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TECHNICAL NOTE

Fast and accurate relatedness estimation from high throughput sequencing data in the presence of inbreeding

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Abstract

Background: The estimation of relatedness between pairs of possibly inbred individuals from high-throughput sequencing (HTS) data has previously not been possible for samples where we can not obtain reliable genotype calls, as in the case of low coverage data.

Results: We introduce ngsRelateV2, a major revision of ngsRelateV1, a program which originally allowed for estimation of relatedness from HTS data among non-inbred individuals only. The new revised version takes into account the possibility of individuals being inbred by estimating the nine condensed Jacquard coefficients along with various other relatedness statistics. The program is threaded and scales linearly with the number of cores allocated to the process. **Conclusion:** The program is available as an open source C/C++ program under the GPL license and hosted at https://github.com/ANGSD/ngsRelate. To facilitate easy analysis, the program is able to work directly on the most commonly used container formats for raw sequence (BAM/CRAM) and summary data (VCF/BCF).

Key words: Relatedness estimation; inbreeding; Jacquard coefficients; high throughput sequencing data; genotype likelihood; NGS;threading.

Introduction

Being able to estimate how related two individuals are and whether they are inbred is important in several different fields ranging from conservation genetics to medical genetics. For this purpose, numerous coefficients, like the kinship coefficient and inbreeding coefficients, have been defined and many programs for estimating these coefficients have been proposed.

Notably, the genetic relationship between two individuals can be quantified by the extent to which the two individuals share their alleles identical-by-descent (IBD); i.e. are identical due to recent common ancestry. More specifically, for two

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cient and inbreeding coefficients, can be expressed as linear combinations of the nine condensed Jacquard coefficients.

In the specific case where neither individual is inbred, only three of the condensed Jacquard coefficients can be positive, namely J_7, J_8 and J_9 , which are often also denoted k_2, k_1 and k_0 , respectively, and known as Cotterman's coefficients [2]. Numerous approaches, based on either method of moments (e.g. [3]) or maximum-likelihood estimation (e.g. [4]), have been devised to estimate these three quantities assuming that the rest are zero and thus that the individuals are not inbred. This includes commonly used methods like PLINK and KING [3, 5]. Importantly, these methods can lead to wrong estimates and conclusions if applied to inbred individuals because the assumption that only J_7 , J_8 , and J_9 can be positive is violated. Hence in the presence of inbreeding one needs to estimate all nine coefficients. Several methods for doing this have been proposed [6, 7, 8]. However, very few current tools allow the user to do this and the few that do all require high quality genotype data as input (e.g. [8, 9]). They can therefore not be applied to HTS data of low depth, which is sometimes the only data available. Until recently the same was the case for all the methods for estimating relatedness between non-inbred individual. E.g. both PLINK and KING only work for genotype data. However, recently a few methods that can be applied to low depth sequencing data have been developed [10, 11]. One of these is ngsRelate [11] (hereafter referred to as ngsRelateV1), which works by integrating over every possible genotypic configurations and assigning these a probability given by their genotype likelihood. We here extend this software (hereafter referred to as ngsRelateV2) so it allows the user to infer all nine Jacquard coefficients, and thus allow for inference of relatedness in the presence of inbreeding as well as inference of the inbreeding coefficients for both individuals.

Materials & Methods

The underlying statistical framework is similar to that from ngsRelateV1 [11]. Given two individuals, *i* and *j*, from the same homogeneous population, we let D_l^i and D_l^j denote the observed HTS data at a biallelic locus *l*, and G_l^i and G_l^j denote the true, unobserved genotypes at the same locus. Furthermore, we let f_l denote the allele frequency at locus *l* in the relevant population and X_l denote the unobserved IBD state of the two individuals at locus *l*. Using this notation we can write the likelihood of the condensed Jacquard coefficients, $J = (J_1, J_2, J_3, J_4, J_5, J_6, J_7, J_8, J_9)$, for *L* independent (i.e.unlinked) biallelic loci as:

$$L(J|D^{i}, D^{j}, f^{A}) = \prod_{l=1}^{L} \sum_{m \in J} P(D_{l}^{i}, D_{l}^{j} \mid X_{l} = m, f_{l}^{A}) P(X_{l} = m|J)$$

Notably, here $P(X_l = m | J) = J_m$ and $P(D_l^i, D_l^j | X_l = m, f_l^A)$ can be rewritten as follows:

$$P(D_1^i, D_1^j | X_1 = m, f_1^A)$$

$$= \sum_{G_l^i, G_l^j \in \{0, 1, 2\}^2} P(D_l^i \mid G_l^i) P(D_l^j \mid G_l^j) P(G_l^i, G_l^j \mid f_l^A, X_l = m).$$

where $P(D_l^i | G_l^i)$ and $P(D_l^j | G_l^j)$ denote the per individual genotypes likelihoods for a biallelic locus *l*, which can be calculated from the sequencing data and $P(G_l^i, G_l^j | f_l^A)$ is given from Table 1. We use this likelihood function as a basis for performing maximum likelihood estimation. A number of useful estimates can be calculated directly from *J*, such as relatedness $(R = J_1+J_7+\frac{3}{4}(J_3+J_5)+\frac{1}{2}J_8)$, defined as the proportion of homologous alleles IDB [12], and per individual inbreeding coefficients, F_1 ans F_2 (as in [13]).

We here model the uncertainty of the sequencing data through the genotype likelihoods, but assume knowledge of population allele frequencies. In the presence of called genotypes (genotypes without uncertainty), our model coincides completely with the approach in [8]. In the absence of inbreeding our model reduces to the work in [11]. We assume that sites are independent, if they are linked our likelihood becomes a composite likelihood that will still have consistent estimates even though it has been shown that it can cause relationships to be overestimated [14, 15].

This novel method assumes that populations allele frequencies are obtainable, and we note that it has been shown in [16] that working in a context of solely diallelic markers the estimation of the the nine condensed Jaquard coefficients can display an issue of non-identifiability. This will have an impact for some of summary statistics that are defined as linear combinations of these coefficients, with the estimators that are invariant being R, F_a , F_b , θ , 2 - 3 - IBD, F_{diff} . Finally ngsRelateV2 also computes three summary statistics (Table 2, IBS) based on the 2D–SFS [17], but note that summary statistics based on the 2D–sfs do not require known population allele frequencies, they assume the individuals to be non–inbred. The 2D–SFS ob– tained in ngsRelateV2 follows the methodology from [18] that is based on genotype likelihoods and does therefore not require called genotypes.

In addition to the raw statistics we have also developed a bootstrapping approach that can be used to recover confidence intervals of all the summary statistics shown in Table 2.

Simulations

To simulate data with L sites and N diploid individuals, we first sampled L allele frequencies from a uniform distribution with a minor allele frequency (MAF) filter on 0.05 and 0.1. For each site for each of the N individuals, we sample two alleles using indepedent Bernoulli trials with the probability of success equal to the allele frequency for the given site, implying the data is generated under the assumption of Hardy–Weinberg equilibrium. The outcome of these two trials represent the genotype. Gametes of these individuals are subsequently generated by sampling either of the two alleles from the two haplotypes for every site with equal probability. We assume that each site is independent, thus, linkage disequilibrium (LD) is not modeled. Allosomes are disregarded as well.

From the N founder individuals, we simulate offspring to generate three different pedigrees. From these pedigrees, we have analyzed pairs of individuals with the expected Jacquard coefficients as shown in Table 3.

We then proceed by calculating genotype likelihoods by assuming different sequencing depths $d = \{1X, 2X, 4X, 8X, 16X\}$, error rate e = 0.001 and number of sites $s = \{10K, 30K, 50K\}$ for the individuals of interest. The per-site-per-individual sequencing depth is given by sampling the depth from a Poisson distribution with parameter d and using the binomial density distribution with e. This approach is similar to the previous approach in [11] which does not model the spatial properties of true recombination and LD.

 Table 1. Probabilities for various allelic states, given modes of IDB from Table 1 in [8]. Triallelic sites are disregarded.

Allelic State	J ₁	J ₂	J ₃	J ₄	J ₅	J ₆	J ₇	J ₈	J ₉
$egin{aligned} A_iA_iA_iA_i\ A_iA_iA_jA_j\ A_iA_iA_iA_j\ A_iA_iA_iA_j\ A_iA_iA_iA_j \end{aligned}$	р _і 0 0 0	p _i ² p _i p _j 0 0	p _i ² 0 p _i p _j 0	p_i^3 $p_i p_j$ $2 p_i^2 p_j$ 0	p _i ² 0 0 P _i P _j	p_i^3 $p_i^2 p_j$ 0 $2p_i^2 p_j$	p _i ² 0 0 0	p_i^3 o $p_i^2 p_j$ $p_i^2 p_j$	$p_i^4 \ p_i^2 p_j^2 \ 2 p_i^3 p_j \ 2 p_i^3 p_j$
$A_i A_j A_i A_j$	0	0	0	0	0	0	$2p_ip_j$	$p_i p_j$	$4p_i^2 p_j^2$

 Table 2.
 Various relatedness statistics estimated by ngsRelateV2 and which summary statistics they are based on.

Statistics	Formula	summary statistic	Reference
r _{ab}	$(J_1 + J_7 + 0.75 * (J_3 + J_5) + .5 * J_8)$	IBD	[12]
Fa	$(J_1 + J_2 + J_3 + J_4)$	IBD	[19]
F_b	$(J_1 + J_2 + J_5 + J_6)$	IBD	[19]
θ	$J_1 + 0.5 * (J_3 + J_5 + J_7) + 0.25 * J_8$	IBD	[19]
F ₁₂	$J_1 + 0.5 * J_3$	IBD	[12]
F ₂₁	$J_1 + 0.5 * J_5$	IBD	[12]
Fraternity	$J_2 + J_7$	IBD	[20]
Identity	J_1	IBD	[20]
Zygosity	$J_1 + J_2 + J_7$	IBD	[20]
2-3-IBD	$J_1 + J_2 + J_3 + J_5 + J_7 + 0.5 * (J_4 + J_6 + J_8)$	IBD	[16]
F _{diff}	$0.5 * (J_4 - J_6)$	IBD	[16]
Ro	(C + G)/E	IBS	[17]
R_1	E/(B + D + H + F + C + G)	IBS	[17]
King	(E - 2(C + G))/(B + D + H + F + 2 * E)	IBS	[17]

Table 3. Expected Jacquard coefficients, relatedness and inbreeding coefficients for three simulated scenarios.

	J_1	J ₂	J ₃	J ₄	J ₅	J ₆	J ₇	J ₈	J ₉	R	F_1	F ₂
scenario1	0	0	0	0	0	0	0	0.25	0.75	0.13	0	0
scenario2	0	0	0	0	0.06	0.19	0	0.38	0.38	0.23	0	0.25
scenario3	0.02	0.02	0.09	0.12	0.06	0.06	0.06	0.38	0.22	0.38	0.25	0.13

Results

To test the performance of ngsRelateV2, we use three simulated scenarios (see Simulations section) and compare it to ngsRelateV1 [11]. For every scenario, we generate 100 independent simulations for every combination of sequencing effort and number of segregating sites. In the first scenario, we compare two outbred cousins (Figure 1). As expected, both versions of ngsRelate find not only the correct level of relatedness, but also the correct estimates of the three relevant Jacquard coefficients (J_7, J_8, J_9) . The second scenario also includes two cousins, but this time we have introduced inbreeding in one of the individuals. The parents of the inbred individual are related corresponding to a parent-child relation. In this scenario, even at low sequencing effort and only 10k sites, ngsRelateV2 correctly estimates the coefficients of relatedness and inbreeding; however, the estimates of the nine Jacquard coefficients are somewhat noisy, and at least 50k segregating sites are needed to increase the accuracy (Figure 2). The final scenario, being the most complex, includes the inbred individual from scenario two and another inbred cousin with its parents being related corresponding to a grandparent-grandchild relation. Interestingly, with such a complex pedigree, ngsRelateV2 still manages to recover the exact estimates for relatedness and individual inbreeding coefficients, even with only 10k segregating sites and a low sequencing depth (Figure 3). Similarly to the results from scenario two, accurate estimates of the nine Jacquard coefficients required increasing the number of informative sites and/or the sequencing effort. We also applied ngsRelateV2 to these three scenarios using a MAF cutoff on 0.05 (Supplementary Figure 1-3). We find that ngsRelateV2 recovers comparable accuracy with a MAF filter on 0.05.

We also applied ngsRelateV2 to real HTS data and compared

the estimates to those obtained with ngsRelateV1. We used six pairwise related genomes, sequenced to low coverage (approximately 4X), from the LWK population generated as part of the 1000 Genomes Project [21]. We calculated genotype likelihoods of the related individuals, using ANGSD [18], at genomic sites with MAF in the LWK population on 0.05, summing up to 4.6m segregating sites. We not only show that ngsRelateV2 obtains comparable relatedness estimates to those obtained by ngsRelateV1, with this novel software, we also show that all the tested individuals show an inbreeding coefficient below 1% (Figure 4).

In extremely complicated pedigrees with symmetric inbreeding, such as multiple generations of full sibling mating, we find multiple global maxima where several combinations of the nine Jacquard coefficients, including the expected coefficients, are equally likely. Albeit observing such identifiability challenges, we, importantly, still find accurate relatedness estimates and individual inbreeding coefficients by summing the relevant Jacquard coefficients.

For every pair of individuals, ngsRelateV2 generates and outputs estimates of the nine Jacquard coefficients, the relatedness, the individual inbreeding coefficients as described above but also other combinations of the nine Jacquard coefficient: the kinship coefficient, fraternity, and the three summary statistics inbred relatedness, identity, and zygosity, suggested by Ackerman and colleagues [20]. It also produces the KING statistic [22] based on the two dimensional site frequency spectrum of pairs of individuals following the methodology in [17]. The latter statistics do not require population allele frequencies.

Cores	Memory Usage	Wall-clock	CPU-time
1	45GB	85.59	87.14
2	46.9GB	41.38	81.39
4	49.3GB	23.36	89.03
8	54.2GB	11.88	88.69
16	63.5GB	6.30	89.73
32	83.2GB	3.39	87.81

Table 4. Run statistics for 34 GB BCF file, as a function of number of cores. Shown is the memory requirement, wall-clock (actual runtime) and CPU time (both in hours), see also Figure 5.

Computational speed and memory requirements

To take advantage of the increasing number of cores of available on modern computers we employ a multilevel threading approach by both parallelizing both the file reading and the actual analysis. In Figure 5 we analyzed a semi-random dataset consisting of 135 samples mainly from the [23] publication. The input for the program was a 34 GB BCF file generated with standard bcftools with a liberal 164 mio number of SNP sites. We timed the actual runtime (wall-clock) and the CPU-time for varying number of cores (1,2,4,8,16,32) and noted the memory usage for each run, since allocating more cores for the process requires additional internal datastructures and does therefore also increases the memory requirements as seen by Table 4. From both the table and figure we observe a near linear correlation between the number of cores and the runtime with the CPU time remaining almost constant.

Conclusion

The tool presented in this technical note allows researchers to perform relatedness analysis for inbred individuals in a statistical framework that is especially suited for low coverage sequence data. The results show that the method performs well for estimating all nine coefficients, at least when the underlying pedigrees are not extremely complex. And even when the underlying pedigree is very complex, compound summaries of the output, like relatedness and inbreeding coefficients, will still be correct. The implementation is a fast multi threaded C++ program that can be directly applied to the most commonly used data files used for high throughput sequencing data.

Implementation Details

The program is implemented in a fast multithreaded c++ program and takes as input either genotype likelihood files and frequencies, bcf/vcf files as produced from standard tools such as GATK[24] or SAMtools [25], and binary-format plink files [3]. We also include an R implementation that we used for simulating data. Of note, the simulations generated in this study do not accounted for LD. In case of LD between genetic variants, the likelihood function becomes a composite likelihood function. The maximum likelihood estimate of such a function is consistent to that found with a likelihood function of independent sites [26].

The optimization follows the approach described in [11]. The optimization is an accelerated expectation maximization (EM) following the squared iterative approach in **S3** in [27] and is initialized with a random start point within the parameter space. The borders of the parameter space are manually examined after convergence. Since the EM algorithm is only guaranteed to find a local optimum, it is recommended to rerun with multiple different seeds though we note that we did not find an issue

with multiple local optima in our examples.

Availability of source code and requirements

- Project name: ngsRelateV2
- Project home page: http://github.com/ANGSD/ngsRelate
- Operating system(s): Platform independent
- Programming language: C++
- Other requirements: htslib (only for parsing VCF/BCF files)
- License: GNU GPL (version 3)
- RRID: SCR_016588
- GigaDB: Snapshots of the code and other supporting data are available in the GigaScience repository, [28]

Declarations

List of abbreviations

IBD: Identity-by-descent. HTS: High-throughput sequencing.

Consent for publication

Not applicable.

Competing Interests

None.

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Authors' Contributions

TSK devised the model. It was first prototyped by PAA as part of a master project under the supervision of TSK and IM. KH implemented and ran all analyses. IM and AM devised test scenarios and improved early versions of the method. All authors wrote the article.

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Figure 1. 100 independent simulations of two outbred cousins across variable sequencing depth and informative sites with a minor allele frequency cutoff on 10%. J_9 to J_1 refer to the nine Jacquard coefficients, R is the relatedness, finally, F_1 and F_2 refer to the individual inbreeding coefficients. Simulation (green) are the true values that we compare ngsRelateV1 (red) and the new program ngsRelateV2 (blue) against.

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Figure 2. 100 independent simulations of two cousins, with one individual being inbred, across variable sequencing depth and segregating sites with a minor allele frequency cutoff on 10%. J_9 to J_1 refer to the nine Jacquard coefficients, R is the relatedness, finally, F_1 and F_2 refer to the individual inbreeding coefficients. Simulation (green) are the true values that we compare ngsRelateV1 (red) and the new program ngsRelateV2 (blue) against.

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Figure 3. 100 independent simulations of two cousins, both being inbred, across variable sequencing depth and segregating sites with a minor allele frequency cutoff on 10%. J_9 to J_1 refer to the nine Jacquard coefficients, R is the relatedness, finally, F_1 and F_2 refer to the individual inbreeding coefficients. Simulation (green) are the true values that we compare ngsRelateV1 (red) and the new program ngsRelateV2 (blue) against.



Figure 4. Estimated Jacquard coefficients from six pairs of related individuals. The estimates are based on low-depth NGS data from the 1000 Genomes Project using ngsRelateV1 and ngsRelateV2. J_9 to J_1 refer to the nine Jacquard coefficients, R is the relatedness, finally, F_1 and F_2 refer to the individual inbreeding coefficients. CO: Cousins, HS: Half Siblings, PC: Parent-Child, AV: Avuncular, FS: Full Siblings.



Figure 5. Running times for ngsRelateV2 on a dataset with 135 individuals 164mio possible SNP sites. Blue line is the overall CPU usage across all threads allocated to the main process. Redline being the runtime for the process to finish. The actual values along with memory usage can be found in Table 4

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