

1 **Climate modulates the effects of tree diversity on forest productivity**

2 Tommaso Jucker<sup>1\*</sup>, Daniel Avăcăriței<sup>2</sup>, Ionuț Bărnoaiea<sup>2</sup>, Gabriel Duduman<sup>2</sup>, Olivier  
3 Bouriaud<sup>2</sup> and David A. Coomes<sup>1</sup>

4 <sup>1</sup>Forest Ecology and Conservation, Department of Plant Sciences, University of Cambridge,  
5 Downing Street, Cambridge, CB2 3EA, UK.

6 <sup>2</sup>Forestry Faculty, University Stefan cel Mare of Suceava, 13 Strada Universității, 720229,  
7 Suceava, Romania.

8 \***Correspondence author.** E-mail: [tj272@cam.ac.uk](mailto:tj272@cam.ac.uk)

9 **Running headline:** Biodiversity and wood production in forests

10

## 11 **Summary**

- 12 **1.** Despite growing evidence that, on average, diverse forests tend to be more productive  
13 than species poor ones, individual studies often report strongly contrasting relationships  
14 between tree species richness and above-ground wood production (AWP). In the attempt  
15 to reconcile these apparently inconsistent results, we explored whether the strength and  
16 shape of AWP – diversity relationships shifts along spatial and temporal environmental  
17 gradients in forests across Europe.
- 18 **2.** We used tree ring data from a network of permanent forest plots distributed at six sites  
19 across Europe to estimate annual AWP over a 15 year period (1997–2011). We then  
20 tested whether the relationship between tree species richness and AWP changes (i) across  
21 sites as a function of large-scale gradients in climatic productivity and tree packing  
22 density, and (ii) among years within each sites as a result of fluctuating climatic  
23 conditions.
- 24 **3.** AWP – species richness relationships varied markedly among sites. As predicted by  
25 theory, the relationship shifted from strongly positive at sites where climate imposed a  
26 strong limitation on wood production and tree packing densities were low, to weakly  
27 negative at sites where climatic conditions for growth were most suitable. In contrast, we  
28 found no consistent effect of interannual fluctuations in climate on the strength of AWP  
29 – species richness relationships within sites.
- 30 **4.** *Synthesis.* Our results indicate that the shape and strength of the relationship between tree  
31 diversity and forest productivity depends critically on environmental context. Across  
32 Europe, tree diversity shows the greatest potential to positively influence forest

33 productivity at either end of the latitudinal gradient, where adverse climatic conditions  
34 limit productivity and lead to the development of less densely packed stands.  
35 **Key-words:** above-ground wood production; biodiversity – ecosystem functioning; context  
36 dependency; FunDivEUROPE project; plant-climate interactions; stress gradient hypothesis;  
37 tree packing density; tree ring data

## 38 **Introduction**

39 As evidence that biodiversity promotes key ecosystem functions and services continues to  
40 amass (Cardinale *et al.* 2012), the argument for conserving biodiversity is increasingly being  
41 framed in terms of developing natural capital (Naeem, Duffy & Zavaleta 2012). In the  
42 context of forests, the consensus is that diverse stands are generally more productive than  
43 species-poor ones (Paquette & Messier 2011; Morin *et al.* 2011; Vilà *et al.* 2013; Jucker *et al.*  
44 2014a). Consequently, promoting tree diversity is seen as a promising strategy for increasing  
45 timber yields and carbon sequestration rates across forest landscapes, while also ensuring a  
46 host of additional ecological co-benefits (Scherer-Lorenzen 2014). Yet despite synthesis  
47 work showing that the overall relationship between diversity and forest productivity is  
48 positive, both the strength and shape of this relationship vary considerably among individual  
49 studies (Zhang, Chen & Reich 2012). Recent work has helped resolve some of these context  
50 dependencies, highlighting how the effects of diversity on productivity are scale-dependent  
51 (Chisholm *et al.* 2013) and can change during stand development (Cavard *et al.* 2011; Lasky  
52 *et al.* 2014). A critical next step is to understand how environmental conditions influence the  
53 relationship between species richness and above-ground wood production (AWP) in forests,  
54 so that we may identify where and when tree diversity has the greatest potential to positively  
55 influence forest productivity (Paquette & Messier 2011; Jucker *et al.* 2014a).

56 Predicting how the relationship between AWP and diversity is likely to change along spatial  
57 environmental gradients requires an understanding of how species interactions are influenced  
58 by environmental conditions. Species interactions are central to explaining positive  
59 biodiversity – ecosystem functioning relationships in plant communities, as they underpin the  
60 role of niche complementarity (Loreau & Hector 2001). A key feature of species interactions

61 is that they are not constant in space, but instead shift in importance and intensity along  
62 environmental gradients (Brooker *et al.* 2008). This is conceptualized in the stress gradient  
63 hypothesis, which predicts that under increasingly stressful environmental conditions  
64 competitive interactions weaken and give way to facilitative processes (see Maestre *et al.*,  
65 2009 for a review). Following this reasoning, it has been hypothesized that biodiversity –  
66 ecosystem functioning relationships should also vary along environmental gradients (Fig. 1a),  
67 becoming progressively stronger as conditions for growth become less favourable (Jucker &  
68 Coomes 2012; Gessner & Hines 2012; Wang *et al.* 2013; Forrester 2014). In forests,  
69 competitive interactions among neighbouring trees have been shown to vary in strength along  
70 environmental gradients (Kunstler *et al.* 2011; He, Bertness & Altieri 2013; Prior & Bowman  
71 2014), and evidence suggests that positive relationships between diversity and productivity  
72 are in fact more common where environmental conditions are most limiting (Paquette &  
73 Messier 2011; Wu *et al.* 2014; Toigo *et al.* 2015). A possible explanation for these patterns is  
74 that the relationship between diversity and productivity shifts in strength as a result of  
75 changes in forest structure which unfold along environmental gradients (Condés, Del Rio &  
76 Sterba 2013). For instance, Potter & Woodall (2014) showed that across the USA forest  
77 biomass is most strongly related to species richness at sites where unfavourable climatic  
78 conditions result in low stem packing densities.

79 In addition to focusing on spatial environmental gradients, it has also been suggested that  
80 interannual variation in climate may influence the strength of the relationship between  
81 diversity and productivity (Forrester 2014). In most cases, the expectation seems to be that  
82 temporal responses to environmental stress should mirror those observed along spatial  
83 environmental gradients. Specifically, years characterized by unfavourable climatic

84 conditions are predicted to exhibit stronger relationships between diversity and AWP (Fig.  
85 1b). However, evidence that the relationship between diversity and forest productivity shifts  
86 consistently between years as a function of climate is far from clear cut, as studies have  
87 reported both stronger and weaker diversity effects in stressful years (del Río, Schütze &  
88 Pretzsch 2013; Jucker *et al.* 2014b; Grossiord *et al.* 2014). Part of the problem is that most  
89 studies so far have only compared the response of two contrasting years (e.g., normal vs  
90 drought year). Consequently, it is hard to know whether underlying patterns are going  
91 unnoticed because of the incomplete temporal resolution of most studies (i.e., comparisons  
92 made exclusively for a select number of years, usually two), or if instead the effects of tree  
93 diversity on productivity simply do not vary systematically in response to interannual  
94 fluctuations in climate.

95 Here we take advantage of the FunDivEUROPE permanent plot network – a novel research  
96 platform designed specifically to quantify the functional significance of biodiversity in  
97 mature forests (Baeten *et al.* 2013) – to test how the effects of tree diversity on productivity  
98 change along both spatial and temporal environmental gradients. We use tree ring data  
99 collected at six sites across Europe to estimate the annual AWP of 209 forest plots over a 15  
100 year period (1997 – 2011). In a first step, we relate trends in productivity to climate data with  
101 the aim of identifying the primary climatic drivers of AWP both across sites (i.e., spatial  
102 patterns) and within them (i.e., temporal patterns). We then test whether the relationship  
103 between tree species richness and productivity changes (i) across sites as a function of large-  
104 scale gradients in climatic productivity and tree packing density (Fig. 1a), and (ii) among  
105 years within each site as a result of fluctuating climatic conditions (Fig. 1b).

## 106 **Materials and Methods**

### 107 STUDY DESIGN

108 The study was conducted in the FunDivEUROPE permanent forest plot network, the design  
109 of which is described in detail in Baeten *et al.* (2013). Here we summarize the key features of  
110 the FunDivEUROPE network, which we supplement with a detailed description of the  
111 rationale and design of the project in Appendix S1 of Supporting Information. The plot  
112 network consists of 209 permanent forest plots (30 × 30 m in size) distributed in six countries  
113 across Europe (hereafter referred to as “sites”). Field sites were chosen to represent major  
114 European forest types, and include boreal forests in Finland, hemiboreal mixed forests in  
115 Poland, beech forests in Germany, mountainous beech forests in Romania, thermophilous  
116 deciduous forests in Italy and Mediterranean mixed forests in Spain. At each site, plots  
117 contain different combinations of locally dominant tree species, and range in species richness  
118 from 1-3 in Finland, 1-4 in Romania, Germany, Italy and Spain, and 1-5 in Poland (see Table  
119 S1 for a full species list). Each target tree species is found in monoculture and is represented  
120 in all species richness levels, with the majority of species’ combinations being replicated at  
121 least twice (59 of 91 combinations). This nested design approach was modelled around that  
122 of the BIOTREE, ORPHEE and FORBIO tree diversity experiments (see Baeten *et al.*, 2013  
123 and references therein). Plots were selected following a set of strict criteria designed to allow  
124 robust comparisons among species richness levels. Specifically, species in mixture share  
125 similar relative abundances (i.e., high species evenness; Fig. S3) and the presence of non-  
126 target species is minimal (< 5% of the total basal area). In addition, plots were established  
127 exclusively in mature forest stands (i.e., those at least in the mid-to-late stages of stem  
128 exclusion) with similar management histories. Lastly – and critically for the purposes of this

129 study – great care was taken to ensure that (i) plots within a site share similar environmental  
130 conditions and (ii) that stand attributes known to influence productivity (e.g., stem density  
131 and stand age; Magnani, Mencuccini & Grace 2000; Kadmon & Benjamini 2006) were not  
132 confounded with the species richness gradient (see Appendix S1 and Fig. S4). Consequently,  
133 while sites differ markedly from one another in terms of climate and forest structure (e.g.,  
134 basal area, stem density, maximum height), plots within a site are spatially clustered and have  
135 similar elevation, topography and soil type.

#### 136 ALLOMETRIC DATA

137 In each plot, all stems  $\geq 7.5$  cm in diameter were identified to species and permanently  
138 marked (12939 stems in total). We recorded the diameter (measured to the nearest 0.1 cm at a  
139 height of 1.3 m off the ground using diameter tape) and height (to the nearest 0.1 m using a  
140 vertex hypsometer, Haglöf AB, Sweden) of each stem, and used these measurements to  
141 estimate the above-ground biomass of each tree based on published biomass functions (see  
142 Jucker *et al.*, 2014a for a complete list of equations used). We also characterized the crown  
143 dominance of each tree using the crown illumination index, which consists in scoring trees  
144 from 1 (suppressed crown receiving no direct light) to 5 (fully exposed dominant crown)  
145 (Clark & Clark 1992). Crown illumination scores are an effective way of gauging the degree  
146 to which a tree is shaded by its neighbours (Jennings, Brown & Sheil 1999; Jucker *et al.*  
147 2014b), and were used to model the effects of competition for light on tree growth (see  
148 following section).



149 CLIMATE DATA

150 We obtained data on air temperature (T; °C), precipitation (P; mm), solar radiation (R; J cm<sup>-2</sup>), actual evapotranspiration (AET; mm), potential evapotranspiration (PET; mm) and the  
151 ratio between precipitation and potential evapotranspiration (P/PET; a measure of water  
152 availability) in the form of daily values covering the period between 1997 and 2011 for each  
153 of the six study sites. With the exception of AET, which was derived from the water balance  
154 model BILJOU (Granier *et al.* 1999; <https://appgeodb.nancy.inra.fr/biljou/>), all climatic data  
155 were obtained directly from the CGMS database of interpolated meteorological data  
156 (AGRI4CAST; <http://mars.jrc.ec.europa.eu/mars>). The AGRI4CAST system combines  
157 meteorological observations, remote sensing imagery and agro-meteorological modelling to  
158 generate climate surfaces for Europe (25 × 25 km resolution).  
159

160 ESTIMATING ANNUAL ABOVE-GROUND WOOD PRODUCTION FROM TREE RING DATA

161 We used tree ring data to obtain a time series of annual above-ground wood production  
162 (AWP; Mg C ha<sup>-1</sup> yr<sup>-1</sup>) for all 209 plots covering the 15 year period between 1997 and 2011.  
163 The approach used to estimate AWP from tree ring data is outlined in detail in Jucker *et al.*,  
164 (2014a). Here we summarize the main analytical steps of this workflow.

165 *Measuring individual tree growth from wood cores*

166 We used 5.15 mm diameter increment borers (Haglöf AB, Sweden) to extract wood cores  
167 from 2950 trees across the plot network following a size-stratified random sampling approach  
168 (Nehrbass-Ahles *et al.* 2014). Once mounted and sanded, wood cores were scanned using a  
169 high resolution flatbed scanner (2400 dpi). From the scanned images we measured annual  
170 radial growth increments for all cored trees using CDendro (Cybis Elektronik & Data,

171 Saltsjöbaden, Sweden). Finally, radial increments for each year between 1997 and 2011 were  
172 converted to annual biomass growth ( $\text{kg C yr}^{-1}$ ) using the biomass functions described above.  
173 Biomass growth was expressed in units of carbon by applying the standard conversion of 0.5  
174 g C per gram of biomass.

#### 175 *Modelling individual tree biomass growth*

176 For each year between 1997 and 2011, we used linear mixed-effects models coded in R  
177 (3.0.1; R Core Development Team, 2013) to model the biomass growth ( $G$ ) of each target  
178 species as a function of tree size, competition for light, species richness and a random plot  
179 effect:

$$180 \quad \log(G_i) = \beta_{0j[i]} + \beta_1 \log(D_i) + \beta_2 CI_i + \beta_3 SR_j + \varepsilon_i \quad (\text{eqn 1})$$

181 where  $D_i$  and  $CI_i$  are, respectively, the stem diameter and crown illumination index of tree  $i$   
182 growing in plot  $j$  (as measured in 2011);  $SR_j$  is the species richness of plot  $j$ ;  $\beta_{0j[i]}$  is a  
183 species' intrinsic growth rate for a tree growing in plot  $j$  (*sensu* Rüger *et al.*, 2012);  $\beta_{1-3}$  are,  
184 respectively, a species' growth response to size, light availability and species richness; and  $\varepsilon_i$   
185 is the residual error. Equation 1 effectively captured variation in biomass growth within and  
186 among species [average  $R^2$  across species = 0.80; for further information see Jucker *et al.*  
187 (2014a)].

#### 188 *Scaling up from tree-level biomass growth to plot-level AWP*

189 Equation 1 was used to estimate the annual biomass growth of all trees that had not been  
190 cored for each year between 1997 and 2011. The annual AWP of each plot was then  
191 quantified by summing the biomass growth of all trees within a plot. This step was repeated  
192 for each year between 1997 and 2011 to generate a 15-year time series of AWP for all 209

193 plots. This approach to estimating AWP was chosen after having ruled out potential biases  
194 associated with temporal autocorrelation in the AWP time series (see Appendix S2).  
195 Similarly to previous diversity – productivity studies in forests (e.g., Paquette & Messier  
196 2011; Vilà *et al.* 2013), estimates of AWP are based exclusively on the growth of trees  
197 present in 2011, with the implicit assumption that mortality rates are unrelated to tree species  
198 diversity (Liang *et al.* 2007).

#### 199 RELATING VARIATION IN AWP ACROSS AND WITHIN SITES TO CLIMATE

200 We combined data on AWP and climate to identify which climatic predictors best explain  
201 variation in wood production across sites (i.e., spatial variation in AWP) and within them  
202 (i.e., interannual variation in AWP). To understand how differences in climate among sites  
203 shape geographic patterns of wood production, we calculated mean annual values of T, P, R,  
204 AET, PET and P/PET for each study site and related these to mean annual AWP using  
205 univariate regressions (Appendix S3). Following the same approach, we also tested the ability  
206 of climatic predictors to explain interannual variation in AWP within sites. For each site, we  
207 aggregated daily climate values to obtain both yearly and seasonal (spring = March, April,  
208 May; summer = June, July, August; autumn = September, October, November) estimates of  
209 T, P, R, AET, PET and P/PET for each year between 1997 and 2011. We then fitted  
210 univariate regressions relating interannual variation in AWP within sites to each climatic  
211 predictor in turn (Appendix S3).

#### 212 STATISTICAL MODELLING OF AWP – DIVERSITY RELATIONSHIPS

213 We devised a two-step approach to test whether AWP – diversity relationships change  
214 predictably (i) across sites and (ii) among years within a site (Fig. 1). We first used a linear

215 regression model to estimate the effects of species richness on AWP (i.e., the slope of the  
 216 AWP – species richness relationship) for (i) each site and (ii) for every year within a site. We  
 217 then related the slope estimates obtained from the model to (i) differences in climate and  
 218 packing density among sites, and (ii) interannual variation in climate within sites (e.g., Reich  
 219 *et al.*, 2012; Wu *et al.*, 2014). We favoured this approach over one in which climatic effects  
 220 on AWP were modelled directly, as it allows the strength of the species richness effect on  
 221 AWP to be intuitively compared both among and within sites (Nakagawa & Cuthill, 2007).

222 *Step 1: Estimate AWP – species richness slopes across and within sites*

223 We modelled AWP as a function of (i) species richness, (ii) plot basal area (to account for the  
 224 effects of stem packing density on productivity), (iii) site (i.e., accounting for variation in  
 225 AWP among study sites), (iv) year nested within site (i.e., allowing AWP to vary among  
 226 years within each site), (v) an interaction term between species richness and site (testing  
 227 whether species richness effects on AWP vary among sites), and (vi) an interaction term  
 228 between species richness and year nested within site (to determine whether species richness  
 229 effects on AWP vary among years within sites):

$$\begin{aligned}
 230 \quad \log(AWP_{ijk}) = & \beta_0 + \beta_1 \log(BA_{ij}) + \beta_2 \log(SR_{ij}) + \beta_3 \text{Site}_j + \beta_4 \text{Year}_{jk} + \\
 231 \quad & \beta_5 [\log(SR_{ij}) \times \text{Site}_j] + \beta_6 [\log(SR_{ij}) \times \text{Year}_{jk}] + \varepsilon_{ijk} \quad (\text{eqn 2})
 \end{aligned}$$

232 where  $AWP_{ijk}$  is the above-ground wood production of plot  $i$  in site  $j$  in year  $k$ ;  $SR_{ij}$  and  $BA_{ij}$   
 233 are, respectively, the species richness and basal area of plot  $i$  in site  $j$ ;  $\text{Site}_j$  and  $\text{Year}_{jk}$  are  
 234 categorical grouping variables which capture how AWP varies among the  $j$  sites and among  
 235 the  $k$  years within each of the  $j$  sites, respectively;  $\beta_0 - \beta_6$  are parameters to be estimated from  
 236 the data (i.e., intercept and slope coefficients); and  $\varepsilon_{ijk}$  is the residual error.

237 In equation 2, both “Site” and “Year” were modelled as fixed effects so that we could  
238 estimate uncertainties for all parameters. To verify the robustness of the parameter estimates  
239 obtained from equation 2 (particularly  $\beta_5$  and  $\beta_6$ , which define how the effect of tree diversity  
240 on AWP varies across and within sites), we first compared them with those estimated using a  
241 mixed-effects modelling approach. In addition to this, we also fitted an additional model in  
242 which we accounted for a number of potentially confounding factors which could  
243 conceivably alter the relationship between tree diversity and AWP (elevation, terrain slope,  
244 soil depth, rock cover and stand age). As all three approaches yielded almost identical results  
245 (see Appendix S4), here we focus on the parameter estimates obtained through equation 2.  
246 Note that support for the interaction terms in equation 2 was assessed through comparison  
247 with simpler models lacking interacting effects.

248 *Step 2: Relate variation in slopes to climate and packing density*

249 From equation 2 we obtained the slope of the AWP – species richness relationship for (i)  
250 each site (i.e.,  $\beta_5$ ) and (ii) for all years within a site (i.e.,  $\beta_6$ ). We then used Pearson’s  
251 correlation coefficients ( $r$ ) to explore whether (i) variation in the strength of the species  
252 richness effect among sites is related to differences in macroclimate and/or basal area (a  
253 measure of tree packing density which captures both the mean size and number of trees  
254 within a plot) (Curtis & Marshall 2000), and (ii) whether the influence of species richness on  
255 AWP changes over time within a site depending on how suitable climatic conditions are in a  
256 given year.

257 **Results**

258 INFLUENCE OF CLIMATE AND PACKING DENSITY ON AWP

259 AWP varied considerably among sites (M2 vs M1 in Table 1), differing almost fivefold  
260 between Spain ( $0.65 \text{ Mg C ha}^{-1} \text{ yr}^{-1}$ ) – the least productive site – and Romania ( $3.11 \text{ Mg C ha}^{-1}$   
261  $\text{yr}^{-1}$ ), where AWP rates were highest. Geographic variation in AWP was strongly related to  
262 differences in evapotranspiration among sites (Fig. 2a; see Appendix S3 for comparison with  
263 other climatic drivers). Sites where annual AET was low – either because of insufficient  
264 rainfall (e.g., Spain) or low temperatures (e.g., Finland) – had much slower rates of AWP  
265 than those where AET was high (e.g., Germany and Romania). However, AET was only a  
266 weak predictor of interannual variation in AWP within sites (see Table S2 and following  
267 paragraph). Instead, we found that increased annual AET was associated with the  
268 development of forests with greater basal areas (Fig. S7). This in turn resulted in markedly  
269 faster rates of AWP – as basal area was a key driver of AWP across the plot network (Fig. 2b  
270 and Table 1) – suggesting that at least in part the link between evapotranspiration and  
271 productivity is mediated through changes in stand packing density.

272 In addition to varying across sites, AWP also fluctuated strongly among years within each  
273 site (M3 vs M2 in Table 1). Climate explained much of this interannual variation in  
274 productivity (Fig. 3), with a clear distinction emerging between Finland – where AWP was  
275 co-limited by temperature and solar radiation – and all other sites – where annual growth was  
276 instead primarily water-limited (Table 2). In Finland, the single best climatic predictor of  
277 annual AWP was spring PET (Fig. 3b and Table 2), with productivity peaking in years  
278 characterized by high solar radiation and mild spring temperatures which together contributed

279 to high PET rates (Table S2). For all other sites, fluctuations in annual AWP was best  
280 captured by P/PET (Table 2), with productivity declining sharply in years when evaporative  
281 demands (i.e., PET) were not met by precipitation (Fig. 3a,c-f). The timing of drought proved  
282 equally important in explaining AWP patterns. For both Mediterranean sites, low P/PET  
283 values during spring months were associated with strong reductions in AWP (Fig. 3a,c),  
284 while summer drought had a much less pronounced influence on productivity (Table S2). In  
285 contrast, for Germany and Poland the best predictor of AWP was P/PET integrated over the  
286 entire year (Fig. 3d-f), while in Romania AWP was influenced by autumn P/PET (Fig. 3e).  
287 Note that for all sites, AWP – climate relationships were best captured by linear functions  
288 (see Table 2 for estimated regression coefficients).

#### 289 AWP – DIVERSITY RELATIONSHIPS ACROSS SITES

290 We found strong support for the inclusion of the interaction term between species richness  
291 and “site” in equation 2 (M4 vs M3 in Table 1), indicating that the effect of species richness  
292 on AWP varies substantially across sites. In accordance with the stress gradient hypothesis  
293 (Fig. 1a), the slope of the AWP – species richness relationship decreased progressively along  
294 the AET climatic productivity gradient ( $r = -0.88$ ,  $P = 0.020$ ,  $n = 6$ ; Fig. 2c), ranging from  
295 strongly positive in Spain ( $0.24 \pm 0.06$ ) – where evapotranspiration and productivity were  
296 low – to weakly negative in Romania ( $-0.05 \pm 0.07$ ) – where AET and AWP were greatest.  
297 This shift in the strength of the diversity effect on productivity was equally well captured by  
298 accounting for differences in basal area across sites ( $r = -0.90$ ,  $P = 0.013$ ,  $n = 6$ ), with  
299 diversity effects being strongest at sites where environmental conditions led to the  
300 development of less densely packed stands (Fig. 2d and Fig. S7).

302 In contrast to patterns across sites, we found little evidence to suggest that species richness  
303 effects on AWP varied strongly among years within sites (M5 vs M4 in Table 1).  
304 Furthermore, we found no consistent relationship between the strength of the species richness  
305 effect on AWP and climatic conditions within a given year when looking across the six sites  
306 (Fig. 4). The only site to show a relationship between the magnitude of the species richness  
307 slope and climate was Finland (Fig. 4b), where species richness effects weakened in years  
308 when spring PET was high and climatic conditions for growth were favourable ( $r = -0.73$ ,  $P$   
309  $< 0.01$ ,  $n = 15$ ). A similar response was found for Romania, although the pattern between  
310 diversity effects and climate (autumn P/PET) was weaker ( $r = -0.44$ ,  $P = 0.10$ ,  $n = 15$ ; Fig.  
311 4f). Instead, in the case of Poland (Fig. 4d), and to a much lesser extent Spain (Fig. 4a), we  
312 observed the opposite trend, with diversity effects strengthening during non-drought years ( $r$   
313  $= 0.49$ ,  $P = 0.06$  and  $r = 0.21$ ,  $P = 0.45$ , respectively;  $n = 15$ ). These observations do not  
314 support the predictions of the stress gradient hypothesis (Fig. 1b). Lastly, in the case of Italy  
315 (Fig. 4c) and Germany (Fig. 4e) we found no discernible pattern linking diversity effects and  
316 climatic conditions across years ( $r = 0.03$ ,  $P = 0.93$  and  $r = 0.05$ ,  $P = 0.87$ , respectively;  $n =$   
317 15).



318 **Discussion**

319 In line with a number of recent studies, we found a generally positive relationship between  
320 tree diversity and forest productivity across Europe (Morin *et al.* 2011; Vilà *et al.* 2013;  
321 Pretzsch *et al.* 2015). Importantly, however, our results also highlighted that the effects of  
322 species richness on productivity are strongly context dependent, varying in space and – to a  
323 lesser extent – time. Specifically, AWP – species richness relationships shifted predictably in  
324 strength across sites, becoming progressively stronger under harsher environmental  
325 conditions (Fig. 2). In contrast, we found that interannual variation in the strength of species  
326 richness effects within sites was rather weak and could not be consistently explained by year-  
327 to-year fluctuations in climate (Fig. 4).

328 **AWP – DIVERSITY RELATIONSHIPS ARE STRONGEST IN STRESSFUL ENVIRONMENTS**

329 Macro-scale patterns in wood production were best captured by differences in  
330 evapotranspiration among sites (Fig. 2a), confirming what has previously been reported in the  
331 literature (Stephenson 1998; Oberle, Grace & Chase 2009). Evapotranspiration integrates the  
332 effects of water availability and temperature on plant growth (Stephenson 1998; Boisvenue &  
333 Running 2006): in order for trees to grow, temperatures need to be sufficient to initiate  
334 photosynthesis and allow the microbially driven mobilization of soil nutrients, and at the  
335 same time enough water needs to be available to meet evaporative demands. Our results also  
336 suggest that the link between evapotranspiration and productivity is largely mediated through  
337 changes in stand basal area which unfold across AET gradients (Fig. S7), as opposed to a  
338 direct influence of AET on AWP (Stephenson 1998; Coomes *et al.* 2014; Michaletz *et al.*  
339 2014). While interannual fluctuations in AET were not a particularly strong predictor of

340 variation in productivity within sites (Tables S2), sites with higher mean annual AET also  
341 had greater basal areas, which in turn drove faster rates of wood production (Fig. 2b).

342 The strength of the species richness effect on productivity also shifted along the  
343 evapotranspiration gradient, and was strongest at sites where AET was low – either as a result  
344 of low annual rainfall (e.g., Spain) or because of low temperatures (e.g. Finland) – and  
345 conditions for growth were poor (Fig. 2c). Conversely, at sites where conditions for growth  
346 were more favourable, we saw a drastic reduction in the importance of diversity as a driver of  
347 wood production. These results are broadly consistent with the predictions of the stress  
348 gradient hypothesis (Fig. 1a), as well as the findings of a number of recent papers (Paquette  
349 & Messier 2011; Jucker & Coomes 2012; Wu *et al.* 2014; Toïgo *et al.* 2015). For instance,  
350 using forest inventory data from France, Toïgo *et al.* (2015) showed that overyielding was  
351 more frequent at low productivity sites. Similarly, Wu *et al.* (2014) found that the strength of  
352 the relationship between tree diversity and biomass weakened when transitioning from  
353 boreal, to temperate and subtropical forests in China.

354 In addition to the stress gradient hypothesis, a number of alternative frameworks have also  
355 been proposed to explain how environmental conditions shape the relationship between  
356 diversity and productivity. For instance, Forrester (2014) noted that the outcome of species  
357 interactions will depend on the type of limiting resource or resources (e.g., water, light,  
358 nutrients), and whether – on average – interactions among neighbouring trees improve the  
359 availability of those resources (also see Pretzsch *et al.* 2015). What our results suggest is that  
360 the net outcome of species interactions shifts along environmental gradients and that – on  
361 balance – competition tends to decrease in importance in more stressful environments  
362 allowing greater room for complementarity (Kunstler *et al.* 2011; Prior & Bowman 2014).

363 Climate can influence interactions among neighbouring trees directly by modulating the  
364 relative performance and strategy of species within a community (Brooker *et al.* 2008;  
365 Maestre *et al.* 2009). In addition to this, climate can influence interactions among  
366 neighbouring trees indirectly through its effect on forest structure. Climate plays an important  
367 role in determining local species pools through environmental filtering (e.g., Stahl *et al.*,  
368 2014), as well as shaping species demographic rates (Lines, Coomes & Purves 2010;  
369 Vanderwel, Lyutsarev & Purves 2013; Coomes *et al.* 2014) and influencing how trees  
370 allocate carbon above and below-ground (Lines *et al.* 2012; Reich *et al.* 2014), all of which  
371 interact to determine forest structure. As our results suggest that positive AWP – diversity  
372 relationships are much more likely in forests with low packing densities (Fig. 2d), the effect  
373 of climate on forest structure appears to be particularly important in modulating the effects of  
374 diversity on productivity (Condés *et al.* 2013; Potter & Woodall 2014). This may explain why  
375 in Mediterranean forests, where drought impedes the development of densely packed stands,  
376 the effects of diversity on productivity tend to be strongly positive (Vilà *et al.* 2007; Ruiz-  
377 Benito *et al.* 2014; Jucker *et al.* 2014b).

#### 378 INTERANNUAL FLUCTUATIONS IN CLIMATE HAVE WEAK AND IDIOSYNCRATIC EFFECTS ON AWP 379 –DIVERSITY RELATIONSHIPS WITHIN SITES

380 AWP responses to interannual fluctuations in climate within sites closely matched  
381 expectations based on published reports, with most forest types in Europe being primarily  
382 limited by water availability, while productivity in boreal forests depends strongly on  
383 temperature and solar radiation (Boisvenue & Running 2006; Babst *et al.* 2013). In terms of  
384 AWP – diversity relationships, we generally found that variation within sites was  
385 considerably weaker than variation among them (Table 1). Moreover, we found no evidence

386 of consistent pattern linking variation in diversity effects across years to climate (Fig. 4).  
387 Certain functional responses match expectations, such as the decreased strength of diversity  
388 effects during warmer years in Finland (Fig. 4b). In boreal systems, warmer springs are  
389 associated with earlier leaf-out dates (Polgar & Primack 2011). While this positively  
390 influences productivity by increasing the length of the growing season (Chen *et al.* 1999;  
391 Polgar & Primack 2011), it is also likely to reduce the degree of phenological mismatch (i.e.,  
392 temporal complementarity) between evergreen conifers and deciduous broadleaf species  
393 (Sapijanskas *et al.* 2014). In contrast, we found no consistent effect of drought on the strength  
394 of diversity effects across forest types (Grossiord *et al.* 2014). For Spain, our results generally  
395 confirm those of a previous study comparing responses in two years with contrasting climate  
396 (a drought year and a wet year), which showed that drought tends to exacerbate competition  
397 among neighbouring trees (Jucker *et al.* 2014b). However, in the case of Poland, Germany  
398 and Romania we find strongly contrasting patterns despite the fact that these sites share  
399 similar species compositions and have broadly similar climates.

400 Differences between responses across sites – which strongly matched theoretical predictions  
401 – and those within sites – which instead exhibited weak and contrasting patterns – suggest  
402 that in addition to climate itself, a key factor in shaping the strength of diversity effects on  
403 productivity is how densely trees pack in space. If the relationship between diversity and  
404 productivity was directly influenced by climate, diversity effects would be expected to shift  
405 consistently both across and within sites. Instead, under the assumption that packing density  
406 is what ultimately determines the strength of the relationship between diversity and  
407 productivity (Condés *et al.* 2013; Potter & Woodall 2014), temporal patterns would be  
408 unlikely to match spatial trends. The reason for this is that although basal area varies strongly

409 among sites (Fig. 2b), changes in packing density are decoupled from interannual variation in  
410 climate in forest ecosystems (i.e., basal area does not increase or decrease markedly from one  
411 year to next based on climatic suitability). Recent work in coastal dunes plant communities  
412 has shown that shifts from competition to facilitation among years depend critically on  
413 differences in vegetation cover: during years characterized by harsh environmental conditions  
414 vegetation cover decreases, resulting in an increase in facilitation, while in more favourable  
415 years increased vegetation cover exacerbates competition (Doxford, Ooi & Freckleton 2013).  
416 However, because trees are long-lived and do not regenerate the bulk of their above-ground  
417 biomass each year, a similar response in terms of basal area cannot be observed in forests  
418 (Stephenson & Mantgem 2005). As a result, the way in which diversity effects change among  
419 years may simply be more context dependent in the case of forests [e.g., highly dependent on  
420 species' traits or on the type of limiting resource; Forrester (2014)].

#### 421 IMPLICATIONS FOR FOREST CONSERVATION AND MANAGEMENT

422 Identifying where and when tree diversity has the greatest potential to positively influence  
423 forest productivity has important implications for forest management and conservation  
424 practises, as well as efforts to mitigate climate change (Zhang *et al.* 2012; Cardinale *et al.*  
425 2012; Scherer-Lorenzen 2014). We found that in terms of maximising rates of wood  
426 production, the benefits of maintaining diverse forests are most pronounced in systems where  
427 environmental conditions strongly limit productivity. While our study provides a useful  
428 framework for predicting under which conditions tree diversity is likely to matter most, there  
429 are however several reasons why practises aimed at maintaining diverse forests should not  
430 necessarily be limited to specific ecological contexts or geographic regions. For instance, in  
431 addition to promoting forest productivity, tree diversity has also been shown to help stabilize

432 wood production over time across a range of forest types, highlighting the fact that mixed-  
433 species forests are able to remain productive under a wider range of environmental conditions  
434 than monocultures (Jucker *et al.* 2014a; Morin *et al.* 2014). Secondly, although we detected  
435 clear differences in the importance of tree diversity as a driver of productivity among forest  
436 types, only one site showed any indication of a negative association between diversity and  
437 productivity. Consequently, even though gains in productivity may be negligible for certain  
438 forest types, maintaining diverse forests is unlikely to adversely affect wood production and  
439 has the advantage of delivering a number of added ecological and economic co-benefits (e.g.,  
440 reduced risk of pest and pathogen outbreaks, increased associated biodiversity, greater soil  
441 carbon storage; Scherer-lorenzen, 2014).

442 In addition to highlighting under which circumstance tree diversity is currently most  
443 important for forest productivity, our study also provides a number of clues as to how climate  
444 change is likely to influence AWP – diversity relationships in future forests. By the end of  
445 this century Mediterranean forests in Europe are expected to suffer more frequent and  
446 prolonged periods of drought, while boreal systems are predicted to warm considerably and  
447 experience longer growing seasons (Jacob *et al.* 2014). As a result, diversity effects may  
448 weaken in strength, particularly in the case of boreal forests where spring warming is  
449 expected to reduce the degree of phenological mismatch among coexisting tree species  
450 (Polgar & Primack 2011). More importantly, however, our results suggest that longer-term  
451 responses of forests to climate change – such as changes in species composition and forest  
452 structure – are what will ultimately determine how the relationship between tree diversity and  
453 forest productivity will look like in the future.

454 **Acknowledgements**

455 We thank FunDivEUROPE site managers and field technicians for establishing the  
456 permanent plots, and are grateful to I. Dănilă, L. Nichiforel and A. Benneter for assisting with  
457 field data collection, and to A. Granier for providing climate data. Four anonymous reviewers  
458 provided insightful comments on an earlier draft of this manuscript. The research leading to  
459 these results received funding from the European Union Seventh Framework Programme  
460 (FP7/2007-2013) under grant agreement n° 265171.

461 **Data Accessibility**

462 Above-ground wood production (AWP) data: uploaded as online supporting information

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674 **Supporting information**

675 Additional supporting information may be found in the online version of this article:

676 **Appendix S1:** FunDivEUROPE project

677 **Appendix S2:** Temporal autocorrelation in AWP

678 **Appendix S3:** Relating variation in AWP to climate

679 **Appendix S4:** Estimating diversity effects across and within sites

680 **Appendix S5:** Above-ground wood production (AWP) data archive

681

682 **Tables**

683 **Table 1:** Model comparison statistics for a set of candidate models testing whether diversity  
 684 effects on above-ground wood production (AWP) vary across and within sites. In the baseline  
 685 model (M1), AWP is expressed as a function of plot basal area (BA) and species richness  
 686 (SR). “Site” indicates a model which accounts for variation in AWP among sites, while  
 687 “Year” tests whether AWP varies among years within each site. Model M4 tests whether the  
 688 effects of SR on AWP varies among sites (i.e., interaction between SR and Site), while model  
 689 M5 further allows species richness effects to vary among years within a site (i.e., interaction  
 690 between SR and Year) and corresponds to equation 2 in the text. Models of increasing  
 691 complexity were compared to simpler ones (e.g., M5 vs M4) using F-tests to determine  
 692 whether the reduction in the residual sum of squares between models was statistically  
 693 significant ( $\alpha = 0.05$ ). In addition to this we also report the  $R^2$  and AIC of each model. Note  
 694 that AWP, SR and BA were log-transformed prior to model fitting.

<b>Model</b>	<b>Model structure</b>	<b><i>P</i> (&gt;F)</b>	<b><math>R^2</math></b>	<b>AIC</b>
M0	AWP ~ BA	–	0.40	4306
M1	AWP ~ BA + SR	<0.0001	0.42	4276
M2	AWP ~ BA + SR + Site	<0.0001	0.74	1738
M3	AWP ~ BA + SR + Site + Year	<0.0001	0.78	1291
M4	AWP ~ BA + SR + Site + Year + SR×Site	<0.0001	0.79	1248
M5	AWP ~ BA + SR + Site + Year + SR×Site + SR×Year	n.s.	0.79	1384

695



696 **Table 2:** Best climatic predictors of interannual variation in above-ground wood production  
697 (AWP) for each study site. Regression equations (with 95% confidence intervals for  
698 parameter estimates in brackets) correspond to fitted relationships illustrated in Fig. 3. To  
699 facilitate the interpretation of regression coefficients, climatic predictors were cantered prior  
700 to model fitting (i.e., for each site, intercepts indicate the mean AWP across years). PET =  
701 potential evapotranspiration; P/PET = precipitation/PET.

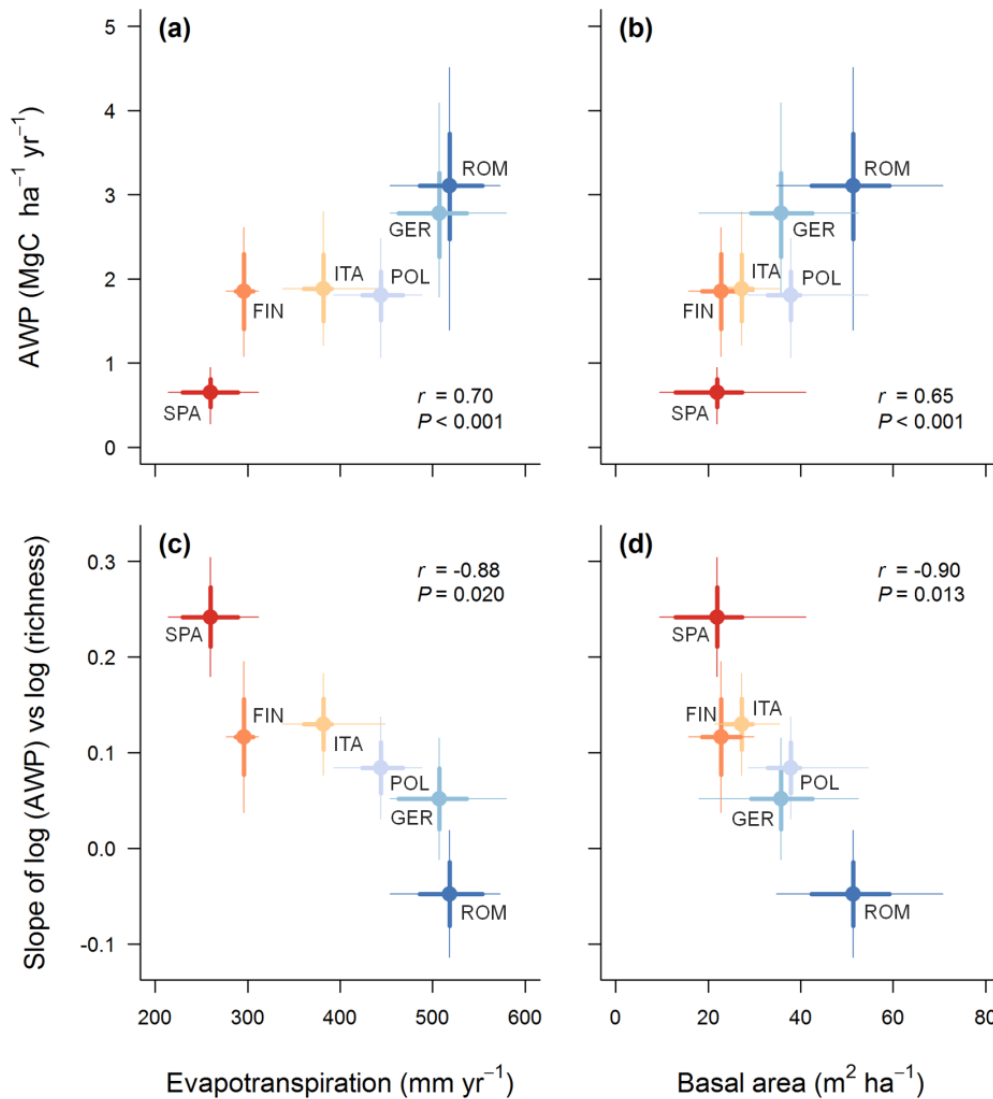
<b>Site</b>	<b>Climatic predictor</b>	<b>Regression equation</b>
Finland	Spring PET	$AWP = 1.85 (0.06) + 0.012 (0.003) \times PET$
Poland	Annual P/PET	$AWP = 1.81 (0.04) + 0.87 (0.23) \times (P/PET)$
Germany	Annual P/PET	$AWP = 2.78 (0.07) + 1.11 (0.44) \times (P/PET)$
Romania	Autumn P/PET	$AWP = 3.11 (0.11) + 0.40 (0.14) \times (P/PET)$
Italy	Spring P/PET	$AWP = 1.88 (0.05) + 0.48 (0.24) \times (P/PET)$
Spain	Spring P/PET	$AWP = 0.65 (0.02) + 0.28 (0.09) \times (P/PET)$

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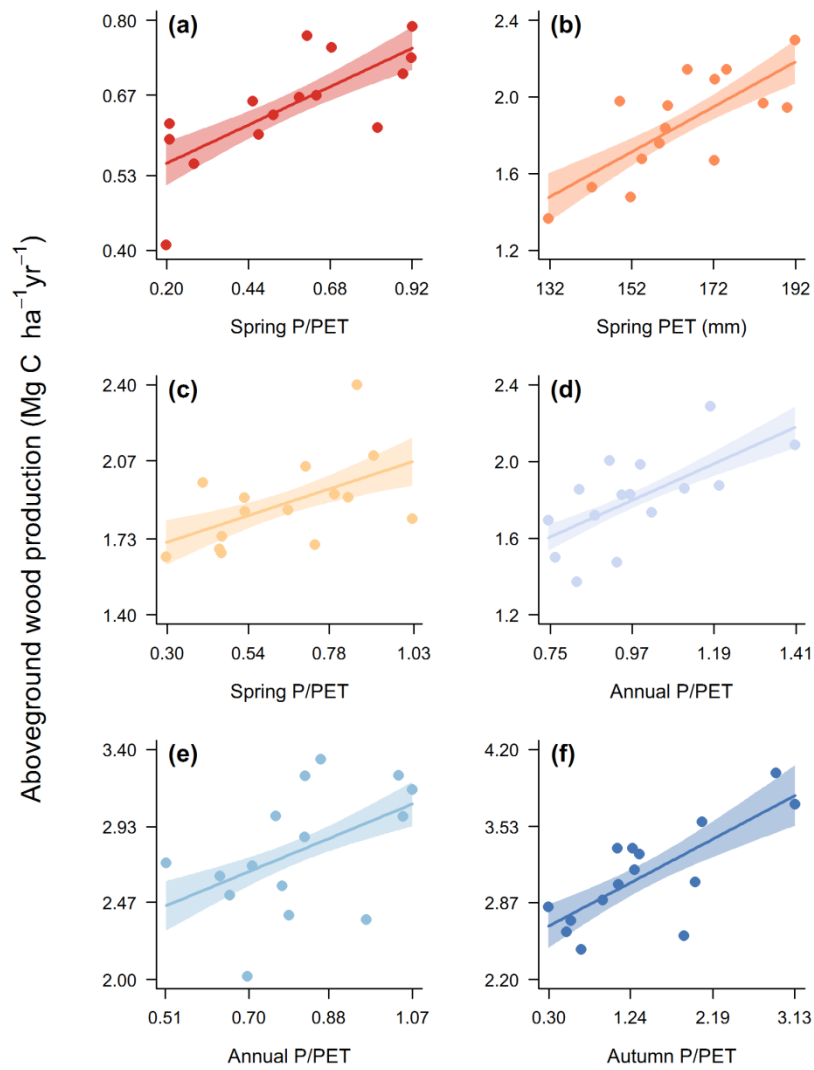
705 **Fig. 1:** Schematic diagram illustrating how the relationship between productivity and  
 706 diversity is expected to change (a) across sites along an environmental gradient and (b)  
 707 among years within a site. Panel (a) depicts different sites along an environmental gradient.  
 708 Sites where productivity is low (e.g., as a result of strong environmental limitations on  
 709 growth) exhibit strong positive effects of diversity on productivity, while sites where growing  
 710 conditions are more favourable show little or no benefit from species mixing. Panel (b)  
 711 illustrates how the effects of diversity on productivity vary among years within a given site.  
 712 The expectation is that the slope of the diversity – productivity relationship will be steeper  
 713 than average in low productivity years, weakening instead when conditions for growth  
 714 improve.



715

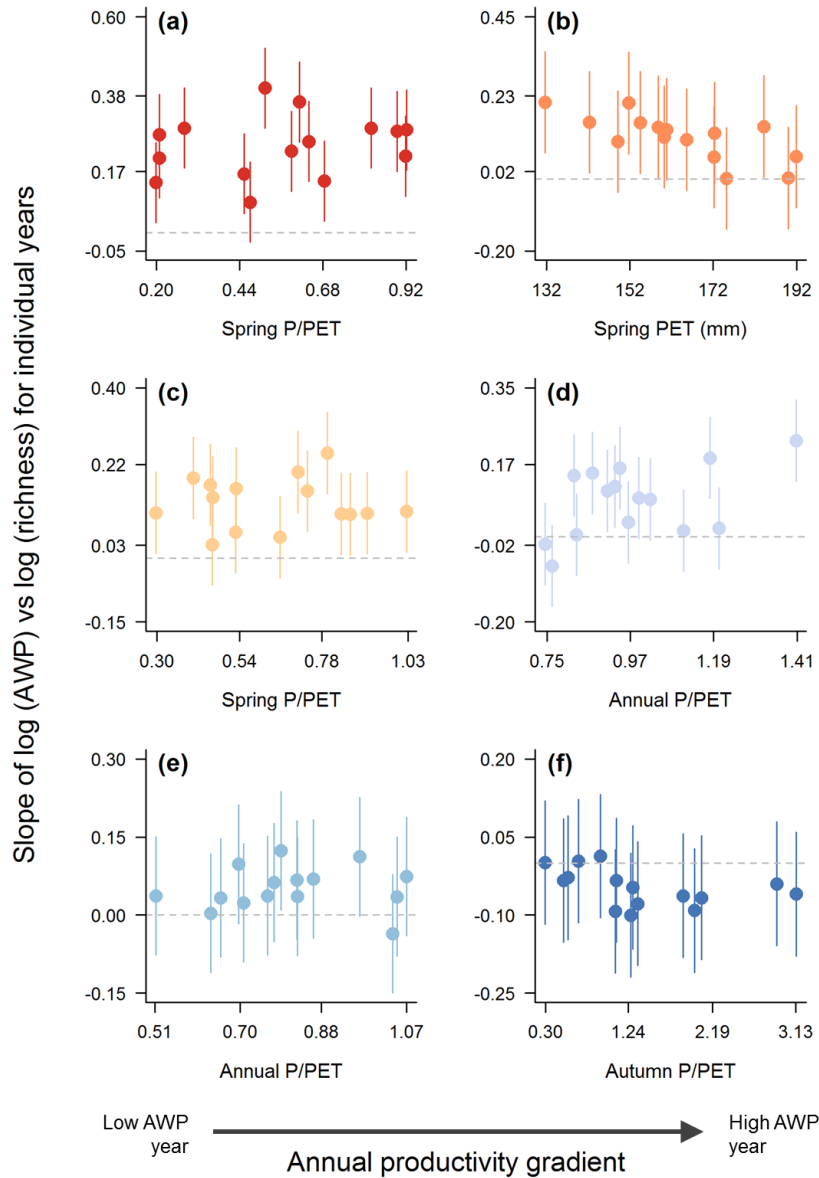
716 **Fig. 2:** Variation in (a–b) above-ground wood production (AWP) and (c–d) the slope of the  
 717 AWP – species richness relationship among sites. Panels on the left show how  
 718 evapotranspiration (AET) influences (a) AWP and (c) the strength of the AWP – species  
 719 richness relationship, while right-hand panels (b and d) illustrate variation in response to plot  
 720 basal area. For AWP and basal area, points represent mean values across all plots within a  
 721 site, while AET is the long-term mean annual evapotranspiration registered at each site  
 722 (between 1997 and 2011). Error bars delimit the interquartile range (thick lines) and 95%  
 723 limits (thin lines) of the data. Slopes of log (AWP) vs log (species richness) were obtained

724 from equation 2 in the main text (thick lines:  $\pm 1$  SE; thin lines:  $\pm 2$  SE). Points are labelled  
725 using the first three letters of each site. Pearson's correlation coefficients ( $r$ ) and  
726 corresponding  $P$ -values reported in the top panels (a–b) were calculated for  $n = 209$   
727 (corresponding to the number of FunDivEUROPE forest plot), while for the bottom panels  
728 (c–d)  $n = 6$  (corresponding to the number of FunDivEUROPE sites).



729

730 **Fig. 3:** Variation in annual above-ground wood production (AWP) as a function of climate  
 731 for (a) Spain, (b) Finland, (c) Italy, (d) Poland, (e) Germany and (f) Romania. Points  
 732 correspond to the mean AWP of each year (1997 to 2011; calculated across all plots within a  
 733 site). Fitted relationships from linear regression models (with shaded 95% confidence  
 734 intervals) are shown for each panel (see Table 2 for regression equations). PET = potential  
 735 evapotranspiration; P/PET = precipitation / PET.



736

737 **Fig. 4:** Interannual variation in the slope of the AWP – species richness relationship as a  
 738 function of climate in (a) Spain, (b) Finland, (c) Italy, (d) Poland, (e) Germany and (f)  
 739 Romania. Slopes ( $\pm 1$  SE) of log (AWP) vs log (species richness) for each year between 1997  
 740 and 2011 are estimates from a linear regression model (equation 2 in the main text). Dashed  
 741 grey lines mark a slope of zero. PET = potential evapotranspiration; P/PET = precipitation /  
 742 PET.