Correlates of long-term land-cover change and protected area performance at priority conservation sites in Africa

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SUMMARY

The loss of natural habitats is a major threat to biodiversity, and protected area designation is one of the standard responses to this threat. However, greater understanding of the drivers of habitat loss, and of the circumstances under which protected areas succeed or fail is still needed. We use visual assessment of satellite images to quantify land-cover change over periods of up to 30 years in and around a matched sample of protected and unprotected Important Bird and Biodiversity Areas (IBAs) in Africa. We modelled the annual survival of forests and other natural land covers as a function of a range of environmental and anthropic predictors of plausible drivers. The best-supported model indicated that survival rates of natural land-cover were highest in steeper areas, at higher altitudes, in areas with lower human population densities and in areas where the cover of natural habitats was already higher at the start of the period. Survival rates of natural land-cover in protected areas were, on average, around twice those in unprotected areas, but the difference between them varied along different environmental gradients. The overall survival rates of both protected and unprotected forests were significantly lower than those of other natural land-cover types, but the net benefit of protection, in terms of the absolute difference in rates of loss between protected and unprotected sites, was higher in forest. Interaction terms indicated that as slope and altitude increased, the natural protection offered by topography increasingly nullified the additional benefits of legislative protection. Furthermore, protected area designation offered reduced additional benefits to the survival of natural land-cover in areas where rates of conversion were higher at the start of the observation period. Variation in the impacts of protected area status along different environmental gradients indicates that targets to improve the world’s protected area network, such as Aichi Target 11 of the CBD, need to look beyond simple area-based metrics. Our methods and results contribute to the development of a protocol for prioritising places where protection is likely to have the greatest effect.
INTRODUCTION

Conversion of natural habitats is one of the biggest threats to global biodiversity (Pereira et al. 2010), and designation of protected areas one of the most widespread approaches to mitigating this threat. It is now well established that protected areas can be effective at reducing rates of loss of natural land cover (e.g. Andam et al. 2008; Gaveau et al. 2009; Selig & Bruno 2010; Beresford et al. 2013; Geldmann et al. 2013; Butsic et al. 2015), but that they vary greatly in the extent to which they achieve this (Andam et al. 2008; Joppa & Pfaff 2011; Francoso et al. 2015; Paiva et al. 2015). Targets for expanding protected area networks are often based largely or solely on the area covered. Aichi target 11 of the CBD sets targets for the coverage of land and sea by protected areas, but then simply states that these need to be effectively managed. It does not set a target against which this effectiveness can be assessed. Thus, designation alone is an insufficient measure of the effectiveness of a protected area network. Conservation requires also an understanding of the effectiveness of protection, and how this varies between locations. Estimates of both extent and effectiveness are needed to develop effective networks of protected areas capable of meeting conservation targets at local to global scales.

As there have been inconsistencies in the way that terms such as “impact” and “effectiveness” of protected areas have been applied and interpreted across different studies, we identify three key elements determining the overall impact of a protected area: (i) the conservation value of the site, which is a function of the biota it holds and its area, (ii) the degree of threat to the site and (iii) the degree to which that threat is reduced by designation (i.e. the effectiveness of protection).

Analysis of the effectiveness of protected areas is not as simple as comparing rates of change with surrounding areas or randomly chosen points due to the non-random distribution of protected areas within the landscape (Joppa & Pfaff 2009; Nelson & Chomitz 2011; Joppa & Pfaff 2011; Beresford et al. 2013; Pfaff et al. 2015a), and to other effects such as the potential ‘leakage’ of conversion from within to outside the protected area boundary (Ewers & Rodrigues 2008; Robalino et al. 2015). Consequently, it is necessary to match protected (treatment) and unprotected (control) sites to control for confounding effects that might drive both the likelihood of designation and of conversion. Matching by appropriate characteristics that may
influence these rates can effectively control for locational bias in the siting of protected areas (Joppa & Pfaff 2011; Nelson & Chomitz 2011; Beresford et al. 2013).

A number of studies have been undertaken from regional to global scales in which sites have been matched by variables such as ecoregion, rainfall, agricultural suitability, elevation, slope and distance from roads and urban areas, with an emerging consensus that impacts in terms of avoided conversion are greatest on flatter land and at sites near to roads and cities (Joppa & Pfaff 2011; Pfaff et al. 2015a). The effectiveness of protected areas in meeting their potential can be influenced by a range of factors such the type of designation and the level and effectiveness of enforcement, as well as the siting and accessibility of the protected area (Nelson & Chomitz 2011; Pfaff et al. 2014; Pfaff et al. 2015b).

Even in studies in which matching is used to assess more accurately protected area impacts, it can still be difficult to disentangle the relative effects of different elements of impact, as there are often complex interactions between them (Geldmann et al. 2013). In particular, there are often negative correlations between threat and effectiveness, with protection being more effective on sites with lower impact in terms of avoided conversion (e.g. Ahrends et al. 2010; Boakes et al. 2010; Brun et al. 2015). On the other hand, threat and conservation value are often positively correlated, with sites containing the most threatened habitats considered those of highest conservation value.

These relationships are further complicated by the fact that a single variable such as slope, elevation or accessibility can often influence, or correlate with, more than one element of impact. For example, protected sites on steeper slopes may be more effective than those on flatter ground, but may have a lower impact in terms of avoided conversion if rates of conversion are already low because of the steeper ground (Joppa & Pfaff 2011). Depending on how conservation value is measured, the tendency for low rates of conversion may mean that the site is considered of high conservation value (i.e. pristine condition) or of relatively low value (if low conversion rates result in it being relatively abundant in the landscape).

Much conservation effort is focused on the designation and management of protected areas across the globe, so a better understanding of the factors influencing their effectiveness in reducing conversion of natural land-cover is needed to assess
the relative benefits of existing protected areas and to identify sites whose future protection would bring the greatest benefits. This is especially true given that they represent the mechanism through which so many conservation issues are tackled. Of the 20 Aichi Biodiversity Targets set under the Convention for Biological Diversity (CBD), Target 11 relates directly to the designation of protected areas. Protected area networks can also make a major contribution to other targets, e.g. Target 1, relating to the awareness of people to the values of biodiversity, Target 5, relating to loss of natural habitats, Target 12, on preventing the extinction of known threatened species, Target 14, on the preservation of ecosystem services, and Target 15, on carbon storage ((Scharlemann et al. 2010; Beresford et al. 2016). They also have a recognised role to play in mitigating the impacts of climate change on people and biodiversity (Loarie et al. 2009; Thomas & Gillingham 2015).

In a previous study (Beresford et al. 2013), we focused on measuring the impact of protection in terms of avoided conversion by comparing rates of loss of natural land cover between protected and unprotected sites of recognised conservation value (Important Bird and Biodiversity Areas, or IBAs) in Africa. IBAs are sites of global significance for the conservation of the world’s birds, identified using semi-quantitative criteria (Fishpool & Evans 2001), and are Key Biodiversity Areas (IUCN 2016). The most prevalent threats to African IBAs are associated with land-cover change (Buchanan et al. 2009). Use of this set of sites eliminates from our study the influence of protected areas of relatively low conservation value that were designated primarily because of their low opportunity cost and low likelihood of conversion.

We previously established that protection is effective at reducing, but not halting, land-cover change on protected IBAs compared to unprotected IBAs (Beresford et al. 2013). Here, we develop this work by using the same database of spatially-explicit information on land-cover change in and around African IBAs to identify the characteristics of points which were (and were not) converted during the study. In doing so, we hope to identify potential drivers of conversion and enable pro-active conservation. This information could be used to better target monitoring of sites. Additionally, and perhaps more importantly, we could identify areas most at risk from habitat loss and potentially increase the rapidity with which threats are tackled on the
ground in these places. We produce statistical models of relationships (GLMMs) and consider the interaction between correlate variables and whether or not the point is protected. By examining output fitted relationships we can determine whether the form of the relationship varies within and outside protected areas. This allows differing conservation strategies to be developed within and outside protected areas, if appropriate. By investigating changes on sites of objectively defined high conservation importance (IBAs), this is the first study to control for variation in potential impact in terms of conservation value, as well as avoided conversion, which we account for using site-level matching.

METHODS

Site selection

Protected and unprotected IBAs were matched at the site level to reduce the known problem of the non-random distribution of PAs, which are often designated in remote and inaccessible areas where the risk of damage is inherently lower than in more accessible areas (Joppa & Pfaff 2009). Matching ensures that protected sites are compared with unprotected sites with a comparable risk of sustaining damage by controlling for some of the most likely correlates of environmental risk. Full details of the selection process are given in Beresford et al. (2013). This selection process resulted in a set of 54 protected and 49 unprotected IBAs from the 793 IBAs in continental sub-Saharan Africa and Madagascar for which digital boundaries were available. Our sample of 103 sites differs slightly from Beresford et al. (2013) because we include IBAs from countries for which only a single site was selected by the matching process. We obtained site protection status by intersecting the boundaries of all IBAs with the boundaries of all 1,580 nationally designated PAs in Africa from the World Database of Protected Areas (IUCN categories I to VI) that had digitised boundaries (IUCN & UNEP-WCMC 2010). We defined protected IBAs as those that fell wholly or mostly (>90%) within the boundaries of protected areas that had been designated before 1985. Partially (<90%) protected sites, and those whose protection status changed immediately before or during the assessment period were excluded from the selection process. We defined unprotected IBAs as those that did not overlap any protected areas, irrespective of PA designation date. We excluded
from our analysis all IBAs smaller than 10 km², because the number of points at which we could have assessed land cover would have been small. We matched protected and unprotected IBAs using the command ‘subclass’ in the MatchIt package (Ho et al. 2011) in R (R Development Core Team 2010). We matched sites on the basis of area, mean altitude (both from BirdLife International (2011)), mean distance from roads (National Imagery and Mapping Agency 2000), mean human population density (CIESIN & CIAT 2005) and the most extensive GLC2000 land-cover class in the IBA (Mayaux et al. 2004). Site-level matching ensured that protected and unprotected IBAs were similar at the landscape level.

**Assessment of land-cover change**

We used visual interpretation of satellite imagery to assess dominant land cover within 300 × 300-m sample boxes (‘points’) in each IBA and in a surrounding 20-km buffer using a dedicated Graphical User Interface (Bastin et al. 2013). Points were distributed on a regular grid and spaced 0.5 km apart in IBAs of <50 km², 1.5 km apart in IBAs >50 km² and 3 km apart in the 20-km buffers. Excluding points where cloud-free images were unavailable for any one time period, this gave a total of 20,481 points assessed within IBAs and 17,870 points in their buffers. We recorded land cover at each point once in each of three time periods: 1981–1994, 1995–2004 and 2005–2009. The years of sampling varied according to the availability of high quality, cloud-free images. Variation between sites in years of sampling was controlled in the analyses by modelling survival of natural land cover as a point-specific exposure period (see below).

For the years from 1981 to 2002, we used freely available imagery from Landsat (http://www.landcover.org; Tucker et al. (2004)). For the years 2003-2009, we used a combination of Landsat imagery (http://glovis.usgs.gov) and purchased Aster images. All images had a spatial resolution of 30 m. We excluded points for which no data were available (because of cloud, poor image quality or the failure of Landsat 7’s scan line corrector after 2003) in one or more of the sampling periods. We also excluded the small number of points in some IBA buffers that overlapped other PAs or IBAs.

The dominant land cover at each point in each time period was allocated to one of 11 broad categories: closed tree cover, open tree cover, mosaic of natural and
agricultural vegetation, shrub, herbaceous, tree and shrub crops, arable crops, open water, flooded shrub and herbaceous, urban, and bare (classification based on Di Gregorio & Jansen (2000)). Of these categories, we considered four as ‘non-natural’ land covers: mosaic of natural and agricultural vegetation, tree and shrub crops, arable crops and urban. The others were considered ‘natural’ land covers. For points at which the initial land cover was closed or open tree cover (forests), subsequent change to any other land cover was counted as conversion. For other natural land covers, only conversion to non-natural land covers was counted as conversion. Across all land cover types, natural and non-natural land covers could be separated with an estimated accuracy of c. 94%; further details on interpretation and validation are given in Beresford et al. (2013).

**Correlates of land-cover change**

In addition to the potential correlates of land-cover change used in the initial site matching (area, altitude, distance to roads, human population density and most extensive land cover), we obtained maps of market accessibility (Nelson 2008), elevation (Jenness et al. 2007), slope (USGS 2006), agricultural suitability (Fischer et al. 2002), cropland cover in 1990 (Ramankutty et al. 2008) and biome (Olson et al. 2001). We included a 4-level ‘class’ factor denoting whether a point fell within a protected IBA or its buffer, or an unprotected IBA or its buffer, and a covariate indicating the distance of each point from the edge of the IBA, with negative values indicating points within the IBA and positive values points outside (Beresford et al. 2013). A summary of the explanatory variables and their sources is given in Table 1. We included the initial land-cover category at each point as a categorical variable, and calculated the proportion of other points within 5 km of each point that were already dominated by non-natural land cover in the initial time period (1981-1994). Spatial data manipulation and processing was undertaken in ArcGIS 10.1. Human population density, slope and altitude were log-transformed prior to analysis.

**Data analysis**

Land-cover conversion was modelled as a Bernoulli process in a generalised linear mixed model framework using the ‘glmer’ package in R (R Development Core Team 2010). The natural land cover at each point was considered to have survived (if it remained as natural land cover) or not (if it was converted to a non-natural land
cover) during a point-specific exposure period, thus controlling for different periods of observation across points. Exposure in the case of points at which natural land cover was not converted was the number of years over which each point was observed (the period between the earliest and latest years of satellite imagery used for each point). For points at which land cover was converted between the first and second or between the second and third assessments, we assumed that loss occurred half way between the first assessment at which conversion was first recorded and the previous assessment, and calculated the exposure period accordingly.

A binary survival/conversion variable was fitted as the dependent variable and the point-specific exposure period (in years) fitted as a binomial denominator, allowing us to derive annual survival probabilities that were comparable across all points. Country and site (IBA code) were included in initial models as random effects, either in separate models or as a nested random effect in the same model. Once the best supported set of random effects had been decided using the ‘anova’ command in ‘glmer’, the best combination of fixed effects was assessed.

Given the large number of explanatory variables and the many plausible interactions between them, we adopted a pragmatic approach to model selection. First, we fitted all possible combinations of up to a maximum of five explanatory variables (excluding the selected random effects, which remained in all models, and without interactions) using the ‘dredge’ function of the R package ‘MuMIn’ (Bartoń 2012). These models were ranked by Akaike’s Information Criterion (AIC) and that with the lowest AIC was used as a base from which to assess support for more heavily parameterised models. To the 5-or-fewer-variable model selected, we next added the quadratic terms of any covariates in that model; these were retained if they significantly improved the model (reflected in an AIC of 2 or more units lower than the model without the quadratic term/s). We then added and removed each remaining explanatory variable in turn to select the best-supported model and compared each of these to the previous model using AIC and likelihood ratio tests, to assess whether fitting an extra variable to the previous model improved the fit. The process was repeated until the addition of any further variable could not reduce the AIC by at least 2. We then fitted a number of plausible interactions to the model, and compared the resulting candidate models using AIC and likelihood ratio tests. Finally, we assessed whether any simplification of the final model was justified by
comparing the AIC of the final model with the AICs of all subsets of that model that lacked each explanatory variable in turn. The relative importance of each predictor was also assessed from this comparison. Once a final model was adopted, we assessed its goodness-of-fit using the method of Nakagawa & Schielzeth (2013) with the ‘r.squaredGLMM’ command in ‘MuMln’.

RESULTS

When accounting for the random effect of IBA, the overall annual survival rate across all natural habitats was 0.996, equating to an overall survival rate of 88.7% over 30 years. The best supported model of five or fewer variables contained slope, human population density, the 4-level protection class, initial land cover type and the extent of already converted land within 5 km at the start of the observation period (Table 2). This model carried an Akaike weight relative to the other competing models of 1 and all other candidate models of five or fewer variables had ΔAIC > 37.5 with respect to this model. Likelihood ratio tests supported simplification of this model by reducing the 4-level protection class variable to a binary classification in which unprotected IBAs and the buffers of both protected and unprotected IBAs were all classed as ‘unprotected’, and protected IBAs were classed as ‘protected’. A model in which the quadratic term of slope was fitted received greater support than a model without, but the same was not true for the quadratic terms of human population density or the proportion of points within 5 km that had been converted to non-natural habitats by the start of the exposure period (Table 2). The addition of altitude and its quadratic term further improved the model (Table 2). The best supported model was that which also included interactions between protected status and slope, protected status and altitude, protected status and initial conversion and protected status and initial land-cover type (Table 2). Removal of each variable in turn from this model confirmed that no simpler model received greater support, although a model that lacked the interaction between protected status and altitude was within 2 AIC units. The marginal $R^2$ of the final model (ie the variation explained only by the fixed effects) was 0.265, and the conditional $R^2$ (fixed + random effects) was 0.382. This model indicated that the survival rate of natural land cover increased with increasing slope and altitude and with decreasing human population density, that it was higher in PAs
then in unprotected sites, that it varied between major land cover classes and that it was lower in already heavily converted areas. The interactions indicated that the effects on survival rates of slope, altitude, initial conversion and land cover types varied between protected and unprotected areas. Predicted values were generated for each parameter included in the final model by holding each of the other variables in the model, except the binary variable relating to protected status, at its mean or, in the case of factors, at its reference level. This revealed that PAs had lower rates of loss than unprotected areas, but the interactions in the model indicated that this pattern varied across the range of values of other parameters (Fig. 1). Survival of natural habitats increased with increasing slope and altitude, and at high values of both there was little difference between protected and unprotected areas. Survival declined with increasing human population density and with the amount of adjacent land that had already been converted by the start of the observation period. There was little effect of protected status on subsequent rates of loss of natural land-cover where land conversion of nearby cells was already high at the start of the exposure period (Fig. 1). Survival varied by major land-cover type, as did the relative benefits of protected status within land-cover types; closed and open forest suffered the greatest overall rates of loss but the difference between rates of loss in protected and unprotected IBAs was greater in these habitats than was the case in other habitats, indicating a greater impact of PAs in terms of avoided conversion (Fig. 2).

**DISCUSSION**

Our analyses identify a number of environmental and anthropic correlates of natural land-cover conversion, which have additive and sometimes interactive effects. Slope, altitude, human population density and initial adjacent conversion prior to the exposure period were all retained in the final model. In addition, as shown by Beresford *et al.* (2013), rates of conversion differed significantly between land cover types.

Our results also confirm many previous assessments (e.g. Andam *et al.* 2008; Gaveau *et al.* 2009; Selig & Bruno 2010; Geldmann *et al.* 2013; Butsic *et al.* 2015), including our own analysis of these data (Beresford *et al.* 2013), in showing that
protection significantly reduces conversion. Rates of conversion in protected areas were less than half those in unprotected areas, even when accounting for a range of covariates. However, the beneficial effects of protected areas are not universal (e.g. Western et al. 2009; Brun et al. 2015; Wendland et al. 2015), and the extent to which protection yields net conservation gains varies greatly, even within relatively small regions (Paiva et al. 2015).

We found that the net impact of protection, in terms of the difference in land-cover conversion rates between protected and unprotected areas (i.e. avoided conversion), declined with both slope and altitude (Fig. 1), such that on steeper slopes and at higher altitudes, the annual survival rates of natural land-cover in protected and unprotected sites converged. This was presumably because the increased natural protection offered by topography increased survival rates in unprotected areas to a level close to that recorded in protected areas, and confirms that the over-representation of protected areas in high and steep areas (Joppa & Pfaff 2009) reduces the potential impact of the protected area network as a whole.

Furthermore, by considering the interaction between protection and proximity of converted points we found that protected area designation did little to halt the loss of natural cover when sites were already heavily degraded. This is consistent with previous findings that protection is least effective in areas where the threat of conversion is greatest, and supports suggestions that degradation is a contagious process (e.g. Ahrends et al. 2010; Boakes et al. 2010; Brun et al. 2015). The lack of statistical support for an interaction between protected area status and human population density suggests, perhaps unexpectedly, that protected areas are as effective in heavily populated areas as in areas with low human populations. This result supports the assertion of Fisher (2010), who suggested that population growth and urbanization alone do not explain deforestation in Africa, as they do in other parts of the developing world.

Some land cover types were more susceptible to conversion than others, with both open and closed forest sustaining higher than average rates of decline (Fig. 2). This supports previous suggestions that forest is a particularly threatened habitat in Africa, with multiple threats that include clearance for agricultural land (including temporary rotational agriculture), small scale collection of firewood and commercial
logging (Buchanan et al. 2009). However, the difference between rates of loss in protected and unprotected forests, and therefore their impact in terms of avoided conversion, was greater than for other natural land cover types. Protection was more effective in closed, compared to open forests. This may reflect an inability to distinguish between forest types with naturally more open canopies, and areas of degraded forest which would naturally have closed canopies, using our methods and Landsat and Aster imagery. If so, the observed difference in effectiveness between open and closed forest would further support the negative correlation between threat and effectiveness.

Our results support the assertions of (among others) Andam et al. (2008) and Paiva et al. (2015) in showing that the effectiveness of protected areas can vary greatly along different environmental gradients, and between land cover types. As protected area networks are further developed, it is essential that all aspects of their impacts are considered: the species and habitats they contain, the ecosystem services they could conserve, the probable losses in the absence of protection, and how likely legal designation is to prevent those losses in practice. The variation in the effectiveness of protection shows the need to have objective, measurable targets on the effectiveness of protected areas, in addition to targets on the extent of their coverage (Woodley et al. 2012). Ideally, these would extend beyond assessment of land cover retention and conversion and include a range of metrics of effectiveness.

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REFERENCES


National Imagery and Mapping Agency (2000) Vector Map Level 0 (Digital Chart of the World)


Table 1. Variables used to model land cover change in and around 103 Important Bird Areas in Africa. Categorical variables are shown in italics.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Source</th>
<th>Description/units</th>
<th>Mean</th>
<th>Range/classes</th>
</tr>
</thead>
<tbody>
<tr>
<td>Distance from edge</td>
<td>This study</td>
<td>Kilometres from edge of protected area or (for unprotected sites) from edge of IBA</td>
<td>6.31</td>
<td>0 – 27.13</td>
</tr>
<tr>
<td>Protection class</td>
<td>IUCN &amp; UNEP-WCMC 2010</td>
<td>Four classes</td>
<td>NA</td>
<td>Protected IBA, unprotected IBA, buffer of protected IBA, buffer of unprotected IBA</td>
</tr>
<tr>
<td>Human population density</td>
<td>CIESIN &amp; CIAT 2005: GPWv3</td>
<td>Thousands of people per km² at 2.5 arc-minute (~5 km) resolution, adjusted to match UN totals (1990)</td>
<td>0.04</td>
<td>0 – 13.99</td>
</tr>
<tr>
<td>Market accessibility</td>
<td>Nelson 2008</td>
<td>Hours of travel time to nearest city of 50,000 people or more, at 30-arc-second (~1-km) resolution</td>
<td>5.74</td>
<td>0 – 43.52</td>
</tr>
<tr>
<td>Elevation</td>
<td>Jenness et al. 2007</td>
<td>Elevation above sea level in metres from SRTM 30 arc-second (~1-km) DEM</td>
<td>959.2</td>
<td>0 – 4774</td>
</tr>
<tr>
<td>Slope</td>
<td>USGS 2006</td>
<td>Mean slope in degrees derived from 300-m resolution SRTM data and averaged to 1 km</td>
<td>4.58</td>
<td>0 – 50.5</td>
</tr>
<tr>
<td>Agricultural suitability</td>
<td>Fischer et al. 2002</td>
<td>Eight-level suitability index, for rainfed crops, at 5-arc-minute (~10-km) resolution</td>
<td>4.50</td>
<td>1 – 8 (Very high suitability – not suitable)</td>
</tr>
<tr>
<td><strong>Cropland cover</strong></td>
<td>Ramankutty <em>et al.</em> 2008</td>
<td>Proportion cropland at 5-arc-minute (~10-km) grid resolution (1990)</td>
<td>0.11</td>
<td>0 – 0.69</td>
</tr>
<tr>
<td>-------------------</td>
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<td>------------------------------------------------------------------</td>
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</tr>
<tr>
<td><strong>Biome</strong></td>
<td>Olson <em>et al.</em> 2001</td>
<td>Simplified into 5 biome classes</td>
<td>NA</td>
<td>Wet forest, dry forest, grassland/savanna, montane grassland, arid</td>
</tr>
<tr>
<td><strong>Initial land cover</strong></td>
<td>This study: Landsat (TM) imagery</td>
<td>Visual interpretation into 7 “natural” land cover classes</td>
<td>NA</td>
<td>Closed tree cover, open tree cover, shrub, herbaceous, open water, flooded, bare</td>
</tr>
<tr>
<td><strong>Initial adjacent conversion</strong></td>
<td>This study: Landsat (TM) imagery</td>
<td>Proportion of points within 5 km of point with artificial land cover in initial time period (1981-1994)</td>
<td>0.10</td>
<td>0 – 0.96</td>
</tr>
</tbody>
</table>
Table 2. Model selection table, showing the relative support for the null model (random effects only), the best supported model of 5 or fewer variables (‘best-5-var’), the same model with quadratic terms fitted for one or more of the covariates, decided using AIC (‘best-5-var-quad’), the best supported model that added an extra variable to the previous model (‘best-6-var’), the same model with quadratic terms fitted for the covariate added in the previous step (best-6-var-quad), and the same model with the best supported combination of interactions, assessed by ΔAIC (‘best-6-var-quad-i’). No model with 7 or more variables received as much support from the data as the best supported 6-variable models. Variable abbreviations: S slope, A altitude, H human population density, I initial land-cover type, P protected area status, C proportion of surrounding points already converted to non-natural land-cover types by start of exposure period. IBA was fitted as a random effect in all models and is not shown. K = number of parameters estimated. Asterisks indicate interactions between variables whose main effects were also included in the model.

<table>
<thead>
<tr>
<th>Model</th>
<th>Terms</th>
<th>K</th>
<th>AIC</th>
<th>ΔAIC</th>
<th>Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>best-6-var-quad-i</td>
<td>P<em>A+P</em>C+P<em>S+ P</em>I+S^2+H</td>
<td>23</td>
<td>22060.2</td>
<td>0</td>
<td>1</td>
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<tr>
<td>best-6-var-quad</td>
<td>P+A+A^2+C+S+S^2+H+I</td>
<td>15</td>
<td>22101.9</td>
<td>41.7</td>
<td>0</td>
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<tr>
<td>best-6-var</td>
<td>P+A+C+S+S^2+H+I</td>
<td>14</td>
<td>22122.5</td>
<td>62.3</td>
<td>0</td>
</tr>
<tr>
<td>best-5-var-quad</td>
<td>P+C+S+S^2+H+I</td>
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<td>22139.4</td>
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Fig. 1. Variation in modelled values (+/- 1 s.e.) of annual survival rate of natural habitats with slope (log degrees), altitude (km), human population density (log people per km$^2$) and the amount of conversion (proportion) that had already taken place by the start of the observation period. Grey: unprotected sites, black: protected areas. Curves were generated using the best-supported model in Table 2 fitted to data in which all variables except the covariate of interest and protected area status were constrained to their mean or reference values.
Fig. 2. Estimates of annual survival of each of 7 broad natural land-cover classes, and of all land-cover types combined (+/- 1 s.e.). Grey: unprotected sites, black: protected sites. Estimates were derived using the best-supported model in Table 2 fitted to data in which all variables except land-cover class and protected area status (or just protected area status, in the case of “All”) were constrained to their mean or reference values.