Utah study was supported by the ALSAM Foundation, an unrestricted grant from Research to Prevent Blindness to the Department of Ophthalmology and Visual Sciences, University of Utah, SOM, Moran Eye Center.

The Seoul National University Bundang group was supported by grants from the National Research Foundation of Korea, funded by the Ministry of Education, Science, and Technology (grant numbers; NRF-2009-0072603 and NRF-2012R1A1A2008943).

The Westmead/Sydney samples were collected in three studies that were supported by the National Health and Medical Research Council (NHMRC), Australia: Grant IDs 974159, 211069, 457349 and 512423 supported the Blue Mountains Eye Study that provided population-based controls; Grant ID 302010 supported the Cataract Surgery and Risk of Age-related Macular Degeneration study that provided clinic-based early and late AMD cases and controls; and Grant ID 571013 supported the Genes and Environment in late AMD study that provided clinic-based late AMD cases. NHMRC Senior Research Fellowship (JJW, 358702, 632909). NHMRC Senior Research Fellowship (JJW, 358702, 632909). The NHMRC had no role in the design or conduct of these studies.

The Columbia study was supported by the National Institutes of Health/NIH grants R01-EY013435 and P30-EY019007, and Research to Prevent Blindness (New York, NY).

The CWRU group was supported by VA Merit Review (NSP), Foundation Fighting Blindness (SAH), Research to Prevent Blindness (SAH); International Retinal Research Foundation (SKI).

CIDR Program contract number HHSN268201200008I.

We thank all participants of all the studies included for enabling this research by their participation to these studies. Computer resources for this project have been provided by the High Performance Computing Centers of the University of Michigan and the University of Regensburg. The Regensburg Team would like to thank Randy Rueckner for technical assistance.

## Figure legends

**Figure 1. Genome-wide search reveals 34 loci and genes with rare variant burden for AMD.** (a) We conducted a genome-wide single variant association analysis for >12 million variants in 16,144 advanced AMD patients versus 17,832 controls. Shown is the Manhattan Plot exhibiting P-values for association highlighting novel ( $P < 5 \times 10^{-8}$  for the first time, green) and known (blue) AMD loci (see **Table 1**). (b) We computed independent effect size (log Odds Ratios) of each of the 52 identified variants (**Supplementary Table 4**). Shown are these effect sizes versus the frequency of the AMD risk increasing allele and a 80% power curve. (c) We conducted a genome-wide gene-based test for disease burden based on the protein-altering variants testing 17,044 RefSeq genes by the variable threshold test<sup>52</sup>. Shown is the Manhattan Plot with P-values, the red horizontal line indicating genome-wide significance ( $P \leq 0.05/17,044 = 2.9 \times 10^{-6}$ ) and the yellow line indicating AMD-locus-wide significance (given 703 genes in the 34 AMD loci,  $P \leq 0.05/703 = 7.1 \times 10^{-5}$ ). No gene outside the 34 loci is genome-wide significant; 14 genes are AMD-locus-wide significant (blue), four remain significant after locus-wide conditioning (bold letters, **Supplementary Table 9**).

**Figure 2. Genes with top priority based on biological and statistical evidence combined.** We queried 368 genes in the 34 narrow AMD regions (index and proxies,  $r^2 \geq 0.5$ ,  $\pm 100$ kb) for biological (red; expression in retina/RPE/choroid, **Supplementary File 6**; ocular mouse phenotype, **Supplementary File 7**), statistical, (blue;  $\geq 1$  credible set variant in gene  $\pm 50$  kb, **Supplementary File 3**; rare variant burden, **Table 2**), putative functional (green;  $\geq 1$  credible set variant in gene  $\pm 50$  kb being protein-altering, 5'/3' UTR, other exonic, or putative promoter, **Supplementary File 3**), and molecular (magenta; enriched molecular pathway, drug target) evidence. We here focus on the gene(s) with the highest gene priority score (GPS) per locus (full list of genes in **Supplementary File 9**). Shown are (a) the 16 genes with highest GPS in the 15 novel AMD loci (one novel locus without any gene), and (b) the 25 genes with highest GPS in the 18 known AMD loci. Colored fields indicate yes and GPS counts number of colored fields per row.

**Figure 3. Comparison of advanced AMD subtypes and intermediate versus advanced AMD.** We compared associations of the 34 lead variants across different AMD phenotypes. Shown are effect sizes (log Odds Ratio) per minor allele in controls as well as 95% confidence intervals (widths and heights of diamonds). (a) Comparison of neovascular disease (10,749 CNV cases vs. 17,832 controls) and GA (3,235 GA cases vs. 17,832 controls) identified four variants (in loci *MMP9*, *ARMS2/HTRA1*, *CETP*, and *SYN3/TIMP3*) with significantly different association comparing CNV with GA ( $P_{\text{diff}} < 0.05/34$ , marked in red,



see also **Supplementary Table 12**). **(b)** Comparison of intermediate AMD (6,657 cases vs. 17,832 controls) with advanced AMD (16,144 cases vs. 17,832 controls) identifies 24 variants with nominally significant ( $P < 0.05$ , marked in red) association with intermediate AMD ( $P_{\text{binomial}} = 4.8 \times 10^{-24}$ ), all of which have the same effect direction and less extreme effect sizes compared to advanced AMD (**Supplementary Table 13**).

**Figure 4. Variance explained and absolute risk of disease based on the 52 identified variants.** **(a)** Absolute disease risk (=proportion of affected) by genetic risk score intervals (deciles and top 10 percentiles in embedded bar plot) based on our cases-control-data weighted to model a general population with 5% disease prevalence (see also **Supplementary Table 16**). **(b)** Shown is disease liability explained by the 52 identified variants (bars) compared to the genomic heritability based on all genotyped variants (red lines) assuming disease prevalence of 1%, 5%, or 10%, respectively.

**Table 1. Thirty-four loci for age-related macular degeneration.** Our genome-wide single-variant association analysis identified 34 loci for advanced AMD with genome-wide significance ( $P < 5 \times 10^{-8}$ ) based on logistic regression in 16,144 cases and 17,832 controls of European ancestry. Shown are P-values and effect sizes (Odds Ratios, OR) for the variant with the smallest P-value per locus (lead variant) and the number of independent signals per locus (see **Supplementary Table 4**)

Lead Variant	Chr	Position <sup>a</sup>	Major/minor allele	Locus name <sup>b</sup>	# Sig-nals <sup>c</sup>	MAF		Association	
						Cases	Controls	OR	P
KNOWN (previously reported with genome-wide significance, $P < 5 \times 10^{-8}$ )									
rs10922109	1	196,704,632	C/A	<i>CFH</i>	8	0.223	0.426	0.38	$9.6 \times 10^{-618}$
rs62247658	3	64,715,155	T/C	<i>ADAMTS9-AS2</i>	1	0.466	0.433	1.14	$1.8 \times 10^{-14}$
rs140647181	3	99,180,668	T/C	<i>COL8A1</i>	2	0.023	0.016	1.59	$1.4 \times 10^{-11}$
rs10033900	4	110,659,067	C/T	<i>CFI</i>	2	0.511	0.477	1.15	$5.4 \times 10^{-17}$
rs62358361	5	39,327,888	G/T	<i>C9</i>	1	0.016	0.009	1.80	$1.3 \times 10^{-14}$
rs116503776	6	31,930,462	G/A	<i>C2/CFB/SKIV2L</i>	4	0.090	0.148	0.57	$1.2 \times 10^{-103}$
rs943080	6	43,826,627	T/C	<i>VEGFA</i>	1	0.465	0.497	0.88	$1.1 \times 10^{-14}$
rs79037040	8	23,082,971	T/G	<i>TNFRSF10A</i>	1	0.451	0.479	0.90	$4.5 \times 10^{-11}$
rs1626340	9	101,923,372	G/A	<i>TGFBR1</i>	1	0.189	0.209	0.88	$3.8 \times 10^{-10}$
rs3750846	10	124,215,565	T/C	<i>ARMS2/HTRA1</i>	1	0.436	0.208	2.81	$6.5 \times 10^{-735}$
rs9564692	13	31,821,240	C/T	<i>B3GALT1</i>	1	0.277	0.299	0.89	$3.3 \times 10^{-10}$
rs61985136	14	68,769,199	T/C	<i>RAD51B</i>	2	0.360	0.384	0.90	$1.6 \times 10^{-10}$
rs2043085	15	58,680,954	T/C	<i>LIPC</i>	2	0.350	0.381	0.87	$4.3 \times 10^{-15}$
rs5817082	16	56,997,349	C/CA	<i>CETP</i>	2	0.232	0.264	0.84	$3.6 \times 10^{-19}$
rs2230199	19	6,718,387	C/G	<i>C3</i>	3	0.266	0.208	1.43	$3.8 \times 10^{-69}$
rs429358	19	45,411,941	T/C	<i>APOE</i>	2	0.099	0.135	0.70	$2.4 \times 10^{-42}$
rs5754227	22	33,105,817	T/C	<i>SYN3/TIMP3</i>	1	0.109	0.137	0.77	$1.1 \times 10^{-24}$
rs8135665	22	38,476,276	C/T	<i>SLC16A8</i>	1	0.217	0.195	1.14	$5.5 \times 10^{-11}$
NOVEL (reported with genome-wide significance, $P < 5 \times 10^{-8}$ , for the first time)									
rs11884770	2	228,086,920	C/T	<i>COL4A3</i>	1	0.258	0.278	0.90	$2.9 \times 10^{-8}$
rs114092250	5	35,494,448	G/A	<i>PRLR/SPEF2</i>	1	0.016	0.022	0.70	$2.1 \times 10^{-8}$
rs7803454	7	99,991,548	C/T	<i>PILRB/PILRA</i>	1	0.209	0.190	1.13	$4.8 \times 10^{-9}$
rs1142	7	104,756,326	C/T	<i>KMT2E/SRPK2</i>	1	0.370	0.346	1.11	$1.4 \times 10^{-9}$
rs71507014	9	73,438,605	GC/G	<i>TRPM3</i>	1	0.427	0.405	1.10	$3.0 \times 10^{-8}$
rs10781182	9	76,617,720	G/T	<i>MIR6130/RORB</i>	1	0.328	0.306	1.11	$2.6 \times 10^{-9}$
rs2740488	9	107,661,742	A/C	<i>ABCA1</i>	1	0.255	0.275	0.90	$1.2 \times 10^{-8}$
rs12357257	10	24,999,593	G/A	<i>ARHGAP21</i>	1	0.243	0.223	1.11	$4.4 \times 10^{-8}$
rs3138141	12	56,115,778	C/A	<i>RDH5/CD63</i>	1	0.222	0.207	1.16	$4.3 \times 10^{-9}$
rs61941274	12	112,132,610	G/A	<i>ACAD10</i>	1	0.024	0.018	1.51	$1.1 \times 10^{-9}$
rs72802342	16	75,234,872	C/A	<i>CTRB2/CTRB1</i>	1	0.067	0.080	0.79	$5.0 \times 10^{-12}$
rs11080055	17	26,649,724	C/A	<i>TMEM97/VTN</i>	1	0.463	0.486	0.91	$1.0 \times 10^{-8}$
rs6565597	17	79,526,821	C/T	<i>NPLOC4/TSPAN10</i>	1	0.400	0.381	1.13	$1.5 \times 10^{-11}$
rs67538026	19	1,031,438	C/T	<i>CNN2</i>	1	0.460	0.498	0.90	$2.6 \times 10^{-8}$
rs142450006	20	44,614,991	TTTTTC/T	<i>MMP9</i>	1	0.124	0.141	0.85	$2.4 \times 10^{-10}$
rs201459901	20	56,653,724	T/TA	<i>C20orf85</i>	1	0.054	0.070	0.76	$3.1 \times 10^{-16}$

Chr = Chromosome; MAF = minor allele frequency; OR = Odds Ratio <sup>a</sup> Chromosomal position is given based on NCBI RefSeq hg19; <sup>b</sup> The locus name is a label of the region using the nearest gene(s), but does not necessarily state the responsible gene; <sup>c</sup> number of independent variants in this locus; hg19 = human genome reference assembly (version 19)

**Table 2. Four genes with a significant rare variant burden within the 34 AMD loci independent from other identified variants.** We computed a gene-based burden test of rare protein-altering variants comparing 16,144 advanced AMD cases and 17,832 controls. Shown are P-values from the variable threshold test (up to 100 million permutations) and Odds Ratios from the collapsed burden test, both adjusted for the other identified variants in the respective locus (locus-wide conditioning). Four genes (among the 703 genes in the 34 AMD locus regions) showed a significant ( $P < 0.05/703 = 7.1 \times 10^{-5}$ ) burden. Details about the corresponding rare variants underlying the observed burden can be found in **Supplementary File 4**. Results for the 14 genes that show significant burden within the 34 AMD loci without locus-wide conditioning are shown in **Supplementary Table 9**. Rare variants were defined here as variants with minor allele frequency in cases and controls  $< 1\%$  in each of the ancestries, European, Asian, and African.

Gene	Optimal Threshold for Rare Variants Count (%)	Number of Variants below Optimal RAC Total (Exome Chip Base + Custom)	Summed Rare Allele Count (Frequency [%])		$P^a$	Odds Ratio
			Cases N = 16,144	Controls N = 17,832		
			<i>CFH</i>	10 (0.015%)		
<i>CFI</i>	46 (0.068%)	43 (17+26)	213 (0.660%)	82 (0.230%)	$1.0 \times 10^{-8}$	2.95
<i>TIMP3</i>	14 (0.021%)	9 (1+8)	29 (0.0898%)	1 (0.00280%)	$9.0 \times 10^{-8}$	31.21
<i>SLC16A8</i>	648 (0.954%)	9 (7+2)	487 (1.51%)	392 (1.10%)	$3.1 \times 10^{-6}$	1.40

RAC = rare allele count; <sup>a</sup> P-values are from the variable threshold test conditioned on other identified variants in the locus (locus-wide conditioned).

## ONLINE METHODS

**Study data and phenotype.** In the International AMD Genomics Consortium (IAMGDC), we gathered 26 studies with each including (i) advanced AMD cases with GA and/or CNV in at least one eye and age at first diagnosis  $\geq 50$  years, (ii) intermediate AMD cases with pigmentary changes in the RPE or more than five macular drusen greater than  $63\mu\text{m}$  and age at first diagnosis  $\geq 50$  years, or (iii) controls without known advanced or intermediate AMD. Recruitment and ascertainment strategies varied by study (**Supplementary Tables 1 and 2**): Advanced and intermediate AMD cases were mostly recruited from ophthalmology clinics (61.6% of cases), but also in spouses and friends of cases (2.1%), from general population (18.5%), or via mixed approaches (17.8%); controls were recruited among elderly individuals at ophthalmology clinics (53.0% of controls), among spouses and friends of cases (2.6%), from general population (26.4%), or via mixed approaches (18.0%). Of all subjects, 94.5% ascertained disease status via fundus photography or fundus exam; one study (5.5% of subjects) validated interview information through the patients' ophthalmologist. Of the 26 studies, 18 studies used Fluorescein Angiography or Optic Coherence Tomography for differentiating GA from neovascular disease. Grading scales, used to ascertain intermediate AMD, differed – as usual – across studies. All groups collected data according to the Declaration of Helsinki principles. Study participants provided informed consent and protocols were reviewed and approved by local ethics committees.

**DNA and chip design.** We gathered DNA samples of more than 50,000 individuals. Groups with very limited amounts of available DNA contributed aliquots after whole-genome amplification (8% of subjects).

We utilized a custom-modified HumanCoreExome array by Illumina, Inc., which includes (i) tagging variants across the genome (genome chip content) and (ii) a catalogue of protein-altering variants (exome chip content). Our customization of the array included three additional tiers to enrich for variants from 22 AMD loci implicated by our previous genome-wide association analysis<sup>6</sup> based on 19 index variants with genome-wide significance, 3 with consistent effect direction in the replication stage and  $4 \times 10^{-7} \leq P \leq 2 \times 10^{-6}$ ) by selecting (iii) tagging variants (pair-wise tagging  $r^2 < 0.8$ ) from Phase I 1000G/HapMap<sup>53,54</sup> common variants (minor allele frequency, MAF,  $\geq 1\%$  in European or East Asian individuals) using Tagger implemented in Haploview<sup>55</sup> within  $\pm 100\text{kb}$  of the 22 index variants expanded to cover all correlated variants ( $r^2 [\text{EUR}] > 0.5$ ) and the complete gene (transcript  $\pm 1\text{ kb}$ ), (iv) protein-altering variants within 500 kb of the 22 index variants as identified from public general population data bases (dbSNP<sup>56</sup>, the NHLBI Exome Sequencing Project<sup>57</sup>, the Phase I 1000 Genomes Project, see **Web Resources**), and (v) protein-altering variants within the 500 kb of the 22 index variants identified by re-sequencing AMD case-control study data

(targeted re-sequencing of 2,335 AMD cases and 789 controls<sup>10,58</sup> and whole-genome sequencing 60 AMD cases and 60 controls; G. Abecasis and A. Swaroop). The customization further included (vi) the 1,000 top independent (> 2 Mb distant) variants from the previous analysis and additional 100 top variants from each the previous CNV only and the previous GA only analysis, (vii) and 375 variants in *ABCA4*, including known variants causing Stargardt disease<sup>59</sup>, benign variants, and those of unknown significance, as well as 10 known and 44 predicted cysteine mutations in *TIMP3*, motivated by the known variants causing Sorsby's fundus dystrophy<sup>33,34</sup> (also B. Weber, personal communication).

**Annotation.** Variant identifiers were based on NCBI dbSNP v137. Chromosomal position and functional annotation of the variant was based on the NCBI Reference Sequence Human Genome Build 19 (RefSeq hg19)<sup>60</sup> and SeattleSeq Annotation 138<sup>61</sup> (see **Web Resources**). We particularly focus on protein-altering variants including non-synonymous coding variants (missense, stop loss, in-frame insertion/deletion, frameshift, premature stop codon) and splice sites. We converted the description of splice site variants to HGVS nomenclature using Mutalyzer version 2.0.beta-33<sup>62</sup> (see **Web Resources**).

**Genotypes.** We genotyped all subjects centrally at the Center for Inherited Diseases Research (CIDR), Johns Hopkins University School of Medicine, Baltimore, MD, USA. From the 569,645 genotyped variants, our stringent quality control procedure excluded poorly genotyped variants as evidenced by genotype call rates < 98.5% (5.8%), deviations from Hardy-Weinberg equilibrium with  $P < 10^{-6}$  (0.34%), variants that mapped at multiple genome locations (0.25%) or variants failing other criteria, resulting in 521,950 (91.6%) variants passing all quality criteria. After excluding monomorphic variants (15.8%), we yielded 264,655 common variants distributed across autosomes, sex chromosomes, and the mitochondria, as well as 163,714 directly genotyped protein-altering variants including 8,290 from previously implicated AMD loci (**Supplementary Table 3A**). For these variants, genotype call rates averaged 99.9% (99.1% for subjects with amplified DNA).

We phased the autosomal and X-chromosomal genotype data using SHAPEIT (200 states, 2.5 Mb windows)<sup>63</sup>, then imputed genotypes based on the 1000 Genomes Project<sup>64</sup> reference panel (1000G Phase I, version 3, SHAPEIT2 Reference) using MINIMAC<sup>65</sup> (reference-based 2.5 Mb chunks, 500 kb buffer regions). We then merged study variants that were excluded during imputation (not found in the reference panel) back into the final data set. We excluded common variants ( $CAF \geq 1\%$ ) with bad imputation quality,  $R^2 < 0.3$ , and adopted a more stringent exclusion criterion for rare variants ( $CAF < 1\%$ ,  $R^2 < 0.8$ , for the initial identification of lead variants. This yielded a total of 12,023,830 genotyped (439,350) or imputed (11,584,480) quality-controlled variants (**Supplementary Table 3A**).

**Analyzed subjects.** Using the genomic information for subject-level quality control, we excluded duplicated and related individuals (kinship coefficient  $\phi \geq 0.0884$ , i.e. 3<sup>rd</sup> degree relatives or closer)<sup>66</sup>, subjects with discrepancies between reported gender and sex chromosomal information or with atypical sex chromosome configurations<sup>67</sup>, or subjects with genotyping call rates < 98.5%; we derived ancestry based on the first two principal components using autosomal genotyped variants together with genotype information of the samples from the Human Genome Diversity Project (HGDP)<sup>68</sup>. Our final data set contained 43,566 successfully genotyped unrelated subjects including 16,144 advanced AMD cases and 17,832 controls of European ancestry, 6,657 intermediate AMD cases of European ancestry, and 2,933 subjects (advanced AMD or controls) of Asian or African ancestries (**Supplementary Table 3B**).

**Genomic heritability and genomic correlation.** Combined contribution of genotyped variants to disease was evaluated using a variance-component based heritability analysis<sup>69</sup>. This analysis used genotypes to build a similarity matrix, summarizing the overall genetic kinship between each pair of individuals, and then examined the correspondence between genetic and phenotypic similarity. We estimated the explained variance on all genotyped, autosomal variants using restricted maximum likelihood (REML) analysis implemented in GCTA<sup>28</sup> (see **Web Resources**). We jointly estimated the contributions of rare (MAF in controls < 1 %) and common (MAF in controls  $\geq 1\%$ ) genotyped variants by first separately calculating their genetic relationship matrices before adding both to the model. Obtained estimates of variance explained were transformed from the observed scale to the liability scale assuming various levels of disease prevalence<sup>69</sup>.

We estimated the genomic correlation between different disease sub-phenotypes using bivariate REML analyses implemented in GCTA and only included common (MAF in controls  $\geq 1\%$ ) genotyped variants<sup>29</sup>. We compared 10,749 cases with CNV versus 3,325 cases with GA (excluding the 2,070 cases with mixed CNV and GA) and we compared 6,657 intermediate AMD cases with 16,144 advanced AMD cases. For both analyses, we used the control subjects as reference and avoided shared controls between traits by randomly splitting the 17,832 unrelated European control individuals into two sub-samples of 8,916 individuals.

**Genome-wide single variant association analysis.** Single-variant association tests analyzing the 16,144 advanced AMD cases and 17,832 controls of European ancestry were based on the Firth bias-corrected likelihood ratio test<sup>70</sup>, which is recommended for genetic association studies that include rare variants<sup>71</sup>, as implemented in EPACTS (see **Web**

**Resources).** Analyses were adjusted for two principal components and source of DNA (whole-blood or whole-genome amplified DNA). Allele dosages of the imputed data were utilized, Sensitivity analyses were conducted to evaluate the influence of alternative association tests, alternative covariate adjustment including age or sex, or up to 10 principal components instead of two, as well as the influence of restricting to population-based controls, or to controls aged 50 years or older. Genomic control correction<sup>72</sup> was used to account for potential population stratification using all genotyped variants with minor allele count  $\geq 20$  outside of 20 previously described AMD loci<sup>6,9</sup>. As usual for genome-wide association studies, we considered P-values  $\leq 5 \times 10^{-8}$  as genome-wide significant.

To identify independently associated variants, we adopted a sequential forward selection approach: We first computed single variant association for each of the  $> 12$  million variants. Then we selected the variant with the smallest P-value and its flanking  $\pm 5$  Mb region, repeating the process until no genome-wide significant variant ( $P \leq 5 \times 10^{-8}$ ) was left yielding a number of 10 Mb regions. Within each of these large regions, we re-analyzed each variant conditioning on the top variant, and repeated this process by adding the previously identified genome-wide significant variant(s) within the respective 10 Mb region. This yielded one or more independently associated genome-wide significant variant(s) per 10 Mb region.

A locus region was defined by a genome-wide significant variant and its correlated variants ( $r^2 \geq 0.5$ )  $\pm 500$ kb; overlapping locus regions were merged to one locus, so some loci contained more than one index variant (details in **Supplementary Figure 3**).

In order to derive independent effect sizes (log odds ratios) for all identified variants, we computed a fully conditioned logistic regression model including all identified variants.

**Bayesian approach to prioritize variants.** In order to summarize the statistical evidence of a variant for its association strength, we computed the Bayes factor for each variant, which is a measure of the strength of the association that is comparable irrespective of variant frequency or study sample size. It provides the probability of the genotype configuration at a variant (in cases and controls) under the alternative hypothesis (association) divided by the probability of the genotype configuration under the null hypothesis (no association). It is computed using the association results per variant<sup>73</sup>. The posterior probability of each variant is then computed as the Bayes factor relative to the sum of all variants' Bayes factors across one locus region and can be thought of as the relative strength of evidence in favor of each SNP studied in the respective region. This assumes that there is one causal variant per region and that the causal variant is in the analyzed data set.

Expanding to loci with multiple association signals and thus a single alleged causal variant per signal, we used the association results per SNP obtained by conditioning on the other independent variants at that locus for computing the Bayes factor.

We derived 95% credible sets of variants per signal, which is the minimal set of variants, for which the sum of the posterior probabilities accumulates beyond 95%. This approach was recommended for fine-mapping of association signals and for prioritizing variants<sup>74</sup>. Assuming that there is only one causal variant in an association signal and that the causal variant is contained among the analyzed variants, such a credible set of variants contains the causal variant with 95% probability.

We annotated functionality of the variants in each of the 95% credible sets (see above).

**Gene-based burden analysis.** Single variant analyses have limited power to depict rare variants with association. Gene-based burden tests evaluating accumulated association from multiple rare variants per gene have been shown to complement such analyses and improve power to detect a burden of disease. We computed the burden of disease using the variable threshold test<sup>52</sup> as implemented in EFACTS. These analysis assume that all variants in a gene either increase or decrease disease risk. When variants with opposite directions of effect reside in the same gene, power will be reduced. An analysis with SKAT and SKAT-O, which both allow for variants with opposite directions of effect to reside in the same gene, did not identify additional signals (data not shown).

We focused this analysis on protein-altering variants, since we assumed that the other (not protein-altering) variants would outnumber these predicted deleterious variants by far and would thus dilute a disease burden from the deleterious variants. Assuming a negative selection against such deleterious variants that cause their frequency to be low across ancestries, we restricted our rare variant definition to variants with MAF < 1% (cases and controls combined) in each of our ancestry groups (African, Asian, and European). We utilized the genotypes of these rare protein-altering variants if genotyped directly, or rounded imputed allele dosages to the next best genotype if imputed; imputed variants were restricted to those of highest imputation quality (RSQ  $\geq$  0.8).

We assessed statistical significance by adaptive permutation testing with variable thresholds (up to 100 million permutations; minimal P-value =  $1 \times 10^{-8}$ )<sup>52</sup>. When rare variants appear on a haplotype associated with disease through a common variant allele already identified for AMD, the rare variant burden would depict a mere shadow of the already identified variant. Therefore, we repeated the variable threshold test conditioned on the variant(s) identified in the respective locus by single variant analysis (locus-wide conditioning), to unravel a gene-based burden of rare variants independent of risk variants identified in single variants tests.

First, we searched for rare variant disease burden genome-wide applying a genome-wide Bonferroni-corrected significance threshold of  $0.05 / 17,044 = 2.9 \times 10^{-6}$  (17,044 genes



genome-wide with at least 1 variant included in the analysis, i.e. with  $\geq 1$  rare protein-altering variant). In a second view on this, we focused on our 34 identified AMD loci and here applied a significance threshold based on the 703 genes overlapping with the locus regions ( $P < 0.05 / 703 = 7.1 \times 10^{-5}$ ). Odds ratio estimates of the burden were derived by logistic regression using the Wald test on the collapsed burden.

There was an overlap of the sequenced subjects with the chip data subjects: of the 3264 subjects in the overlap, 3084 had passed our quality control including 2959 unrelated subjects of European ancestry with either late AMD (858), early AMD (1451), or no AMD (650). We conducted a sensitivity analysis for the burden test excluding the 858 advanced AMD subjects and the 650 control subjects (thus comparing 15,286 advanced AMD subjects to 17,182 control subjects).

**Follow-up queries for genes underneath the association signals.** In order to derive information for all genes underneath our 52 identified association signals (spread across the 34 AMD loci), we built a gene list containing all genes that overlapped with a more narrow definition of locus regions: We have been using a particularly comprehensive definition of the locus region during the signal identification step (index variants and proxies,  $r^2 \geq 0.5$ ,  $\pm 500\text{kb}$ ), to avoid far-reaching linkage disequilibrium that may generate shadow signals (particularly in the light of strong associations in the *CFH*, *C3*, *C2/CFI*, and *ARMS2/HTRA1* loci) and to optimally differentiate independent signals within a locus. We have also used this wide locus region definition for the rare variant burden test again to fully correct for independent signals in the respective wider locus regions and to be conservative in the multiple testing corrections for the AMD-locus-wide burden test search. However, this wide definition is less adequate when prioritizing genes around the identified signals under the assumption that most protein-altering or regulating variants exert their effects in cis<sup>42</sup>. We thus focused the gene list for further queries to a more narrow locus region definition (index variants and proxies,  $r^2 \geq 0.5$ ,  $\pm 100\text{kb}$ ) and yield 368 overlapping RefSeq genes (**Supplementary File 5**).

**Gene expression.** For the 368 genes in our gene list (see above), we sought to obtain gene expression in relevant tissues, retina, RPE, and choroid, in two independent data sets.

In the first laboratory (Dwight Stambolian Lab; University of Pennsylvania), we used RNA-Seq to characterize the chorioretinal transcriptomes in a discovery set of eight normal human eyes (two eyes from each of four persons)<sup>75</sup>. For each eye, we sequenced four RNA-Seq samples and generated close to 100 million 101-bp paired-end reads per sample. We mapped the sequence reads to the reference human genome (hg19) using GSNAP<sup>76</sup>. Our data are of high quality with 76–94% of the reads mapped to the human genome and 60–

81% retained after stringent quality control filtering, among which 86–93% mapped to genes defined by RefSeq. We considered the overall gene expression in each sample. Using filtered mapped reads, we estimated the expression levels of 23,569 RefSeq protein-coding genes using the fragments per kilobase of gene per million mapped fragment (FPKM) metric<sup>77</sup>. With coverage depth ranging from 66 to 133 million paired-end reads per sample, we detected expression of the majority of known protein-coding genes<sup>75</sup>. We considered genes and transcripts to be expressed, if FPKM > 0. Among the 23,569 genes with available expression for retina tissue or RPE/choroid/sclera, respectively, 290 genes for retina and 300 for RPE/choroid/sclera tissue overlapped with the 368 genes in the gene list.

In the second independent laboratory (Weber lab; University of Regensburg), we used RNA-Seq to estimate the relative abundance of known and novel transcripts in human retina, RPE and RPE-related cell types. Each tissue/cell line was sequenced as biological replicate, e.g. RNA was retrieved from cells of two individuals. The NextFlex Directional RNASeq library preparation kit (UDP based) from Bioo was used and between 30 and 60 million 75bp paired-end reads for each library were generated. The Tuxedo Tools pipeline (BowTie, TopHat, and Cufflinks) was used to map the reads to the genome and transcriptome and to quantify the abundance of transcripts measured as fragments per kilobase of gene per million mapped fragments (FPKM)<sup>77</sup>. We considered genes and transcripts to be expressed, if the respective FPKM value of the gene/transcript was greater than the first quartile of all FPKM values obtained from the tissues. Among the 20,590 with available expression available in retina tissue or RPE/choroid, respectively, 316 genes for retina and for REP/choroid overlapped with the 368 genes in gene list.

A consensus rating of gene expression observed in the two labs was derived as follows: Expression of a gene in one set of tissues (retina or RPE/choroid) was inferred, if both labs detected expression in the respective set of tissues; if at least one of the labs did not observe expression, the gene was considered as not expressed; gene expression of all other genes (one lab observing expression and the other with missing, or both labs with missing data) was regarded as missing.

**Mouse model phenotypes.** For the 368 genes in our gene list, we queried the Mouse Genome Informatics (MGI)<sup>78</sup> and the International Mouse Phenotyping Consortium (IMPC)<sup>79</sup> data bases (see **Web Resources**), and manually curated results by information from published literature. We determined whether a gene exhibited a relevant eye-phenotype (i.e. retina, RPE, or choroid phenotypes) in established genetic mouse models (knock-out, knock-in, or trans-genic mice).

**Enrichment for molecular pathways.** For the 368 overlapping genes, we performed functional enrichment analysis using INRICH<sup>80</sup> with default settings unless stated otherwise. Target intervals of this analysis were the narrow AMD locus regions (index variants and proxies,  $r^2 \geq 0.5$ ,  $\pm 100\text{kb}$ , **Supplementary Table 5**). Since there is no consensus approach to pathway analysis, we queried multiple data bases: (i) Kyoto Encyclopedia of Genes and Genomes (KEGG)<sup>81</sup>, (ii) Reactome<sup>82</sup>, and (iii) Gene Ontology (GO) Consortium<sup>83</sup> (see **Web Resources**). For example, while KEGG is a manually curated database on metabolic pathways, GO also includes automatic annotations and more comprehensive set of cellular processes and molecular functions. To reduce the multiple testing burden, we used gene sets with 5 to 200 genes that overlapped at least three overlapping target intervals. All our imputed/genotyped common (MAF in cases and controls combined  $\geq 1\%$ ) variants genome-wide in these target regions were used to inform this analysis regarding variant density; no P-value threshold was used. We carried out the analysis with 1,000,000 replicates and 50,000 bootstrap rounds to yield corrected P-values, matching selected target regions in terms of gene count, variant density (80-120%) and total number of variants.

**Drug pathways and targets.** In order to derive information on whether the product of a gene among the 368 genes in our gene list was a direct drug target, we searched the DrugBank database (Version 4.1) which contains 4,207 drug targets (= genes) and 7,740 drugs<sup>84</sup> (see **Web Resources**).

**Explained variability in disease liability.** Based on the 52 identified AMD variants, we estimated the explained proportion of disease liability explained by these variants (see **Web Resources**)<sup>85</sup> using the log Odds Ratio estimates from the model including all 52 identified variants (fully conditioned) to derive independent effect sizes. We compared this proportion explained by the 52 variants with the earlier derived genomic heritability based on all genotyped variants (see above).

**Genetic risk score and relative and absolute genetic risk of AMD.** For each individual, we computed a genetic risk score (GRS) as the effect size weighted sum of the AMD risk increasing alleles for all 52 independent variants divided by the sum of all effect sizes. For the weighting, the log Odds Ratios for each of the 52 variants were derived from the fully adjusted model (including all 52 variants), to assure independence of effect sizes.

In order to also derive a realistic genetic risk score distribution, we modeled a general population based on our case-control data by weighing each case individual using

$$w_{\text{case}} = \textit{Prevalence} / (N_{\text{cases}} / (N_{\text{cases}} + N_{\text{controls}}))$$

and each control individual using

$$w_{\text{control}} = (1 - \text{Prevalence}) / (N_{\text{controls}} / (N_{\text{cases}} + N_{\text{controls}})),$$

with *Prevalence* being an assumed prevalence of advanced AMD in the general population. We computed several scenarios using prevalence estimates of 1%, 5%, or 10% reflecting approximate prevalence of advanced AMD in the general population above the age of 50, 75, or 85 years of age, respectively. For this modeled general population, we derived the GRS distribution and its deciles.

We derived relative risk estimates (as Odds Ratios) for each GRS decile with the first decile as reference. This relative risk estimate per se is independent of the prevalence except that the decile to form the genetic risk groups used the GRS distribution as expected in a general population (which requires a prevalence assumption). We also computed absolute risk estimates per GRS decile, which is given by the proportion of advanced AMD cases applying the weights, again, as described above. This estimate depends on the prevalence.

## References - Online Methods:

53. The International HapMap Consortium *et al.* A second generation human haplotype map of over 3.1 million SNPs. *Nature* **449**, 851-61 (2007).
54. 1000 Genomes Project Consortium *et al.* An integrated map of genetic variation from 1,092 human genomes. *Nature* **491**, 56-65 (2012).
55. Barrett, J.C., Fry, B., Maller, J. & Daly, M.J. Haploview: analysis and visualization of LD and haplotype maps. *Bioinformatics* **21**, 263-5 (2005).
56. Sherry, S.T. *et al.* dbSNP: the NCBI database of genetic variation. *Nucleic Acids Res* **29**, 308-11 (2001).
57. Tennessen, J.A. *et al.* Evolution and functional impact of rare coding variation from deep sequencing of human exomes. *Science* **337**, 64-9 (2012).
58. Age-Related Eye Disease Study Research Group. A randomized, placebo-controlled, clinical trial of high-dose supplementation with vitamins C and E and beta carotene for age-related cataract and vision loss: AREDS report no. 9. *Arch Ophthalmol* **119**, 1439-52 (2001).
59. Fritsche, L.G. *et al.* A subgroup of age-related macular degeneration is associated with mono-allelic sequence variants in the ABCA4 gene. *Invest Ophthalmol Vis Sci* **53**, 2112-8 (2012).
60. Pruitt, K.D. *et al.* RefSeq: an update on mammalian reference sequences. *Nucleic Acids Res* **42**, D756-63 (2014).
61. Ng, S.B. *et al.* Targeted capture and massively parallel sequencing of 12 human exomes. *Nature* **461**, 272-6 (2009).
62. Wildeman, M., van Ophuizen, E., den Dunnen, J.T. & Taschner, P.E. Improving sequence variant descriptions in mutation databases and literature using the Mutalyzer sequence variation nomenclature checker. *Hum Mutat* **29**, 6-13 (2008).
63. Delaneau, O., Marchini, J. & Zagury, J.F. A linear complexity phasing method for thousands of genomes. *Nat Methods* **9**, 179-81 (2012).
64. Ristau, T. *et al.* Allergy is a protective factor against age-related macular degeneration. *Invest Ophthalmol Vis Sci* **55**, 210-4 (2014).
65. Howie, B., Fuchsberger, C., Stephens, M., Marchini, J. & Abecasis, G.R. Fast and accurate genotype imputation in genome-wide association studies through pre-phasing. *Nat Genet* **44**, 955-9 (2012).
66. Manichaikul, A. *et al.* Robust relationship inference in genome-wide association studies. *Bioinformatics* **26**, 2867-73 (2010).
67. Turner, S. *et al.* Quality control procedures for genome-wide association studies. *Curr Protoc Hum Genet* **Chapter 1**, Unit1 19 (2011).

68. Cavalli-Sforza, L.L. The Human Genome Diversity Project: past, present and future. *Nat Rev Genet* **6**, 333-40 (2005).
69. Lee, S.H., Wray, N.R., Goddard, M.E. & Visscher, P.M. Estimating missing heritability for disease from genome-wide association studies. *Am J Hum Genet* **88**, 294-305 (2011).
70. Firth, D. Bias reduction of maximum likelihood estimates. *Biometrika* **80**, 27-38 (1993).
71. Ma, C., Blackwell, T., Boehnke, M., Scott, L.J. & Go, T.D.i. Recommended joint and meta-analysis strategies for case-control association testing of single low-count variants. *Genet Epidemiol* **37**, 539-50 (2013).
72. Devlin, B. & Roeder, K. Genomic control for association studies. *Biometrics* **55**, 997-1004 (1999).
73. Stephens, M. & Balding, D.J. Bayesian statistical methods for genetic association studies. *Nat Rev Genet* **10**, 681-90 (2009).
74. Wellcome Trust Case Control Consortium *et al.* Bayesian refinement of association signals for 14 loci in 3 common diseases. *Nat Genet* **44**, 1294-301 (2012).
75. Li, M. *et al.* Comprehensive analysis of gene expression in human retina and supporting tissues. *Hum Mol Genet* **23**, 4001-14 (2014).
76. Wu, T.D. & Nacu, S. Fast and SNP-tolerant detection of complex variants and splicing in short reads. *Bioinformatics* **26**, 873-81 (2010).
77. Trapnell, C. *et al.* Transcript assembly and quantification by RNA-Seq reveals unannotated transcripts and isoform switching during cell differentiation. *Nat Biotechnol* **28**, 511-5 (2010).
78. Blake, J.A., Bult, C.J., Eppig, J.T., Kadin, J.A. & Richardson, J.E. The Mouse Genome Database: integration of and access to knowledge about the laboratory mouse. *Nucleic Acids Res* **42**, D810-7 (2014).
79. Brown, S.D. & Moore, M.W. Towards an encyclopaedia of mammalian gene function: the International Mouse Phenotyping Consortium. *Dis Model Mech* **5**, 289-92 (2012).
80. Lee, P.H., O'Dushlaine, C., Thomas, B. & Purcell, S.M. INRICH: interval-based enrichment analysis for genome-wide association studies. *Bioinformatics* **28**, 1797-9 (2012).
81. Kanehisa, M. & Goto, S. KEGG: kyoto encyclopedia of genes and genomes. *Nucleic Acids Res* **28**, 27-30 (2000).
82. Croft, D. *et al.* The Reactome pathway knowledgebase. *Nucleic Acids Res* **42**, D472-7 (2014).
83. Ashburner, M. *et al.* Gene ontology: tool for the unification of biology. The Gene Ontology Consortium. *Nat Genet* **25**, 25-9 (2000).
84. Law, V. *et al.* DrugBank 4.0: shedding new light on drug metabolism. *Nucleic Acids Res* **42**, D1091-7 (2014).
85. So, H.C., Gui, A.H., Cherny, S.S. & Sham, P.C. Evaluating the heritability explained by known susceptibility variants: a survey of ten complex diseases. *Genet Epidemiol* **35**, 310-7 (2011).

## Web Resources:

**Full GWAS results:** <http://csg.sph.umich.edu/abecasis/public/amd2015/>

The following Web Resources have been utilized:

GWAS catalog <http://www.ebi.ac.uk/gwas/home>),

Exome Variant Server, NHLBI GO Exome Sequencing Project:

<http://evs.gs.washington.edu/EVS/>

EPACTS: <http://www.sph.umich.edu/csg/kang/epacts/index.html>

SHAPEIT: [https://mathgen.stats.ox.ac.uk/genetics\\_software/shapeit/shapeit.html](https://mathgen.stats.ox.ac.uk/genetics_software/shapeit/shapeit.html)

MINIMAC: <http://genome.sph.umich.edu/wiki/Minimac>

1000 Genomes Reference Panel:

<http://www.sph.umich.edu/csg/abecasis/MACH/download/1000G.2013-09.html>

The Human Genome Diversity Project data:

<http://genome.sph.umich.edu/wiki/LASER> and <http://www.hagsc.org/hgdp>

SeattleSeq: <http://snp.gs.washington.edu/SeattleSeqAnnotation138/index.jsp>

Mutalyzer: <https://mutalyzer.nl>

NCBI Reference Sequence (RefSeq, downloaded December, 2012):

<http://www.ncbi.nlm.nih.gov/refseq/>

Human Splicing Finder 3.0: <http://www.umd.be/HSF3/index.html>

PubMed (retrieved November 11, 2014): <http://www.pubmed.org>

Mouse Genome Informatics (MGI) databases: <http://www.informatics.jax.org>

International Mouse Phenotyping Consortium Database: <https://www.mousephenotype.org>

INRICH: <http://atqu.mgh.harvard.edu/inrich/>

KEGG: Kyoto Encyclopedia of Genes and Genomes (KEGG): <http://www.genome.jp/kegg/>

MSigDB database v4.0: <http://www.broadinstitute.org/gsea/index.jsp>

Reactome (downloaded January 12th, 2015): <http://www.reactome.org>

Gene Ontology (GO) Consortium (downloaded January 12th, 2015): <http://geneontology.org>

DrugBank (downloaded June 4, 2014): <http://www.drugbank.ca>

GCTA: <http://www.complextraitgenomics.com/software/gcta/>

Variance explained by genetic variants:

<https://sites.google.com/site/honcheongso/software/varexp>