Original article

Prognostic gene expression signature for high-grade serous ovarian cancer

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Running head: prognostic signature for high grade serous ovarian cancer

1 Abstract

2 Background

Median overall survival (OS) for women with high-grade serous ovarian cancer (HGSOC) is approximately four
years, yet survival varies widely between patients. There are no well-established, gene expression signatures
associated with prognosis. The aim of this study was to develop a robust prognostic signature for overall
survival in HGSOC patients.

7 Patients and methods

8 Expression of 513 genes, selected from a meta-analysis of 1455 tumours and other candidates, were measured

9 using NanoString technology from formalin-fixed, paraffin-embedded (FFPE) tumour tissue from 3,769 women

10 with HGSOC from multiple studies. Elastic net regularization for survival analysis was applied to develop a

11 prognostic model for 5-year OS, trained on 2702 tumours from fifteen studies and evaluated on an

12 independent set of 1067 tumours from six studies.

13 Results

14 Expression levels of 276 genes were associated with OS [false discovery rate (FDR) < 0.05] in covariate-adjusted

single gene analyses. The top five genes were TAP1, ZFHX4, CXCL9, FBN1, and PTGER3 (P << 0.001). The best

16 performing prognostic signature included 101 genes enriched in pathways with treatment implications. Each

- 17 gain of one standard deviation in the gene expression score (GES) conferred a greater than two-fold increase in
- 18 risk of death [HR = 2.35 (2.02, 2.71); *P* ≪ 0.001]. Median survival by GES quintile was 9.5 (8.3, --), 5.4 (4.6, 7.0),
- 19 3.8 (3.3, 4.6), 3.2 (2.9, 3.7) and 2.3 (2.1, 2.6) years.
- 20

22 Conclusion

- 23 The OTTA-SPOT (Ovarian Tumor Tissue Analysis consortium Stratified Prognosis of Ovarian Tumours) gene
- 24 expression signature may improve risk stratification in clinical trials by identifying patients who are least likely
- 25 to achieve 5-year survival. The identified novel genes associated with the outcome may also yield
- 26 opportunities for the development of targeted therapeutic approaches.

27

Key words: high grade serous ovarian cancer, gene expression, prognosis, overall survival, formalin fixed
 paraffin embedded

30

31 Highlights

- A gene expression signature for high-grade serous ovarian cancer prognostic for two and five-year
 overall survival (OS).
- The 101 gene expression signature performs substantially better than age and stage alone.
- Median survival by quintile was 9.5, 5.4, 3.8, 3.2 and 2.3 years.
- The top five genes associated with OS were TAP1, ZFHX4, CXCL9, FBN1, and PTGER3 (P << 0.001).

37 Introduction

38 Epithelial ovarian cancer (EOC) causes approximately 125,000 deaths globally every year, and long-term 39 survival rates have changed little in the past three decades[88]. Approximately 70% of women with EOC are 40 diagnosed with advanced stage disease (stages III/IV), and fewer than 50% will survive more than 5 years[89]. 41 There are five major EOC histotypes – high-grade serous; low-grade serous; endometrioid; clear cell and 42 mucinous[90]. High-grade serous ovarian cancer (HGSOC) comprises about two-thirds of cases, is responsible 43 for most deaths and is characterized by profound genomic and clinical heterogeneity. 44 The most informative prognostic factors for HGSOC are International Federation of Gynecology and Obstetrics 45 (FIGO) stage, residual disease following debulking surgery[91], BRCA1 or BRCA2 germline mutation[92, 93] and 46 tumour-infiltrating lymphocyte scores[94, 95]. Patients with HGSOC who carry a loss-of-function germline 47 mutation in BRCA1 or BRCA2 have an increased sensitivity to platinum-based chemotherapy and PARP 48 inhibitor treatment[96, 97] and a medium-term survival advantage[92]. However, the frequent development 49 of drug resistant disease[93] limits the effectiveness of current therapies. 50 Gene-expression data have been used to define four tumour molecular subtypes of HGSOC (C1/mesenchymal, 51 C2/immune, C4/differentiated and C5/proliferative)[98, 99]. Using transcriptome-wide data from fresh frozen 52 tissues, The Cancer Genome Atlas (TCGA) project used 215 tumours to identify an overall survival (OS) 53 expression signature of 193 genes that has been validated on three other HGSOC gene expression data 54 sets[99].

Despite these findings, gene expression biomarkers have not been implemented clinically owing to several important shortcomings. The majority of the individual markers comprising the 193 gene signature were not statistically significant across all studies, suggesting that the signature may not be robust. The sample sizes in other discovery efforts have been too small for robust statistical inference [99]. Also, previous studies used

- 59 fresh frozen samples, resulting in logistic and cost barriers to examining large clinically relevant data sets, and
- 60 translation to the clinical setting.
- 61 The aim of this study was to identify a robust and clinic-ready prognostic HGSOC profile that can be applied to
- 62 formalin fixed paraffin embedded (FFPE) tumour tissue.

63 Patients and methods

- 64 Twenty studies provided pre-treatment, FFPE tumour samples from 4,071 women diagnosed with HGSOC
- 65 (Supplemental Table S1). All HGSOC cases with available tissue were included. During this time period HGSOC
- 66 patients were treated with chemotherapy (carboplatin and paclitaxel) after primary debulking surgery. Study
- 67 protocols were approved by the respective Institutional Review Board / ethics approval committee for each
- 68 site (Supplemental Table S1).
- A schematic of the overall study design is shown in Figure 1. There were four main components: gene
- 50 selection, gene-expression assay, development of prognostic gene signature in a training set and validation of
- 71 prognostic signature in an independent validation set.

72 Gene selection

73 Candidate prognostic genes were identified by carrying out an individual participant meta-analysis of six 74 transcriptome-wide microarray studies[98-103], which included tumour samples from 1,455 participants. 75 Gene expression association with overall survival was evaluated by Cox proportional hazards regression 76 adjusted for molecular subtype (Supplemental Table S2). In total, 200 genes from the meta-analysis, most 77 achieving a permutation-based FDR[104] of less than 0.05, and an additional 313 candidate genes based on the 78 literature and unpublished data were selected (Supplemental Tables S3 and S4, Figure S1; for more details see 79 Supplemental Material). Five genes, RPL19, ACTB, PGK1, SDHA, and POLR1B, were included as house-keeping 80 genes for normalization.

81 Gene expression assay in study participants samples

82 FFPE tumour samples were processed with the NanoString nCounter technology at 3 different locations, 83 Vancouver, Los Angeles and Melbourne. A control set of 48 FFPE tumour samples were run at each location 84 and the average intraclass correlation coefficient (ICC) was 0.987. Approximately 2 percent of the samples 85 were run in duplicate and the average Spearman correlation r² was 0.995. Single-patient classification methods 86 were used with reference samples to control for batch effects [105]. The data in this publication have been 87 deposited in NCBI's Gene Expression Omnibus[106]; GEO Series accession number GSE132342 88 (https://www.ncbi.nlm.nih.gov/geo/query/acc.cgi?acc=GSE132342). Three thousand eight hundred and 89 twenty-nine samples passed quality control of which 3,769 had survival data and assessable gene expression 90 for 513 genes. Data can be found in NCBI GEO: Accession numbers GSE132342 and GPL26748.

91 Overall survival analysis of individual genes

92 Samples that contributed to the meta-analysis data set (n=211) were removed from subsequent selected 93 analyses to enforce independence of study samples between the gene selection and final survival analysis. 94 Time-to-event analyses were carried out for OS with right-censoring at 10 years and left-truncation of 95 prevalent cases. Associations between log-transformed normalized gene expression and survival time were 96 tested using likelihood ratio tests with Cox proportional hazards models adjusted for age, race, and stage, and 97 stratified by study. Patients with missing race or stage information were assigned to 'unknown' categories. Age 98 was modeled using a B-spline with a knot at the median age, which yielded a better fit than using knots at 99 quartiles or categorical variables. Stage was dichotomized into early (International Federation of Gynecology 100 and Obstetrics [FIGO] stage I/II) and advanced (FIGO stage III/IV). Genes were scaled to have a standard 101 deviation of one, so hazard ratios correspond to a change of one standard deviation. A Benjamini-Hochberg 102 (BH) false discovery rate (FDR) of less than 0.05 was used to identify notable associations. Since the expression 103 of genes can be correlated, an analysis of correlated genes was performed using data from TCGA. Advanced

104 stage ovarian cancer usually has disease spread throughout the abdomen, therefore sensitivity analyses were

105 performed to assess effects of the anatomical location of tumor samples included in the study by removing

106 observations corresponding to samples known to be extraovarian (n = 437).

107 Prognostic signature development and validation

108 Studies were initially randomized to training set (N = 14) and validation set (N = 6). The TRI study was 109 randomized to the validation set, but, because 107 of the samples were part of the meta-analysis data used for 110 gene selection, the study was split, so those 107 samples were included in the model training data set. Thus 2,702 samples from 15 studies were used for model training and 1,067 samples from 6 studies were used for 111 112 validation (Supplemental Table S1). In the training set, four modelling approaches (stepwise regression, elastic 113 net regularized regression, boosting and random survival forests) were applied to construct competing gene 114 expression-based biomarkers. Each was evaluated in the training data using 10-fold cross-validation for its 115 prognostic value for OS at two and five years of follow-up using an area under the curve (AUC) measure 116 derived from receiver operator characteristic (ROC) analysis (see Supplemental Material for additional details). 117 The best performing method, elastic net regularized regression, was applied to the full training set to 118 determine the final gene signature and scoring method, which was then evaluated using the independent 119 testing set. All models were constrained to include age and stage, where age was modelled as categorical 120 based on quartiles of the training dataset with groups: less than 53 years old, 53 to 59, 60 to 66, and 67 or 121 greater. Stage was modelled as described above for the OS individual gene analysis.

122 **Results**

123 Association of expression of individual genes with OS in HGSOC.

In a gene-by-gene analysis of the full data set adjusted for age, race, and stage, and stratified by study, 276 of
 the 513 selected genes were associated with OS (FDR < 0.05). Of these, 138 were selected from the meta-

126 analysis of six published microarray studies (Supplemental Table S2)[98-103] and 144 from candidate gene 127 approaches (Supplemental Tables S5 and S6). Hazard ratios (HR) for one standard deviation change in gene 128 expression ranged from 0.84–1.19, with multiple genes exhibiting associations at very stringent significance 129 levels (e.g., 19 genes with $P < 1 \times 10^{-8}$; Supplemental Tables S5 and S6). The five most significant genes were 130 TAP1, ZFHX4, CXCL9, FBN1 and PTGER3 (Table 1). We did not find extensive evidence of high co-expression 131 between these five genes and genes measured in TCGA project (Supplemental Table S7). In sensitivity analyses 132 we found that excluding samples from omentum and other extra-ovarian sites did not substantially affect the 133 results (Supplemental Tables S8 and S9).

134 Development of a novel prognostic gene signature

135 The four predictive modelling approaches that were evaluated in the training data using 10-fold cross-136 validation yielded median AUCs that ranged from 0.69 to 0.73 for two-year OS and 0.69 to 0.74 for five-year 137 survival (Supplemental Figure S2) with better prediction of 5-year overall survival than at two years. The 138 elastic net approach yielded the highest median AUC for both two and five-year OS and was selected for final 139 development of the signature. Using the model on the full training data set resulted in a prognostic signature 140 of 101 genes in addition to age and stage (Supplemental Table S10). Of these, 66 genes were associated with OS (FDR < 0.05) in the single gene models. There was no obvious subset of signature genes that performed as 141 142 well or nearly as well as the full 101 gene signature (Supplemental Figure S3).

Performance of the signature including age and stage was AUC = 0.69 (95% CI 0.65-0.73) and 0.75 (95% CI 0.72-0.78) for 2-yr and 5-yr OS, respectively (Figure 2, Figure 3, Supplemental Figure S4). This was substantially better than age and stage alone with AUC = 0.61 (95% CI 0.57-0.65) and 0.62 (95% CI 0.59- 0.67) for 2-yr and 5yr OS, respectively), particularly for the 5-yr OS outcome with non-overlapping 95% CI. One standard deviation change in the gene expression score was associated with a hazard ratio of 2.35 [95% CI = (2.02, 2.71); *P* = $5.1x10^{-31}$], and median survival varied substantially across quintiles of the gene expression score [9.5 (8.3, ---),

5.4 (4.6, 7.0), 3.8 (3.3, 4.6), 3.2 (2.9, 3.7) and 2.3 (2.1, 2.6) years, respectively, from smallest to largest quintile;
Table 2].

151 For a subset of cases, there was clinical and experimental data for known prognostic factors. All samples had 152 molecular subtype classification (Talhouk et al. submitted), residual disease was known for 1,771 cases, 153 primary treatment for 687, germline BRCA mutation status for 904, and nuclear CD8 TIL counts[95] for 1,111 154 (Supplemental Table S11). When examined by quintile of gene expression score there were differences, as 155 expected, for each of the known prognostic factors, including age and stage that were included in the model 156 (Table 3). However, in sensitivity analyses, applying the signature to specific patient groups, a robustness of 157 stratification was demonstrated, suggesting that the prognostic power of the signature is not explained by the 158 individual factors, residual disease, treatment, BRCA status, or CD8 score (Figure 3, Supplemental Figures S5-159 S7). The signature score showed modest differences by molecular subtype (Supplemental Figure S8), and 160 adjusting for molecular subtype in the Cox analysis resulted in only minor changes to the HR estimates for 161 signature quintiles (Table 2). The signature was shown to be prognostic within a homogenous group of 316 162 stage 3C cases with no residual disease, within early stage cases (FIGO 1a and 1b), and within patients whose 163 samples were collected from the omentum (Supplemental Figures S9-S10). Analysis of the signature score for 164 paired ovary and omental tissue from 42 of the cases showed a highly significant Pearson correlation coefficient, r = 0.79 ($p = 5.4 \times 10^{-10}$) (Supplemental Figure S11). 165

A geneset enrichment analysis was performed for the 101 genes in the signature, as well as for genes correlated with signature genes achieving r2 > 0.75 (Supplemental Table S12). For the correlated gene analysis, the three most significant pathways involved the immune system, including the adaptive immune system and cytokine signalling. A further ten immune pathways were significantly enriched and included interferon signalling, innate immune system, and TCR signalling and antigen presentation pathways. Restricting to the signature genes only, there was also enrichment in the immune system, but the top two pathways were PI-3K cascade and GPCR ligand binding. Four other pathways were related to the cell cycle and mitosis, with the

- 173 remaining enriched for fibroblast growth factor receptor (FGFR) and epidermal growth factor receptor (ERRB)
- signalling, and one pathway related to homologous combination repair.

175 **Discussion**

In a large-scale study of HGSOC patients, we identified a 101 gene expression signature able to predict
clinically relevant differences in OS. Using methods that are both economical and applicable to standard
clinical sampling techniques, we showed that the signature performs substantially better than age and stage
alone for prognosis of both two and five-year OS. The number of patients and samples included in this study is
an order of magnitude greater than previous comparable studies of gene expression and OS in HGSOC
patients[99, 107, 108]. Thus, we have been able to more precisely quantify the prognostic value of gene
expression.

183 We report definitive associations between OS and expression of 276 genes. Of the five most significant genes (TAP1, ZFHX4, CXCL9, FBN1, and PTGER3), four have been previously reported to be associated with survival in 184 185 HGSOC. The top prognostic gene, TAP1, is involved in the antigen presenting pathway. Expression was reduced 186 in metastatic HGSOC, positively associated with OS[109] as observed here, and linked to tumour regression in 187 response to treatment[110]. Also, hypomethylation of TAP1 was associated with improved time to disease 188 recurrence[111]. CXCL9 is a chemokine that mediates the recruitment of T-cells to solid tumours[112]. High 189 expression of intratumoural CXCL9 was associated with higher OS[113] and higher lymphocytic infiltration, 190 which is also a robust prognostic factor in HGSOC[95, 98, 114] and a feature of the immunoreactive HGSOC 191 molecular subtype[98]. CXCL9 has also been proposed as a therapeutic target due to evidence that it inhibits 192 angiogenesis and promotes antitumour adaptive immunity[115-117]. Strikingly, the signature was able to 193 further refine prognostic groups within patients with high TIL counts suggesting that CXCL9 and TAP1 194 expression may be strong indicators of immune competency in HGSOC.

FBN1 is an extracellular matrix (ECM) protein previously found to be a biomarker associated with early
recurrence in ovarian cancer patients who are initially sensitive to chemotherapy[118] and strongly correlated
with desmoplasia in HGSOC. The prostaglandin E2 receptor *PTGER3* is expressed in ovarian tumour cells and is
associated with relapse-free survival[119]. In contrast, *ZFHX4* does not have previous associations with HGSOC.

199 Associations between the expression of specific genes in tumour tissues and OS in HGSOC patients may 200 suggest new drug targets and lead to insights into biological variation in treatment response. For example, 201 cases in the Q5 quintile with the poorest outcome had increased expression of IGF2, FGFR1, and MYC, a 202 possible argument for the use of IGFR1, FGFR, Bromodomain (MYC), or a combination of PARP and CDK4/6 203 inhibitors (MYC) [33]. More immediately, the signature may help clinicians identify patients most in need of 204 intervention, patients that could potentially benefit from neo-adjuvant chemotherapy (NACT). Alternatively, in 205 clinical trials it could be used to stratify randomization by patients' risk, thereby reducing heterogeneity within 206 subgroups and increasing heterogeneity between subgroups. The signature will be incorporated into future 207 prospective clinical trials to determine if it can predict response to specific treatments.

208 Measurement of the signature required standard FFPE tissue used in routine histopathology. Also, data 209 preprocessing and normalization were conducted on an individual level, thus translatable to a general patient 210 population. That is, 5-year OS prognosis of future patients can be evaluated against the patient population 211 reported here by i) following the same steps described here for generating the normalized gene expression 212 data, 2) computing an individual signature score, and 3) assigning an HR based on the score or comparing it to 213 the reported quintiles (Supplemental Material). NanoString gene expression is highly reproducible as seen by 214 our quality control metrics (Supplemental Material) and the FDA approval of the ProSigna test for breast 215 cancer.

The question of heterogeneity by ancestry or ethnicity was beyond the scope of this study but should be
 pursued in future research. Another important question is whether molecular subtype can improve biomarker

performance. A substantial proportion of signature genes were identified by the subtype adjusted metaanalysis, suggesting that the strong performance of the signature is not solely attributable to differences
among molecular subtypes. Additionally, all of the individual genes used in the molecular subtype classification
were included in development of the signature.

Although the cases received chemotherapy, the FFPE samples used in this study were chemo-naïve, as few patients had NACT during the calendar period in which these samples were collected. Because the signature appears to be prognostic in omentum samples, future studies may assess the value in NACT patients, using pre-treatment omental biopsies or post treatment tumour samples. Future work will also address if the signature can predict platinum-refractory patients.

We have developed a robust prognostic signature for HGSOC that can be used to stratify patients and identify those in need of alternative treatments. Gene set enrichment analysis applied to the signature indicates an important role for the immune system in overall survival and supports further investigation of immune-therapy in ovarian cancer. More generally, the identification here of high-confidence prognostic genes may lead to new hypotheses for targeted treatments.

232

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377 Figure Legends

- 382 Figure 1. Schematic of study design. * The TRI study was split across the training and validation sets due to 107
- 383 samples overlapping with the meta-analysis.



Figure 2. ROC curves for prognostic performance of the gene expression signature in independent HGSOC
patients (testing data). There was no overlap between studies or patient data used to develop models (training
data) and compute ROC curves and AUC values shown here (testing data). All models included age and stage as
described in Methods. TP denotes the true positive rate (sensitivity) and FP denotes the false positive rate (1 –
specificity).



FP

Figure 3. KM curves of overall survival for patients A) in the training and B) testing sets. Patients were assigned
 to quintiles (Q1-Q5) of the signature score including age and stage. Shaded areas indicate 95 percent
 confidence regions, only included for plots representing larger sample sizes. Due to limited sample size, the
 following plots represent all such patients in the entire data set, training or testing, C) no macroscopic residual
 disease after debulking surgery, D) primary treatment ≥ 4 cycles of IV carboplatin AUC 5 or 6 & paclitaxel 135
 or 175 mg/m² every 3 weeks (actual dose known or presumed), E) *BRCA1* or *BRCA2* germline mutation, and F)
 CD8 > 19.



Table 1. Hazard ratios and 95% CIs for top 5 prognostic genes in covariate-adjusted single gene analyses.

Gene	HR (95% CI)	Р	Selection	Correlated gene*	rs
TAP1	0.84 (0.80, 0.87)	8.3x10 ⁻¹⁸	Meta	PSMB9	0.89
ZFHX4	1.19 (1.14, 1.25)	1.4x10 ⁻¹⁵	Meta	LOC100192378	0.74
CXCL9	0.85 (0.82, 0.88)	1.8x10 ⁻¹⁵	Meta and candidate	CXCR6	0.89
FBN1	1.18 (1.13, 1.24)	4.2x10 ⁻¹⁴	Candidate	SPARC^	0.91
PTGER3	1.18 (1.13, 1.24)	1.2x10 ⁻¹³	Meta	COL8A1	0.67

*Most correlated gene according to Spearman's rank correlation coefficient, *r*_s, computed in The Cancer
 Genome Atlas (TCGA) Ovarian Serous Cystadenocarcinoma RNA-seq data set.

405 • SPARC was included in this project and was less significant.

Table 2. Hazard ratios and 95% CIs for quintiles of the gene expression signature score in validation data.

	Quintile	N	Deaths	Median Survival*	HR (95% CI)	Adjusted for <u>Age and Stage</u> HR (95% CI)	Adjusted for M. Subtype <u>Age and Stage</u> HR (95% CI)
	Q1	214	81	9.47 (8.32,)	0.44 (0.33, 0.58)	0.34 (0.22, 0.55)	0.37 (0.23, 0.59)
	Q2	213	117	5.38 (4.63, 6.97)	0.73 (0.57, 0.93)	0.71 (0.55, 0.91)	0.74 (0.58, 0.96)
	Q3	213	145	3.80 (3.34, 4.60)			
	Q4	213	158	3.23 (2.85, 3.68)	1.56 (1.25, 1.96)	1.56 (1.24, 1.97)	1.56 (1.24, 1.96)
	Q5	214	179	2.27 (2.09, 2.62)	2.23 (1.78, 2.78)	2.11 (1.67, 2.67)	2.07 (1.63, 2.61)

409 *Median survival (95% CI) in years for patients in the validation set.

- 410 **Table 3.** Clinical data for the 3769 patients that passed quality control and the percentage of patients in each
- 411 quintile of the gene expression score.

	Total	Q1	Q2	Q3	Q4	Q5	p-value
N	3769	754	754	753	754	754	
median survival (years)	4.1	9.5	5.4	3.8	3.2	2.3	
% 5-year survival	41	75	57	39	25	10	
Age median	63	58	57	61	64	70	
Age range	25-89	39-78	25-86	36-82	27-89	39-86	
Age quartile q1	894	30.8	31.3	20.0	13.4	4.5	<1x10 ⁻⁵⁰
Age quartile q2	838	21.5	20.0	22.9	21.2	14.3	
Age quartile q3	961	16.0	20.2	21.4	23.6	18.7	
Age quartile q4	1076	13.5	10.4	16.4	21.3	38.5	
FIGO stage I / II	607	97.4	2.6	0.0	0.0	0.0	<1x10 ⁻⁵⁰
FIGO stage III/IV	3067	3.8	23.0	24.1	24.4	24.6	
Primary chemo* 1	136	16.2	22.1	23.5	19.1	19.1	0.163
Primary chemo* 2	190	16.3	20.0	21.6	22.1	20.0	
Primary chemo* 3	361	11.1	16.9	22.4	20.5	29.1	
Residual disease No	614	32.4	22.1	17.8	15.5	12.2	<1x10 ⁻⁵⁰
Residual disease Yes	1157	6.0	19.2	24.1	24.5	26.2	
germline BRCA1 mutation	130	23.8	31.5	26.2	11.5	6.9	2.22x10 ⁻⁷
germline BRCA2 mutation	71	28.2	26.8	18.3	18.3	8.5	
germline no mutation	663	19.6	16.7	18.7	20.7	24.3	
CD8 TIL score 0	192	19.8	14.6	12.5	21.4	31.8	2.46x10 ⁻¹⁴
CD8 TIL score 1-2	186	18.3	14.0	18.8	21.5	27.4	
CD8 TIL score 3-19	515	19.8	24.1	20.8	17.9	17.5	
CD8 TIL score >20	218	34.4	31.2	16.5	11.5	6.4	
Molecular subtype C1.MES	1105	5.4	10.4	20.7	27.4	36.0	<1x10 ⁻⁵⁰
Molecular subtype C2.IMM	907	23.2	28.8	21.2	16.2	10.7	
Molecular subtype C4.DIF	1144	32.6	25.5	17.9	12.8	11.2	
Molecular subtype C5.PRO	613	18.1	14.0	20.7	25.8	21.4	
FIGO stage 1A & 1B	111	96.4	3.6	0.0	0.0	0.0	<1x10 ⁻⁵⁰
FIGO stage 3C	1979	3.1	23.7	24.6	24.1	24.6	<1x10 ⁻⁵⁰
FIGO stage 3C Residual	316	6.3	31.0	24.4	20.9	17.4	6.24x10 ⁻⁴⁵
FIGO stage 3C Residual	846	2.6	21.5	25.3	24.6	26.0	

412 Q1 is the quintile with the best survival and Q5 the worst survival. Samples with missing data are reported in

413 Supplementary Table S11. P-values for BRCA1/2 mutation status were calculated for BRCA1 or BRCA2 mutation

414 vs no mutation. * Treatment: 1 = known to have received first line chemotherapy treatment of ≥ 4 cycles of IV

415 carboplatin AUC 5 or 6 & paclitaxel 135 or 175 mg/m² every 3 weeks. 2 = known to have received first line

416 chemotherapy treatment of \geq 4 cycles of IV carboplatin & paclitaxel 3-weekly but at doses presumed to be

417 carboplatin AUC 5 or 6 & paclitaxel 135 or 175 mg/m². 3 = all remaining cases with chemo regimens that do

418 not fit criteria 1 or 2 and include unknown or no chemotherapy.

420 Appendix

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511 Supplemental Methods

512 Gene selection based on transcriptome-wide meta-analysis

513 A meta-analysis was conducted of six transcriptome-wide microarray studies, including 1,455 participant's 514 tumors, to investigate the role of gene expression in HGSOC OS (Supplemental Table 2)[98-103]. The objective 515 was to identify prognostic genes whose effects were consistent across studies and were not surrogates for 516 molecular subtype. The total number of genes tested was 15,345, however, the number of genes in each study 517 differed due to differing microarray platforms used across the studies as well as different data processing and 518 quality control work flows. Expression data were normalized, batch corrected, and extreme outliers removed 519 on a probe-specific level for each study. Extreme outliers were defined as values greater than 2.5 times the 520 interquartile range from the upper or lower quartile, under the constraint that no more than three percent be 521 classified as such. That is, no more than three percent of observations were removed as outliers. According to 522 the microarray platform design, some genes are represented by more than one expression feature. To yield a 523 single expression feature per gene, principal components analysis (PCA) was applied to probe sets within each 524 study, taking the first PC to represent the gene. This approach is similar to that used for the The Cancer 525 Genome Atlas (TCGA) unified gene expression data[99]. To reduce the dependency of the identified genes on 526 the analytic approach, four types of analyses were conducted to evaluate genes for selection.

1) Consistency of OS association across studies. Cox proportional hazards regression was conducted separately for each study, adjusting for age, stage and study-provided molecular subtype (C1/mesenchymal, C2/immune, C4/differentiated and C5/proliferative). The median of the study-specific p-values for association between gene expression and prognosis was used as the statistic for an omnibus test across studies. Missing p-values were set to one. Consistency in direction of effect was accounted for by setting the median p-value to one if the signs of the effects differed for any of the studies with p-values equal to or smaller than the median pvalue. To account for multiple-testing and possible unknown characteristics of the null distribution of the

534 median p-value, we employed a permutation-based false discovery rate (FDR) approach[104] with 100 535 replicate permutations, considering the median p-value to be our test statistic. Permutation analyses were 536 conducted by permuting sample labels on the expression data within each study, while maintaining the 537 observed relationships between genes and between the outcome and adjustment covariates, including 538 molecular subtypes. Thus, for each gene and each of the 100 permutations, a median p-value was computed 539 under the null by grouping the permutation p-values according to the permutation index. This analysis resulted 540 in 115 genes (Supplemental Table S2) identified at FDR < 0.05 level (Supplemental Figure S1. A), all of which 541 were selected for the follow-up study. Note that this significance threshold corresponds to a median p-value 542 that is slightly greater than 0.05. Nevertheless, this threshold defines statistical significance because it is 543 evaluated against the distribution of median p-values under the null hypothesis of independence between 544 gene expression and outcome.

2) *Stratified analysis of marginal effects*. A Cox model was fitted and a likelihood ratio test (LRT) conducted for each gene, adjusted for age, stage, and molecular subtype (differentiated, immunoreactive, mesenchymal, and proliferative)[120], and stratifying on study. The result was an additional 12 genes at an FDR significance level of 0.05.

3) Evidence of interaction with molecular subtype on OS. For each gene, a stratified Cox model adjusted for
covariates was fitted as in (1), but molecular subtype×gene interaction terms were included to identify genes
whose effects differed across subtypes. LRTs were conducted for interaction terms, yet no additional genes
were identified in this analysis at the 0.05 FDR significance level.

4) Stratified analysis with multi-degree-of-freedom tests of main effects and interactions. For fitted models in
(3), LRTs were conducted to jointly assess gene main effects and interactions. This analysis yielded an
additional eight genes.

An additional 65 genes were added that were suggestive in one of the meta-analysis described above, but that did not meet the 0.05 significance level, based on evidence in prior literature and public knowledge databases such as MSigDB and REACTOME. In total 200 genes were selected from the meta-analysis (Supplemental Table S3).

560 Selection of additional candidate genes

561 An additional 304 candidate genes were selected based on one of the following criteria (a) evidence of 562 association with prognosis or potential drug targets from the literature, (b) residing within a 1 MB region of a 563 potential survival GWAS hit $p < 5x10^{-6}$ and showing a survival association in the publicly available TCGA data (c) 564 utility in molecular subtype classification (Talhouk et al, in preparation), (d) other specific hypothesis. Five 565 genes, RPL19, ACTB, PGK1, SDHA, and POLR1B, were included as house-keeping genes for normalisation[105]. 566 An additional six genes, TBP, GAPDH, KIF3B, GUSB, BMS1, and RPL41, were included to evaluate consistency 567 with previous codeset analysis but were not used in the normalisation[105]. Finally, ten genes were selected as 568 a "tagging" approach to increase representation of the gene expression patterns of other genes that are 569 correlated using the methods in Rudd et al[121]. For this study, we chose a threshold of 99% correlation, 570 observed in all four of the largest publicly-available HGSC ovarian cancer gene expression datasets[98, 99, 103, 122]. We determined that the 503 genes already selected included 99% correlated gene expression 571 572 information for an additional 2,617 genes. Another 10 genes were selected in order to maximize gene 573 expression data in other parts of the transcriptome that were not represented, and these 10 genes represent 574 gene expression for another 49 genes. Seven of the candidate genes overlapped with genes selected from the 575 meta-analysis, therefore 313 additional genes were added to the custom code set (Supplemental Table 4).

576 Single gene analysis of associations with OS.

577 Time-to-event analyses were carried out for OS with right-censoring at 10 years and left-truncation of
578 prevalent cases. For most genes, the association between log-transformed normalized gene expression for

579 each gene and survival time was evaluated using Cox proportional hazards models applied to the full data set. 580 However, in the analysis of those genes that were selected due to the meta-analysis results, the 211 cases that 581 were also represented in the meta-analysis data set were excluded. All single-gene models were adjusted for 582 age, race, and stage, and stratified by study. Patients with missing race or stage information were assigned to 583 'unknown' categories. Age was modelled using a B-spline with a knot at the median age, which yielded a 584 better fit than using knots at quartiles or categorical variables. Stage was dichotomized into early 585 (International Federation of Gynecology and Obstetrics [FIGO] stage I/II) and advanced (FIGO stage III/IV). 586 Expression of each gene was scaled to have a standard deviation of one, so hazard ratios correspond to a 587 change of one standard deviation. A Benjamini-Hochberg (BH) false discovery rate (FDR) of less than 0.05 was 588 used to identify notable associations. Advanced stage ovarian cancer usually has disease spread throughout 589 the abdomen, therefore sensitivity analyses were performed to assess effects of the anatomical location of 590 tumor samples included in the study by removing observations corresponding to samples known to be 591 extraovarian (N = 437).

592 Gene-expression profiling

593 For RNA extractions, all sites performing NanoString reactions followed standard operating procedures 594 outlined in advance. RNA was extracted from FFPE tumor samples using the Qiagen miRNeasy FFPE kit and 595 were processed with the NanoString nCounter technology using a custom codeset. Briefly, each day sites 596 processed a maximum of 24 samples. Our standard operating procedure called for 500ng of total RNA, as 597 measured from NanoDrop, combined with hybridization buffer and a custom NanoString reporter and capture 598 CodeSet allowing hybridization for exactly 16 hours (short-hyb, 12 samples per day) or 20 hours (long-hyb, 12 599 samples per day) at 65°C in a pre-heated thermal cycler. Immediately at the end of the prescribed 600 hybridization period samples were processed on an nCounter prep-station (NanoString) following standard 601 procedures. Loaded cartridges (12 samples) were scanned at maximum resolution on an nCounter Digital 602 Analyser (NanoString). The BC Cancer (Vancouver) site performed scanning on a Gen1 Digital Analyzer, while

both USC (Los Angeles) and PMC (Melbourne, sometime denoted as AOC or Australian Ovarian Cancer study)
performed scanning on Gen2 Digital Analyzers. Relevant variables including processing date, operator, site,
and hybridization time were recorded/embedded into specimen information (CDF) and data files (RCC). In
addition to unique HGSOC samples a number of controls and sample replicates were run at all sites to enable
evaluation of data quality.

608 Reference Pools. To monitor for technical bias across sites and allow for cross-CodeSet comparisons 35, we ran 609 3 distinct control RNA pools. This reference-based normalization strategy is considered best practice for 610 development of NanoString based clinical tests and is similar to the implementation already in use for 611 Prosigna[123] and a number of other in development tests. Pools consisted of high-quality RNA from fresh-612 frozen ovarian cancer samples believed to be representative of all molecular subtypes and/or various ovarian 613 cancer histotypes. Pools were assembled en-mass and aliquoted (5ul, 100ng total RNA) for single use without 614 multiple freeze thaws at all sites. Control aliquots were stored at -80°C until ready for use and shipped on dry 615 ice to all processing sites. Pool1 was run approximately every month at each site. Pool2 and Pool3 were run 616 alternatingly, every other month, at each site.

617 Cross-Site Controls. In addition to control pools, a subset of 48 samples were run once at each of the three 618 processing centres (144 individual run files created, 1 failed QC). The first 36/48 consisted of randomly 619 selected high-grade serous ovarian carcinoma specimens, 12 from each processing centre. In addition, the 620 Vancouver site selected 12 samples from non-High-grade serous histology samples (3 clear cell, 3 621 endometrioid, 3 low-grade serous, 3 mucinous). Aliquots of RNA chosen at each site were sent on dry ice to 622 the other two processing centres. RNA from the 48 of the tumor samples were run on all three instruments to 623 assess concordance and the average r-squared was 0.981 (range 0.758-0.996). RNA from 1-2% of the samples 624 were randomly selected as technical replicates and run a second time to access concordance and to identify 625 any systematic problems with sample labelling. All 98 pairs of samples were concordant and the average r-626 squared was 0.978 (range 0.753-0.998).

627 Quality control and normalization of gene-expression data

628 Raw data were assessed using several quality assurance (QA) metrics to measure imaging quality,

629 oversaturation and overall signal to noise.

Imaging quality controls: Samples were flagged as imaging failures if the percentage of lane images FOV
 obtained was less than 75% of the requested number of fields.

Linearity of the assay: Samples were flagged as linearity failures if spiked-in positive control probes at
 different concentrations had R²<0.95.

Barbon Strailest Positive Control: Samples were flagged when the 0.5 fM positive control probe
 smaller than 2 standard deviations from the mean of the negative controls probes.

636 4. Sample Quality. Thresholds were set to maximize the number of samples of high quality included in the
 637 analysis. Sample Quality fails if either the Limit of Detection or Signal to Noise thresholds are not met.

a. % of Genes above Limit of Detection (LOD) of negative controls: LOD is an upper bound of the
background noise in the system, computed as two standard deviations above the mean of the spikedin negative control probes. Samples below a 50% threshold were deemed of poor quality and
considered failures.

b. Signal to noise ratio (S/N): calculated as a ratio between the geometric mean of housekeeping
genes and lower limit of detection: geometric mean/LOD. Samples with signal to noise ratio below a
170 threshold were deemed of poor quality and considered failures.

645 5. Overall QC. This is an overall quality control flag which fails if any of the Imaging, Linearity, or Smallest
646 Positive Control conditions fail.

647 Batch correction using control pools

648	The reference sample methods described in Talhouk et al was used. Briefly, assuming two batches A and B. To
649	calibrate samples with gene expressing X^{β} , that were run in batch B to samples with gene expression X^{A} , that
650	were run in batch A,
CE1	

Some number of reference samples (R) would be run in both batches A and B, resulting in expression
R^A and R^B.

653 • To remove Batch Effect: X^{B} - R^{B} and X^{A} - R^{A}

654 · Or alternatively: $X^{B} + (R^{A} - R^{B})$ would result in calibrating batch B to batch A.

As the same CodeSet was observed at all three sites, little difference was observed across sites; for

656 consistency, everything was calibrated to the Vancouver batch. The count data were log₂ transformed and

657 centered by the arithmetic mean of the selected housekeeping genes, *RPL19*, *ACTB*, *PGK1*, *SDHA*, and *POLR1B*.

658

659

660 **Prognostic signature**

To compare and evaluate competing methods for development of the signature, we used an AUC approach implemented in the "survivalROC" R software package, designed for a time-dependent setting with censoring where a lag exists between measurement of the biomarker and the disease outcome[124]. A ten-fold crossvalidation approach was used, thus each modelling approach was applied to a randomly selected nine-tenths of the training data, then the AUC was computed on the remaining one-tenth (out-of-bag sample; OOB) based on the continuous biomarker generated by the model. These AUC values guided selection of the best method, which was then applied to the full training dataset to identify the signature. Elastic net was the best 668 performing approach, hence the approach described below in B was applied to the full training dataset to 669 develop the final model. This signature was then evaluated in the testing data that were set aside.

A. *Stepwise*. Each gene was initially modelled separately, then those significant at a Bonferroni corrected 0.05
level were jointly modelled using a backwards stepwise approach, sequentially removing genes that did not
achieve a nominal 0.05 significance level. This process yielded a final model with coefficients that defined a
linear combination of adjustment covariates and expression features, forming a prognostic biomarker for OS.

674 B. *Elastic net*. In this regression method, a mixture of l_1 (lasso) and l_2 (ridge regression) penalties are applied. A 675 version has been developed for the Cox proportional hazards model, available in the glmnet R software package[125]. For each step of the overall cross-validation, the penalty parameter, λ , was selected as that 676 677 with the minimum mean cross-validated error as measured by the Cox partial likelihood for each randomly 678 selected nine-tenths of the training data, using the R function cv.glmnet from the glmnet R package. Thus, 679 nested cross-validation was conducted. The covariates age and stage were included in the model as mandatory 680 categorical variables by applying no penalization. Genes with non-zero coefficients in the penalized Cox model 681 were selected for the final model.

C. Boosting. Component-wise gradient boosting, a machine learning method, employs gradient descent 682 683 techniques to optimize a combined variable selection and model building strategy. It has been implemented 684 for general linear models and Cox proportional hazards model in the R package mboost[126]. To constrain the 685 model to account for age and stage we included an offset computed from the linear predictors from a Cox proportional hazards model. Boosting is an iterative strategy where at each step, parameters are updated by 686 687 weak estimators according to a pre-specified loss function. An important tuning parameter that can influence 688 over-fitting is the number of iterations, m_{stop} . We used 10-fold cross-validated estimates of empirical risk to 689 choose m_{stop} . Thus, nested cross-validation was performed for parameter tuning, similar to the elastic net 690 approach.

691 D. Random Survival Forests. This machine learning approach designed for survival outcomes[127] is an 692 ensemble method based on multiple decision trees, which are weakly predictive individually but combine to 693 yield what can be a strong predictor. The analysis was conducted using the rfsrc function in the R package, 694 randomForestSRC. To enhance performance in settings with large numbers of predictors, a preliminary feature 695 selection step is sometime performed. We conducted the analysis restricting to genes significant at a 0.05 696 Bonferroni corrected alpha level, similar to the stepwise approach described above. The tuning parameters for 697 these analyses included the number of trees, set to 200, and the maximum number of splits for each tree, set 698 to 20.

699 **Prognostic signature performance in relation to gene set size**

To further assess the need for all 101 genes included in the signature, we computed AUC sequentially adding genes according to the magnitudes of the corresponding coefficients. The gene with the most extreme coefficient was added first, followed by the gene with the next most extreme coefficient, and so on. Age and stage were included in the signature, and the AUC was computed using individuals from the test set only.

704 Genes correlated with genes in the prognostic signature

To find additional genes co-expressed with genes within our NanoString panel a correlation analysis was
 performed within The Cancer Genome Atlas (TCGA) Ovarian Serous Cystadenocarcinoma RNA-seq data set. A
 correlation coefficient was estimated for each NanoString panel gene vs all other genes included in the TCGA
 dataset. Genes with FDR < 0.05 were considered truly coexpressed and were considered strongly co-expressed

- if both Pearson and Spearman correlation coefficients were greater than 0.75 in absolute value.
- 710 Biological pathways that are enriched in the prognostic gene signature were investigated by computing
- overlaps between the 101 genes in the prognostic signature and the Reactome database1[128] using the
- 712 Broad Institute's Molecular Signature Database (MSigDB)[129]. Results were obtained for the top 20 pathways
- enriched within the gene set that had a FDR < 0.05. A similar analysis was performed to examine the biological

pathways that are enriched in genes that strongly co-expressed with the prognostic signature genes in the

715 TCGA data set.

716 Assigning signature score and hazard ratio to new patients

For future use of this signature with individual patients, it is necessary to compute the signature score based
on the coefficients for age, stage, and gene expression that are specified in Supplemental Table S10,
GeneSignatureCoefficients. Once a new FFPE sample is obtained, the following steps are taken to generate the

720 score:

An FFPE tumor sample is processed and profiled using the NanoString platform and codeset developed
 for the current study (GEO GPL26748) according to methods described above. The quality controls
 steps and normalization must be adhered to as described, including use of the housekeeping genes
 included in the codeset.

725 2. Patients are assigned to an age quartile group, and the corresponding variables, age.fq2, age.fq3,

age.fq4, are assigned a 1 if the patient fits this category or 0 otherwise. Cut-points were based on our

727 training data, ages 53, 60 and 67. Specifically 53 <= age.fq2 < 60, 60 <= age.fq3 < 67, age.fq4 >= 67.

Patients under age 53 are in the reference group, thus the three age.fq variables are all assigned 0 for
these patients.

Patients are assigned to a stage group, and the corresponding variables, stage.f1 and stage.f8 are
assigned a 1 if the patients fits this category or 0 otherwise. stage.f1 is defined as FIGO I-II and stage.f8
is assigned when FIGO is unknown. Patients with FIGO III-IV are in the reference category, thus stage.f1
and stage.f8 are both assigned 0 for these patients.

4. Each coefficient in Table S10 is multiplied by its corresponding age, stage or gene expression value.

5. The products are then summed to compute the scores. That is, letting *x* denote the vector of values

and β denote the vector of coefficients, the score, *s*, is equal to the matrix product, $s = x\beta^T$.

737 The score can be used to estimate the HR relative to the median score observed here	, which was -
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- 738 0.1664587 (mean = -0.2965882). The estimated HR for a given patient relative to the median would be
- 739 computed as exp((observed score median) * 1.215), where 1.215 was the COX regression coefficient for
- the signature score in the testing group in the present study. Scores for future patients can be compared
- to the patient characteristics reported here by the quintile groups shown in Supplemental Table S11 by
- 742 noting that the cut-points for the scores defining quintiles in the full dataset were -0.7320, -0.3126, -
- 743 0.0255, and 0.2658.



Figure S1. Permutation-based FDR for microarray meta-analysis results. Plot A displays FDR estimates and 95
 percent confidence regions for a series of increasingly stringent significance thresholds defined by the median
 p-value across the six studies. Plots B-D are permutation-based FDR results for the joint analyses across the six
 studies. The Omnibus tests, C, allow the gene effect to vary across molecular subtype, whereas GxS tests, D,
 assess evidence for differences in effect across molecular subtypes.



753

754 Figure S2. Boxplots of AUC values generated from 10-fold cross validation for competing gene expression 755 prognostic biomarker modelling approaches for two-year and five-year overall survival. Each approach was 756 applied to 10 datasets, each of which was a randomly sampled nine tenths of the training data. AUC values 757 were computed using these models applied to the remaining one tenths of the training data. Methods include boosting (Bst), elastic net (EN), random forests with BH feature selection (RF.BH), random forests with 758 759 Bonferroni feature selection (RF.Bn), stepwise with BH selection (stp.BH), and stepwise with Bonferroni 760 selection (stp.Bn). The superior performance of the elastic net method as compared to the backward stepwise 761 approach may be explained by the ability of elastic net to account for dependencies among the genes. The fact 762 that it performed better than random survival forests may indicate that Cox proportional hazards models are 763 relatively good at modeling the relation between gene expression and OS.



Figure S3. Graph of AUC from sequentially adding genes according to the magnitudes of the corresponding
coefficients. Age and stage were included in the signature, and the AUC was computed using individuals from
the test set only. The figure shows increasing AUC as the number of genes increases to the final size of 101,
with no clear indication of an asymptote. While it may be possible to reduce the number of genes in the
signature without impacting signature performance, the above results do not suggest specific genes that could
be removed.



Figure S4. KM curves of overall survival for patients in the training (A) and testing (B) sets. Patients were
assigned to quintiles by risk according to age and stage as estimated in Cox models. The smallest quintile has
the lightest shade and increasingly darker shades correspond to larger quintiles. Quintiles were calculated
independently for the training (A) and testing (B) sets. Shaded regions indicate 95 percent confidence regions.







Figure S6. KM curves of overall survival for patients with A) BRCA1, B) BRCA2, and C) no BRCA1 or BRCA2
 germline mutations. Patients were assigned to quintile groups, Q1 to Q5, based on the signature score, with
 Q1 having the lightest shade and increasingly darker shades corresponding to quintiles with greater scores.
 Quintiles were calculated independently for each of the three BRCA groups.



Figure S7. KM curves of overall survival for patients with A) CD8 score equal to 0, B) CD8 score equal to 1 or 2,
and C) 2 < CD8 score < 20. Patients were assigned to quintile groups, Q1 to Q5, based on the signature score,
with Q1 having the lightest shade and increasingly darker shades corresponding to quintiles with greater
scores. Quintiles were calculated independently for each of the three CD8 groups.



Figure S8. Violin plots of 101-gene signature scores for all 3769 HGSOC patients according to their molecular subtype. The data used to assign molecular subtypes overlapped with the data used to compute the signature score, however, hazard ratios for overall survival across quintiles of the signature were only minimally impacted by molecular subtype adjustment (Table 2), implying that the predictive value of the signature score

817 is largely independent of subtype.

818





Figure S9. KM curves of overall survival for patients with A) FIGO stage 1A, B) FIGO stage 1B, and C) FIGO stage
1A and 1B, D) FIGO stage 3C, E) FIGO stage 3C patients with no residual disease, F) FIGO stage 3C patients with
residual disease, and G) residual disease present, any FIGO stage. Patients were assigned to quintile groups, Q1
to Q5, based on the signature score, with Q1 having the lightest shade and increasingly darker shades
corresponding to quintiles with greater scores. Quintiles were calculated independently for each group.



Figure S10. KM curves of overall survival for patients with tissue from A) ovary, B) omentum, C) other
extraovarian sites, and D) omentum combined with other extraovarian sites. Patients were assigned to quintile
groups, Q1 to Q5, based on the signature score, with Q1 having the lightest shade and increasingly darker
shades corresponding to quintiles with greater scores. Quintiles were calculated independently for each of the
four tissue groups.



846 **Figure S11**. Signature scores in paired omentum vs. ovary tumor tissue samples collected from 42 HGSOC

- patients. The dashed line represents the line of identity, and the shaded area shows the 95% confidence region
- around the solid line, which is the least-squares fit to the points. These paired samples were processed and the
- 849 101-gene score computed using the Nanostring platform and computations described in Methods. The
- apparent linear relation observed in the plot demonstrates a strong correspondence between tissue types as
- suggested by the high Pearson correlation coefficient, r = 0.79 (p = 5.4 x 10⁻¹⁰).
- 852

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