- 1 Title:
- 2 Genetic risk for Parkinson's disease and progression: an analysis of 13 longitudinal cohorts
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### 4 Authors:

- 5 Hirotaka Iwaki M.D.<sup>1,2</sup>, Cornelis Blauwendraat Ph.D.<sup>1</sup>, Hampton L. Leonard M.S.<sup>1</sup>, Ganqiang Liu
- 6 Ph.D.<sup>3,4,5</sup>, Jodi Maple-Grødem Ph.D.<sup>6,7</sup>, Jean-Christophe Corvol M.D. Ph.D.<sup>8</sup>, Lasse Pihlstrøm M.D.
- 7 Ph.D.<sup>9</sup>, Marlies van Nimwegen Ph.D.<sup>10</sup>, Samantha J. Hutten Ph.D.<sup>11</sup>, Khanh-Dung H. Nguyen Ph.D.<sup>12</sup>,
- **8** Jacqueline Rick Ph.D.<sup>13</sup>, Shirley Eberly M.S.<sup>14</sup>, Faraz Faghri M.S.<sup>1,15</sup>, Peggy Auinger M.S.<sup>16</sup>, Kirsten M.
- **9** Scott MRCP, M.Phil.<sup>17</sup>, Ruwani Wijeyekoon MRCP<sup>17</sup>, Vivianna M. Van Deerlin M.D. Ph.D.<sup>18</sup>, Dena G.
- 10 Hernandez Ph.D.<sup>1</sup>, Aaron G. Day-Williams Ph.D.<sup>19,20</sup>, Alexis Brice M.D.<sup>21,22,23</sup>, Guido Alves M.D.,
- 11 Ph.D.<sup>6,24,25</sup>, Alastair J. Noyce MRCP, Ph.D.<sup>26,27</sup>, Ole-Bjørn Tysnes M.D., Ph.D.<sup>28,29</sup>, Jonathan R. Evans
- 12 MRCP, Ph.D.<sup>30</sup>, David P. Breen MRCP, Ph.D.<sup>31,32,33</sup>, Karol Estrada Ph.D.<sup>12</sup>, Claire E. Wegel MPH<sup>34</sup>,
- 13 Fabrice Danjou M.D., Ph.D.<sup>21</sup>, David K. Simon M.D., Ph.D.<sup>35,36</sup>, Bernard Ravina M.D.<sup>37,38</sup>, Mathias Toft
- 14 M.D., Ph.D.<sup>9,39</sup>, Peter Heutink Ph.D.<sup>40,41</sup>, Bastiaan R. Bloem M.D., Ph.D.<sup>10</sup>, Daniel Weintraub M.D.<sup>42,43</sup>,
- 15 Roger A. Barker MRCP, Ph.D.<sup>44</sup>, Caroline H. Williams-Gray MRCP, Ph.D.<sup>44</sup>, Bart P. van de Warrenburg
- 16 M.D., Ph.D.<sup>10</sup>, Jacobus J. Van Hilten M.D., Ph.D.<sup>45</sup>, Clemens R. Scherzer M.D.<sup>3,4,5</sup>, Andrew B. Singleton
- 17 Ph.D.<sup>1</sup>, Mike A. Nalls Ph.D.<sup>1,2</sup>
- 18

<sup>1</sup> Laboratory of Neurogenetics, National Institute on Aging, National Institutes of Health, Bethesda, MD, USA, 2 Data 19 Tecnica International, Glen Echo, MD, USA, 3 Precision Neurology Program, Harvard Medical School, Brigham and 20 Women's Hospital, Boston, MA, USA, 4 Neurogenomics Laboratory, Harvard Medical School, Brigham and Women's 21 Hospital, Boston, MA, USA, 5 Ann Romney Center for Neurologic Diseases, Brigham and Women's Hospital, Boston, 22 MA, USA, 6 The Norwegian Centre for Movement Disorders, Stavanger University Hospital, Stavanger, Norway, 7 The 23 Centre for Organelle Research, University in Stavanger, Stavanger, Norway, 8 Assistance-Publique Hôpitaux de Paris, 24 ICM, INSERM UMRS 1127, CNRS 7225, ICM, Department of Neurology and CIC Neurosciences, Pitié-Salpêtrière 25 Hospital, Paris, France, 9 Department of Neurology, Oslo University Hospital, Oslo, Norway, 10 Department of 26 Neurology, Donders Institute for Brain, Cognition, and Behaviour, Radboud University Medical Centre, Nijmegen, The 27 Netherlands, 11 Michael J Fox Foundation, New York, NY, USA, 12 Translational Genome Sciences, Biogen, Cambridge, 28 MA, USA, 13 Department of Neurology University of Pennsylvania, Philadelphia, PA, USA, 14 Department of 29

Biostatistics and Computational Biology, University of Rochester, Rochester, NY, USA, 15 Department of Computer 1 Science, University of Illinois Urbana-Champaign, Champaign, IL, USA, 16 Department of Neurology, Center for Health 2 + Technology, University of Rochester, Rochester, NY, USA, 17 Department of Clinical Neurosciences, University of 3 Cambridge, John van Geest Centre for Brain Repair, Cambridge, UK, 18 Department of Pathology and Laboratory 4 Medicine, Center for Neurodegenerative Disease Research, Parelman School of Medicine at the University of 5 Pennsylvania, Philadelphia, PA, USA, 19 Genetics and Pharmacogenomics, Merck Research Laboratory, Boston, MA, 6 USA, 20 Statistical Genetics, Biogen, Cambridge, MA, USA, 21 Institut du cerveau et de la moelle épinière ICM, Paris, 7 France, 22 Sorbonne Université SU, Paris, France, 23 INSERM UMR1127, Paris, France, 24 Department of Neurology, 8 Stavanger University Hospital, Stavanger, Norway, 25 Department of Mathematics and Natural Science, University of 9 Stavanger, Stavanger, Norway, 26 Preventive Neurology Unit, Wolfson Institute of Preventive Medicine, Queen Mary 10 University of London, London, UK, 27 Department of Molecular Neuroscience, UCL Institute of Neurology, London, 11 UK, 28 Department of Neurology, Haukeland University Hospital, Bergen, Norway, 29 University of Bergen, Bergen, 12 Norway, 30 Department of Neurology, Nottingham University NHS Trust, Nottingham, UK, 31 Centre for Clinical Brain 13 Sciences, University of Edinburgh, Edinburgh, Scotland, 32 Anne Rowling Regenerative Neurology Clinic, University of 14 Edinburgh, Edinburgh, Scotland, 33 Usher Institute of Population Health Sciences and Informatics, University of 15 Edinburgh, Edinburgh, Scotland, 34 Department of Medical and Molecular Genetics, Indiana University, Indianapolis, IN, 16 USA, 35 Department of Neurology, Beth Israel Deaconess Medical Center, Boston, MA, USA, 36 Harvard Medical 17 School, Boston, MA, USA, 37 Voyager Therapeutics, Cambridge, MA, USA, 38 Department of Neurology, University of 18 Rochester School of Medicine, Rochester, NY, USA, 39 Institute of Clinical Medicine, University of Oslo, Oslo, 19 Norway, 40 German Center for Neurodegenerative Diseases-Tubingen, Tuebingen, Germany, 41 HIH Tuebingen, 20 Tubingen, Tuebingen, Germany, 42 Department of Psychiatry, University of Pennsylvania School of Medicine, 21 Philadelphia, PA, USA, 43 Department of Veterans Affairs, Philadelphia, PA, USA, 44 Department of Clinical 22 Neurosciences, University of Cambridge, Cambridge, UK, 45 Department of Neurology, Leiden University Medical 23 Center, Leiden, The Netherlands. 24 25 26 27 **Corresponding author:** 28 Mike A. Nalls Ph.D. 29 Laboratory of Neurogenetics, National Institute on Aging, National Institutes of Health 30 35 Convent Drive, Bethesda, MD 20892, USA 31 +1-202-468-1533 32 nallsm@mail.nih.gov 33 34 Statistical analysis was conducted by: 35 Hirotaka Iwaki M.D. Ph.D.

- 1 Laboratory of Neurogenetics, National Institute on Aging, National Institutes of Health
- 2 35 Convent Drive, Bethesda, MD 20892, USA
- 3
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1 Abstract

2 Objective To determine if any association between previously identified alleles that confer risk for
3 Parkinson's disease and variables measuring disease progression.

4 Methods We evaluated the association between 31 risk variants and variables measuring disease

5 progression. A total of 23,423 visits by 4,307 patients of European ancestry from 13 longitudinal cohorts

6 in Europe, North America, and Australia were analyzed

7 **Results** We confirmed the importance of *GBA* on phenotypes. *GBA* variants were associated with the

8 development of daytime sleepiness (p.N370S: HR 3.28 [1.69, 6.34]) and possible REM sleep behavior

9 (p.T408M: OR 6.48 [2.04, 20.60]). We also replicated previously reported associations of *GBA* variants

10 with motor/cognitive declines. The other genotype-phenotype associations include an intergenic variant

11 near *LRRK2* and the faster development of motor symptom (Hoehn and Yahr scale 3.0 HR 1.33 [1.16,

12 1.52] for the C allele of rs76904798); an intronic variant in *PMVK* and the development of wearing-off

13 effects (HR 1.66 [1.19, 2.31] for the C allele of rs114138760). Age of onset was associated with

14 *TMEM175* variant p.M393T (-0.72 [-1.21, -0.23] in years), the C allele of rs199347 (intronic region of

15 *GPNMB*, 0.70 [0.27, 1.14]), and G allele of rs1106180 (intronic region of *CCDC62*, 0.62 [0.21, 1.03])

16 Conclusions This study provides evidence that alleles associated with Parkinson's disease risk, in

17 particular *GBA* variants, also contribute to the heterogeneity of multiple motor and non-motor aspects.

18 Accounting for genetic variability will be a useful factor in understanding disease course and in

19 minimizing heterogeneity in clinical trials.

# 1 Introduction

2 Parkinson's disease is one of the most common neurodegenerative diseases, with an estimated lifetime risk as high as 1-2%.<sup>1</sup> Parkinson's disease is traditionally characterized by motor features such as 3 4 bradykinesia, rigidity, and tremor. However, in addition to these motor symptoms, patients with 5 Parkinson's disease also develop non-motor symptoms, which include depression, cognitive decline, sleep 6 abnormalities, reduced olfaction, and autonomic dysfunction.<sup>2</sup> Collectively, the combined spectrum of 7 motor and non-motor symptoms more accurately reflects the multisystem nature of the disease. Patients 8 with Parkinson's disease may present with various combinations of symptoms and show differences in 9 the rates of progression.<sup>3</sup> The application of modern molecular genetic approaches over the last decade 10 has revealed a significant number of genetic risk loci for idiopathic Parkinson's disease.<sup>4-7</sup> However, in 11 comparison with case-control GWAS, analyzing how genetic factors influence clinical presentation and 12 progression requires longitudinal cohorts with much more detailed observations. Such data are sparse and 13 individual cohorts are often small in size and quite varied, posing a challenge both in sample size and 14 heterogeneity.

In an attempt to address these issues, we collected data from 13 distinct longitudinal Parkinson's disease cohorts with detailed clinical data, including assessment of disease progression. We sought to determine if Parkinson's disease genetic risk factors, either in the form of known GWAS variants or an aggregate genetic risk score, are associated with changes in clinical progression and the disease features.

19

#### 20 Materials and methods

### 21 Study design and participants

A total of 13 Parkinson's disease cohorts from North America, Europe, and Australia participated in the
 study. Nine were prospective observational cohorts and the rest were from randomized clinical trials. The
 observational cohorts were Drug Interaction with Genes in Parkinson's Disease (DIGPD), Harvard

1 Biomarkers Study (HBS). Oslo Parkinson's Disease study (partly including retrospective data). The Norwegian ParkWest study (ParkWest), Parkinson's Disease Biomarker Program (PDBP), Parkinsonism 2 3 Incidence and Cognitive and Non-motor heterogeneity In Cambridgeshire, Parkinson's Progression 4 Markers Initiative (PPMI), Profiling Parkinson's disease study (ProPark), and the Morris K. Udall 5 Centers for Parkinson's Research (Udall). The four cohorts from randomized clinical trials were Deprenyl 6 and Tocopherol Antioxidative Therapy of Parkinsonism (DATATOP), NIH Exploratory Trials in 7 Parkinson's Disease Large Simple Study 1 (NET-PD\_LS1), ParkFit study (ParkFit), and Parkinson 8 Research Examination of CEP-1347 Trial with a subsequent prospective study (PreCEPT/PostCEPT). 9 Information on these cohorts can be found in the Appendix e-1. Subsets of participants from the cohorts 10 who provided DNA and were non-related participants with PD, diagnosed at the age of 18 or later, and of 11 European ancestry, were included in the study. Participants' information and genetic samples were 12 obtained under appropriate written consent and with local institutional and ethical approvals.

13

## 14 Genotyping SNPs and calculation of genetic risk score

15 Oslo samples were genotyped on the Illumina Infinium OmniExpress array, DIGPD samples were 16 genotyped by Illumina Multi-Ethnic Genotyping Array, and all other samples were genotyped on the 17 NeuroX array.<sup>8</sup> The quality control process of variant calling included GenTrain score < 0.7, minor allele 18 frequency (MAF) > 0.05 (for sample QC but not in our analysis of rare risk factors), and Hardy-Weinberg 19 equilibrium test statistic >  $10^{-6}$ . Sample-specific quality control included a sample call rate of > 0.95, 20 confirmation of sex through genotyping, homozygosity quantified by F within  $\pm 3$  SD from the population 21 mean, European ancestry was confirmed by principal-components analysis with 1000 genomes data as the 22 reference, and genetic relatedness of any two individuals < 0.125. Detailed information regarding NeuroX and the quality control process has been described previously.<sup>9</sup> In the present study, we investigated 31 23 24 SNPs previously shown to be significantly associated with Parkinson's disease.<sup>10–12</sup> In addition, we also 25 calculated a genetic risk score (GRS) for each participant based on these variants. The scores were

transformed into Z-scores within each cohort and treated as an exposure, with effect estimates based on
one standard deviation change from the population mean. The list of 31 SNPs and the GRS calculation
method are provided in the Table e-1.

Furthermore, principal components (PCs) were created for each dataset from genotypes using PLINK. For the PC calculation, variants were filtered for minor allele frequency (> 0.05), genotype missingness (< 0.05) and Hardy–Weinberg equilibrium (*P*-value  $\ge 10^{-5}$ ). The remaining variants were pruned (using a 50 kb window, with a 5 SNP shift per window and  $r^2$  threshold of 0.5) and PCs were calculated using the pruned variants.

9

#### 10 Measurements

11 The following clinical measurements and binomial outcomes were recorded longitudinally. (Table e-2) 12 Total and sub-scores of the Unified Parkinson's Disease Rating Scale (UPDRS) or the Movement 13 Disorder Society revised UPDRS version (MDS-UPDRS; modified Hoehn and Yahr scales (HY); 14 modified Schwab and England Activities of Daily Living Scale (SEADL); and scores for the Mini-Mental 15 State Examination (MMSE), SCOPA-cognition, and Montreal Cognitive Assessment (MoCA). Each was 16 treated as a continuous outcome. For the UPDRS and MDS-UPDRS scores specifically, we took Z-scores 17 of the total and sub-scores (except for part 4 at baseline) to compare the original and revised UPDRS 18 versions. The conversion was applied to the scores for all subsequent visits. For UPDRS part 4, most 19 participants had very low scores or 0 at baseline, so we normalized across all observations within each 20 cohort. We also analyzed binomial outcomes. If we had access to the raw data, we used common cut-off 21 values which had been tested and reported specificity of 85% or more in patients' population. The 22 binomial outcomes include: existence of family history (1st degree relative. 1st and 2nd degree relatives 23 in HBS, PreCEPT, ProPark, and Udall), hyposmia (University of Pennsylvania Smell Identification Test; UPSIT < 21,<sup>13</sup> or answering "yes" to question 2 in the Non-motor Symptoms (NMS) questionnaire), 24 cognitive impairment (SCOPA-cognition < 23, MMSE < 27 or MoCA < 24,<sup>14,15</sup> or diagnosed with DSM-25 IV criteria for dementia), wearing off (UPDRS/MDS-UPDRS part 4 off time > 0 or physician's 26

1	diagnosis), dyskinesia (UPDRS/MDS-UPDRS part 4 dyskinesia time > 0 or physician's diagnosis),
2	depression (Beck Depression Inventory (BDI) > 14 (PICNICS used 9 instead of 14), Hamilton Depression
3	Rating Scale (HDRS) $> 9$ , Geriatric Depression Scale (GRS) $> 5$ , <sup>16</sup> or physician's diagnosis), constipation
4	(MDS-UPDRS part 1 item $11 > 0$ , or answering "yes" to question 5 in the NMS questionnaire), excessive
5	daytime sleepiness (Epworth Sleepiness scale (ESS > 9), <sup>17</sup> insomnia (MDS-UPDRS part 1, item $7 > 0$ ),
6	REM sleep behavior disorder (RBD) (answered "yes" to question 1 on the Mayo Sleep Questionnaire
7	(MSQ), <sup>18</sup> or REM sleep behavior disorder screening questionnaire (RBDSQ > 5), <sup>19</sup> restless legs syndrome
8	(RLS) (answered "yes" to MSQ question 3, <sup>20</sup> or RLS diagnosis positive by RBDSQ), and the progression
9	to HY $\geq$ 3 (HY3, representing moderate to severe disease) and death. The individual definitions of these
10	binomial outcomes are summarized in Table e-2. Age, sex, years of education, age at motor symptom
11	onset, and whether or not the patient was treated with levodopa or dopamine agonists at each visit were
12	also recorded for adjustments.

#### 14 Statistical analysis

#### 15 Cohort-level analysis

16 We analyzed the association between exposures and outcomes using appropriate additive models. 17 Covariates of interest were not available for all cohorts; therefore, the model specifications were slightly 18 different between cohorts (detailed in Table e-3). Briefly, the associations between a SNP/GRS and age at 19 onset were analyzed by linear regression modeling adjusting for population stratification (PC1 and PC2). 20 The association between family history of Parkinson's disease and SNP/GRS was analyzed with a logistic 21 regression model adjusting for PC1/2. For continuous variables, linear regression modeling adjusting for 22 sex, education, PC1/2, age at onset, years from diagnosis, family history, and treatment status were 23 applied. For those who had multiple observations, random intercept was added to adjust for repeated 24 measurements of the same individual. For binomial outcomes, the logistic regression at baseline

25 observation was applied using the same covariates as the continuous models. Those that were negative at

- baseline were further analyzed by a Cox regression with the same covariates but with treatment status as a
  time-varying covariate. Observations with missing variables were excluded from the analyses.
- 3

#### 4 Meta-analysis

5 We applied inverse weighting (precision method) for each combination of outcome-predictor association 6 and combined the estimates from the 13 different cohorts in a fixed-effect model. Multiple test correction 7 for SNPs was controlled with an overall false discovery rate (FDR) of 0.05 per outcome being considered 8 significant. Similarly, multiple testing of outcomes for GRS was corrected with an FDR of 0.05, but across all traits. In addition, as a test of homogeneity,  $I^2$  indices and forest plots were used for quantitative 9 10 assessment. As a sensitivity analysis, we conducted up to 13 iterations of the meta-analyses for the 12 11 cohorts excluding each cohort per iteration. This analysis provides information regarding heterogeneity of 12 the cohorts and how one specific cohort exclusion impacts the results. The range of estimates and 13 maximum P-values for the iterations were included. Finally, we conducted the 13-cohort meta-analysis in 14 a random effects model with REML estimation using the same multiple testing correction. 15 16 All of the above analyses were conducted with PLINK version 1.9, and R version 3.4.4 (64-bit). Statistical tests were all two-sided. 17 18 19 Data availability 20 Qualified investigators can request raw data through the organizations' homepages (PDBP: 21 https://pdbp.ninds.nih.gov/, PPMI: https://www.ppmi-info.org/) or collaboration. 22

## 23 **Results**

1 A total of 23,423 visits by 4,307 patients with a median follow-up period of 2.97 years (quartile range of 2 [1.63, 4.94] years) were eligible for the analysis. The baseline characteristics of the cohorts are shown in 3 Table 1. The mean ages at onset varied from 54 to 69 years old; the average disease durations at cohort 4 entry ranged from less than one to 10 years, and the mean observation periods were between 1.2 and 6.8 5 years. All DATATOP, ParkWest, PPMI, and PreCEPT participants were dopaminergic therapy-naive at 6 baseline; patients in the other cohorts were not. In the primary analysis of 13 cohorts, 17 associations 7 were identified as significant after FDR correction (Table 2, and more information in Table e-4). 8 Overwhelmingly, 10 were associated with GBA variants. In particular, GBA p.E365K (rs2230288) was 9 associated with 2.37- [1.53, 3.66] (95% CI) fold higher odds of having cognitive impairment at baseline 10  $(P = 1.09 \times 10^4)$  and 2.78- [1.88, 4.11] fold higher hazard ratio of developing cognitive impairment during 11 follow-up among those who were negative for cognitive impairment at baseline ( $P = 2.97 \times 10^{-7}$ ). This 12 SNP was also associated with a higher mean on the HY Scale at 0.10 [0.04, 0.16] ( $P = 1.53 \times 10^{-3}$ ), but the test of homogeneity was rejected (P = 0.017,  $I^2 = 48.9\%$ ). In addition, it was associated with the 13 14 development of a REM sleep behavior disorder among those who didn't have the disorder at baseline. 15 Other GBA mutations, p.N370S (rs767763715) and p.T408M (rs75548401), were both associated with a 16 higher HR of reaching HY3 (4.59 [2.60, 8.10] for p.N370S ( $P = 1.58 \times 10^{-7}$ ) and 1.93 [1.34, 2.78] for 17 p.T408M ( $P = 4.40 \times 10^{-4}$ )). GBA p.N370N was also associated with a higher risk of developing wearing-18 off, dyskinesia, and daytime sleepiness. p.T408M was associated with a 6.48 [2.04, 20.60] times higher 19 odds ratio of having a REM sleep behavior disorder symptom at baseline ( $P = 1.53 \times 10^{-3}$ ). 20 Two LRRK2 variants in our 31 SNPs of interest were significantly associated with outcomes. LRRK2 21 p.G2019S (rs34637584) was associated with higher odds of having a family history of Parkinson's disease 22 (OR 3.54 [1.72, 7.29],  $P = 6.06 \times 10^{-4}$ ) and the T allele of rs76904798 (intergenic at the 5' end of *LRRK2*) 23 was associated with a higher HR of reaching HY3 (HR 1.33 [1.16, 1.52] for the T allele,  $P = 5.27 \times 10^{-5}$ ). 24 Age at onset was inversely associated with the Z-value of the genetic risk score (-0.60 [-0.89, -0.31] years 25 per +1 SD,  $P = 5.33 \times 10^{-5}$ ). Moreover, it was associated with rs34311866 (*TMEM175* p.M393T), the C

allele of rs199347 (intronic region of *GPNMB*), and the G allele of rs1106180 (intronic region of
 *CCDC*62).

The majority (14/17) of associations showed good accord across cohorts (I2 < 50%) and the forest plots (Fig.1 and Fig.2) also illustrate this qualitatively. Furthermore, up to 13 iterations of the leave-one-out analysis assessed 15 associations of which outcomes were measured in more than two cohorts and showed a small range of betas. The max *P*-value of 13 iterations was less than 0.05 for all associations except for rs114138769 (intron of *PMVK*) and rs76763715 (*GBA* p.N370S) for wearing-off. A metaanalysis with a random effect model also detected nine associations after the same FDR correction, even though the model is more conservative than a fixed model.

10

#### 11 **Discussion**

12 We conducted a meta-analysis with 13 longitudinal patient cohorts and identified multiple associations

13 between genotypes and clinical phenotypic characteristics, including progression rates. Among these,

14 GBA coding variants showed clear associations with the rate of cognitive decline (binomial outcome, or

15 UPDRS part 1 score) as well as motor symptom progression (HY, HY3), consistent with previous

16 studies.<sup>12,21–25</sup>

17 In addition, we found associations between *GBA* variants and RBD and daytime sleepiness. A previous

18 cross-sectional study with 120 Ashkenazi-Jewish patients reported a higher frequency of RBDSQ-

19 detected RBD symptoms in *GBA* variant carriers,<sup>26</sup> Our finding suggests that *GBA* is associated not only

20 with baseline clinical presentation but also with disease progression.

21 This is the first study, to our knowledge, noting an association between GBA and daytime sleepiness. One

study reported an association between sleep problems (as assessed by the Parkinson's Disease Sleep Scale

23 (PDSS)) and GBA.<sup>27</sup> However, this scale is a combined measure of daytime sleepiness and other aspects

of sleep problems.

1 Finally, a GBA variant (p.N370S) was also associated with treatment-related complications of wearing-off 2 and dyskinesia. Two studies have reported the association of GBA variants with these complications, with one positive and one negative result.<sup>28,29</sup> The negative result may be due to insufficient power with only 3 4 19 patients with GBA mutations. 5 Overall, our study provides a distinct clinical profile of patients with GBA variants compared to those 6 without. We note that with 63 carriers for p.N370S, 166 for p.T408M, and 217 for p.E365K, we have a 7 reasonable power, but the number is yet not enough. And this may affect the results in seemingly different 8 magnitudes of associations and the association for different traits per variants (e.g., motor complications 9 with p.N370S and cognitive impairment with p.E365K). Another possible explanation is that even though 10 the effects are associated with the same gene, the biological activity or molecular mechanism could be 11 different. Such an example has already been reported for LRRK2 p.G2019S and p.G2385R.<sup>30</sup> 12 Aside from GBA variants, the associations between close intergenic (5' end) variant of LRRK2, 13 rs76904798, and the faster development of motor symptom. This variant is 4.3 kb upstream from 5' end 14 of LRRK2 and reported to be associated with LRRK2 gene expression changes in recent blood cis-eQTL study from the eQTLGen Consortium.<sup>31</sup> In contrast, we did not find an association between rs34637584, 15 16 LRRK2 coding mutation (p.G2019S) and motor progression. The p.G2019 variant is a rare variant (MAF 17 0.5% in our study) and our sample size was not adequate barring an extremely large effect size. The 18 intronic region variant of *PMVK*, rs114138760, and the development of wearing-off was another finding. 19 The biological effect of *PMVK* on PD has not been reported, but the variant is also located at close 20 proximity of the *GBA-SYT11* locus, so it is possible that its association was through a similar mechanism 21 as GBA. Including the results of cross-sectional analysis, the associations of age at onset with rs34311866 22 (*TMEM175*, p.M393T), rs199347 (intron of *GPNMB*), and rs11060180 (intron of *CCDC62*) were found. 23 *TMEM175* has been reported to impair lysosomal and mitochondrial function and increase  $\alpha$ -synuclein aggregatio,<sup>32</sup> although no functional data for this missense variant was studied. Interestingly, the variant 24 has recently been reported in another study as being associated with the age at onset.<sup>33</sup> rs199347 is an 25

eQTL increasing the brain expression of *GPNMB*,<sup>34</sup> suggesting a causal link. Regarding rs1160180, no
 functional data is available in this locus.

We also evaluated the association between genetic risk variants and clinical outcomes by two-step metaanalysis. This analysis is exploratory, and we acknowledge that this is biased towards the null due to power issues when partitioning studies randomly. However, we believe that it is helpful to assess the rigorousness of the associations we found in the primary analysis as well as to explore potential missed associations.

8 A strength of the current study was its design, incorporating multiple distinct independent Parkinson's 9 disease cohorts with longitudinal follow-ups. Although the cohorts contained patients at different disease 10 stages, and some of the definition of outcomes were not identical, we analyzed each cohort separately and 11 combined the results. Thus, the significant findings are consistent and applicable to the wider Parkinson's 12 disease populations. The forest plots showed that most of the estimates agree with each other despite the 13 relative differences in the cohort characteristics. Another strength is the size of the study. The total 14 number of genotyped and phenotyped Parkinson's disease patients (N = 4,307) is one of the largest to date 15 for an investigation of disease progression.

16 The limitations of our study were as follows. First, we only included patients of European ancestry. It is 17 uncertain whether the associations in the current study are also applicable to people from different ethnic 18 backgrounds and further research is needed. Second, the current analysis could not distinguish causality, 19 only basic associations. Different approaches, such as molecular-level assessment and Mendelian 20 randomization, are crucial. Third, interaction effects between genes and other factors are another 21 important research target not addressed in this report due to power constraints. For example, gene-by-22 smoking interactions for Parkinson's disease were indicated recently,<sup>35</sup> and highlight the importance of 23 correctly modeling gene-environment interactions. Finally, compared with the typical GWAS analysis 24 (which includes tens of thousands of cases) the number of participants was small, and the outcomes of 25 interest were not as simple or easily defined as with case-control distinctions in GWAS. Acknowledging 26 the limitations, the list of associations provided here is valuable as a foundation for further studies and as

an example that illustrates the potential of efforts to define the genetic basis of variability in presentation
 and course. Accounting for this variability, even in part, has the potential to positively impact etiology based clinical trials by reducing variability between placebo and treatment groups, and by providing
 better predictions of expected individual progression.

5

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25

#### 26 Appendix 1 Authors

NAME	ALL	Role	Contribution
Hirotaka Iwaki M.D. Ph.D.	Laboratory of Neurogenetics, National Institute on Aging, National Institutes of Health, Bethesda, MD, USA	Author	Literature search; study design; data analysis; data interpretation: writings
Cornelis Blauwendraat Ph.D.	Laboratory of Neurogenetics, National Institute on Aging, National Institutes of Health, Bethesda, MD, USA	Author	Literature search: data analysis; data interpretation; critical review
Hampton L. Leonard M.S.	Laboratory of Neurogenetics, National Institute on Aging, National Institutes of Health, Bethesda, MD, USA	Author	critical review
Ganqiang Liu Ph.D.	Precision Neurology Program, Harvard Medical School, Brigham and Women's Hospital, Boston, MA, USA	Author	data collection; critical review
Jodi Maple-Grødem Ph.D.	The Norwegian Centre for Movement Disorders, Stavanger University Hospital, Stavanger, Norway	Author	data collection; critical review
Jean-Christophe Corvol M.D. Ph.D.	Assistance-Publique Hôpitaux de Paris, ICM, INSERM UMRS 1127, CNRS 7225, ICM, Department of Neurology and CIC Neurosciences, Pitié-Salpêtrière Hospital, Paris, France	Author	data collection; critical review
Lasse Pihlstrøm M.D. Ph.D.	Department of Neurology, Oslo University Hospital, Oslo, Norway	Author	data collection; critical review
Marlies van Nimwegen Ph.D.	Department of Neurology, Donders Institute for Brain, Cognition, and Behaviour, Radboud University Medical Centre, Nijmegen, The Netherlands	Author	data collection; critical review
Samantha J. Hutten Ph.D.	Michael J Fox Foundation, New York, NY, USA	Author	data collection; critical review
Khanh-Dung H. Nguyen Ph.D.	Translational Genome Sciences, Biogen, Cambridge, MA, USA	Author	data collection; critical review
Jacqueline Rick Ph.D.	Department of Neurology University of Pennsylvania, Philadelphia, PA, USA	Author	data collection; critical review
Shirley Eberly M.S.	Department of Biostatistics and Computational Biology, University of Rochester, Rochester, NY, USA	Author	data collection; critical review
Faraz Faghri M.S.	Laboratory of Neurogenetics, National Institute on Aging,	Author	data collection; critical review

	National Institutes of Health		
	Bethesda, MD, USA		
Peggy Auinger	Department of Neurology, Center	Author	data collection; critical
M.S.	for Health + Technology, University		review
	of Rochester, Rochester, NY, USA		
Kirsten M. Scott	Department of Clinical	Author	data collection; critical
MRCP, M.Phil.	Neurosciences, University of		review
	Cambridge, John van Geest Centre		
	for Brain Repair, Cambridge, UK		
Ruwani	Department of Clinical	Author	data collection; critical
Wijeyekoon,	Neurosciences, University of		review
MRCP	Cambridge, John van Geest Centre		
	for Brain Repair, Cambridge, UK		
Vivianna M. Van	Department of Pathology and	Author	data collection; critical
Deerlin M.D. Ph.D.	Laboratory Medicine, Center for		review
	Neurodegenerative Disease		
	Research, Parelman School of		
	Medicine at the University of		
	Pennsylvania, Philadelphia, PA,		
Dono C. Hormondoz	USA Laboratory of Nouro consting	Author	data collections anitical
Dena G. Hernandez	National Institute on A sing	Author	
PII.D.	National Institutes of Health		leview
	Ratharda MD USA		
Aaron G. Dav-	Genetics and Pharmacogenomics	Author	data collection: critical
Williams Ph D	Merck Research Laboratory	Aution	review
winnanis i n.D.	Boston MA USA		
Alexis Brice M D	Institut du cerveau et de la moelle	Author	data collection: critical
	épinière ICM, Paris, France	1 1001101	review
Guido Alves M.D	The Norwegian Centre for	Author	data collection: critical
Ph.D.	Movement Disorders, Stavanger		review
	University Hospital, Stavanger,		
	Norway		
Alastair J. Noyce	Preventive Neurology Unit,	Author	data collection; critical
MRCP, Ph.D.	Wolfson Institute of Preventive		review
	Medicine, Queen Mary University		
	of London, London, UK		
Ole-Bjørn Tysnes	Department of Neurology,	Author	data collection; critical
M.D., Ph.D.	Haukeland University Hospital,		review
	Bergen, Norway		
Jonathan R. Evans	Department of Neurology,	Author	data collection; critical
MRCP, Ph.D.	Nottingham University NHS Trust,		review
	Nottingham, UK		
David P. Breen	Centre for Clinical Brain Sciences,	Author	data collection; critical
MRCP, Ph.D.	University of Edinburgh,		review
	Edinburgh, Scotland		

Karol Estrada Ph.D.	Translational Genome Sciences,	Author	data collection; critical
Claims E. Wasal	Biogen, Cambridge, MA, USA	Anthon	review
Claire E. wegel	Department of Medical and	Author	data collection; critical
MPH	Molecular Genetics, Indiana		review
Estaine Dealers	University, Indianapolis, IN, USA	A (1	
Fabrice Danjou	Institut du cerveau et de la moelle	Author	data collection; critical
M.D., Ph.D.	epiniere ICM, Paris, France	A .1	review
David K. Simon	Department of Neurology, Beth	Author	data collection; critical
M.D., Ph.D.	Israel Deaconess Medical Center,		review
	Boston, MA, USA	A .1	1 / 11 / 1
Bernard Ravina	Voyager Therapeutics, Cambridge,	Author	data collection; critical
M.D.	MA, USA		review
Mathias Toft M.D.,	Department of Neurology, Oslo	Author	data collection; critical
Ph.D.	University Hospital, Oslo, Norway		review
Peter Heutink Ph.D.	German Center for	Author	data collection; critical
	Neurodegenerative Diseases-		review
	Tubingen, Tuebingen, Germany		
Bastiaan R. Bloem	Department of Neurology, Donders	Author	data collection; critical
M.D., Ph.D.	Institute for Brain, Cognition, and		review
	Behaviour, Radboud University		
	Medical Centre, Nijmegen, The		
	Netherlands		
Daniel Weintraub	Department of Psychiatry,	Author	data collection; critical
M.D.	University of Pennsylvania School		review
	of Medicine, Philadelphia, PA, USA		
Roger A. Barker	Department of Clinical	Author	data collection; critical
MRCP, Ph.D.	Neurosciences, University of		review
	Cambridge, Cambridge, UK		
Caroline H.	Department of Clinical	Author	data collection; critical
Williams-Gray	Neurosciences, University of		review
MRCP, Ph.D.	Cambridge, Cambridge, UK		
Bart P. van de	Department of Neurology, Donders	Author	data collection; critical
Warrenburg M.D.,	Institute for Brain, Cognition, and		review
Ph.D.	Behaviour, Radboud University		
	Medical Centre, Nijmegen, The		
	Netherlands		
Jacobus J. Van	Department of Neurology, Leiden	Author	data collection; critical
Hilten M.D., Ph.D.	University Medical Center, Leiden,		review
	The Netherlands		
Clemens R.	Precision Neurology Program,	Author	data collection; critical
Scherzer M.D.	Harvard Medical School, Brigham		review
	and Women's Hospital, Boston.		
	MA, USA		
Andrew B.	Laboratory of Neurogenetics.	Author	study design: critical
Singleton Ph.D.	National Institute on Aging,		review

			National Institutes of Health, Bethesda, MD, USA		
	Mik	ke A. Nalls	Laboratory of Neurogenetics,	Author	study design; data
	Ph.	D.	National Institute on Aging,		analysis; data
			National Institutes of Health,		interpretation; critical
4			Bethesda, MD, USA		review
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15	Supp	blementary materials
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<ol> <li>15</li> <li>16</li> <li>17</li> <li>18</li> <li>19</li> <li>20</li> <li>21</li> <li>22</li> <li>23</li> <li>24</li> <li>25</li> </ol>	Supr Appen Table Table Table Table FDR-c Fig. 1: DATA	Dementary materials dix e-1: The cohort information and details e-1: Allele frequency of 31 SNPs and the effect size e-2: Cohort specific definitions of binomial outcomes e-3: Model specifications e-4: Meta-analysis for 13 cohorts and the results of sensitivity analysis (The estimated betas and corrected P) re legends Forest plots for GBA variants and symptoms of Parkinson's disease TOP, Deprenyl and Tocopherol Antioxidative Therapy of Parkinsonism; DIGPD, Drug Interaction
<ol> <li>15</li> <li>16</li> <li>17</li> <li>18</li> <li>19</li> <li>20</li> <li>21</li> <li>21</li> <li>22</li> <li>23</li> <li>24</li> <li>25</li> <li>26</li> </ol>	Supp Appen Table Table Table Table FDR-c Fig. 1: DATA with C	<ul> <li>blementary materials</li> <li>dix e-1: The cohort information and details</li> <li>e-1: Allele frequency of 31 SNPs and the effect size</li> <li>e-2: Cohort specific definitions of binomial outcomes</li> <li>e-3: Model specifications</li> <li>e-4: Meta-analysis for 13 cohorts and the results of sensitivity analysis (The estimated betas and corrected P)</li> <li>re legends</li> <li>Forest plots for GBA variants and symptoms of Parkinson's disease</li> <li>ATOP, Deprenyl and Tocopherol Antioxidative Therapy of Parkinsonism; DIGPD, Drug Interaction Genes in Parkinson's Disease; HBS, Harvard Biomarkers Study; NET-PD_LS1, NIH Exploratory</li> </ul>

1	ParkWest, The Norwegian ParkWest study; PDBP, Parkinson's Disease Biomarker Program; PICNICS,
2	Parkinsonism Incidence and Cognitive and Non-motor heterogeneity In Cambridgeshire; PPMI,
3	Parkinson's progression markers initiative; PreCEPT, Parkinson Research Examination of CEP-1347
4	Trial and PostCEPT; ProPark, Profiling Parkinson's disease study; and Udall, Morris K. Udall Centers for
5	Parkinson's Research.
6	
7	Fig. 2: Forest plots for non-GBA risk variants/genetic risk score and symptoms or features of Parkinson's
8	disease.
9	DATATOP, Deprenyl and Tocopherol Antioxidative Therapy of Parkinsonism; DIGPD, Drug Interaction
10	with Genes in Parkinson's Disease; HBS, Harvard Biomarkers Study; NET-PD_LS1, NIH Exploratory
11	Trials in Parkinson's Disease Large Simple Study 1; Oslo, Oslo PD study; ParkFit, ParkFit study;
12	ParkWest, The Norwegian ParkWest study; PDBP, Parkinson's Disease Biomarker Program; PICNICS,
13	Parkinsonism Incidence and Cognitive and Non-motor heterogeneity In Cambridgeshire; PPMI,
14	Parkinson's progression markers initiative; PreCEPT, Parkinson Research Examination of CEP-1347
15	Trial and PostCEPT; ProPark, Profiling Parkinson's disease study; and Udall, Morris K. Udall Centers for

16 Parkinson's Research.

# Table 1. Summary characteristics of 13 cohorts

				NET-PD							PreCEPT/		
	DATATOP	DIGPD	HBS	LS1	Oslo	ParkFit	ParkWest	PDBP	PICNICS	PPMI	PostCEPT	ProPark	Udall
Cohort size, n	440	311	580	406	317	335	150	422	120	357	321	296	252
Follow-up duration, years	1.22 (0.41)	2.19 (1.51)	1.53 (0.87)	4.48 (1.45)	4.64 (3.10)	1.97 (0.00)	3.04 (0.09)	2.06 (1.70)	3.04 (1.63)	4.87 (1.35)	6.79 (0.95)	4.62 (1.14)	3.77 (1.81)
Female, n (%)	146 (33.2)	121 (38.9)	201 (34.7)	148 (36.5)	107 (33.8)	110 (32.8)	57 (38.0)	174 (41.2)	43 (35.8)	121 (33.9)	106 (33.0)	105 (35.5)	73 (29.0)
Family history, n (%)	86 (20.9)	69 (22.3)	148 (25.5) 62.16	59 (14.5)	43 (14.0) 54.33	-	17 (11.3)	54 (12.8) 58.51	19 (15.8)	48 (13.5) 61.45	93 (29.2)	76 (25.9) 53.14	71 (28.4)
Age at onset, years	58.65 (9.17)	59.41 (9.80)	(10.46)	60.64 (9.45)	(10.06)	60.79 (8.65)	67.27 (9.26)	(10.28)	68.94 (9.34)	(9.55)	59.47 (9.22)	(10.60)	64.26 (8.64)
Baseline from diagnosis, years	1.14 (1.17)	2.60 (1.57)	4.09 (4.63)	1.50 (1.00)	10.13 (6.04)	5.18 (4.44)	0.13 (0.12)	5.68 (5.64)	0.23 (0.48)	0.54 (0.54)	0.80 (0.83)	6.56 (4.67)	6.21 (5.38)
Levodopa use, n (%)	0 (0.0)	198 (63.9)	415 (71.6)	207 (51.2)	-	-	0 (0.0)	255 (60.4)	36 (30.0)	0 (0.0)	0 (0.0)	202 (68.2)	215 (85.3)
Dopamine agonist use, n (%)	0 (0.0)	228 (73.3)	224 (38.6)	280 (69.3)	-	-	0 (0.0)	61 (14.5)	22 (18.3)	0 (0.0)	1 (0.3)	222 (75.0)	118 (46.8)
Modified Hoehn and Yahr Scale	1.61 (0.53)	1.75 (0.55)	2.14 (0.64)	-	2.19 (0.64)	2.08 (0.33)	1.86 (0.58)	2.04 (0.69)	1.64 (0.67)	1.55 (0.50)	1.75 (0.48)	2.51 (0.79)	2.29 (0.68)
UPDRS1	-	7.69 (4.50)	1.70 (1.59)	1.31 (1.45)	-	-	1.95 (1.76)	9.90 (6.11)	-	5.40 (3.97)	0.84 (1.19)	-	1.92 (1.99)
UPDRS2	-	7.72 (4.66)	9.21 (5.23)	7.29 (3.86)	-	-	8.19 (4.22)	11.14 (8.01)	-	5.80 (4.11)	6.11 (3.20)	-	10.74 (7.13)
UPDRS3	-	(10.23)	(9.58)	(8.32)	(10.30)	-	(9.77)	(13.08)	-	20.88 (9.00)	(7.65)	-	(11.09)
UPDRS4	-	0.66 (2.56)	2.25 (2.05)	1.34 (1.49)	-	-	0.57 (1.14)	2.20 (3.17)	-	-	-	-	2.02 (2.75)
MDS_UPDRS total	-	36.43 (16.02)	-	-	-	-	-	46.88 (24.04)	47.27 (17.97)	-	-	-	-
UPDRS total	24.08 (11.56)	-	32.33 (14.28)	(11.62)	-	(10.10)	(13.91)	-	-	-	(10.10)	-	32.04 (18.28)
MMSE	28.99 (1.35)	28.38 (1.73)	28.35 (2.17)	-	-	28.09 (1.61)	27.88 (2.27)	-	28.71 (1.43)	-	29.29 (1.07)	27.05 (2.50)	26.83 (3.50)
MoCA	- 91.55	- 80.55	-	-	-	-	-	25.44 (3.40)	-	27.17 (2.23) 93.18	-	-	24.37 (3.63)
SEADL	(6.49)	(29.02)	-	(6.06)	-	-	(7.35)	(13.10)	-	(5.91)	(5.26)	-	(17.56)
Hyposmia, n (%)	-	89 (28.9)	-	-	-	-	54 (36.0)	276 (65.4)	-	164 (45.9)	-	173 (63.8)	69 (67.0)
Cognitive impairment, n (%)	26 (5.9)	3 (1.0)	74 (13.0)	29 (7.1)	-	55 (16.4)	27 (18.0)	96 (22.7)	11 (9.2)	28 (7.8)	3 (0.9)	77 (27.0)	29 (11.5)
Motor fluctuation, n (%)	-	40 (12.9)	228 (39.9)	103 (25.4)	-	-	4 (2.7)	129 (48.1)	1 (0.8)	-	-	94 (32.4)	75 (35.4)
Dyskinesia, n (%)	4 (0.9)	13 (4.2)	207 (36.2)	5 (1.2)	-	-	2 (1.3)	196 (46.4)	0 (0.0)	-	-	81 (27.6)	44 (22.8)
Depression, n (%)	12 (2.7)	97 (31.6)	35 (10.9)	40 (9.9)	-	-	20 (13.3)	49 (11.6)	27 (22.5)	113 (31.7)	73 (22.7)	48 (16.3)	63 (25.0)
Restless legs syndrome, n (%)	-	44 (14.5)	37 (10.9)	-	-	-	-	91 (23.3)	-	23 (6.4)	-	-	-

Constipation, n (%)	9 (2.0)	62 (20.3)	-	-	-	-	17 (11.3)	239 (56.6)	29 (24.2)	113 (31.7)	-	138 (46.6)	-
REM sleep behavior disorder, n	_	_	_	_	_	_	_	197 (50 5)	_	03 (26 1)	_	_	_
	-	-	-	-	-	-	-	197 (30.3)	-	93 (20.1)	-	-	-
Daytime Sleepiness, n (%)	5 (1.1)	138 (44.8)	-	-	-	-	25 (16.7)	165 (39.1)	25 (20.8)	55 (15.4)	-	126 (42.6)	-
Insomnia, n (%)	11 (2.5)	107 (35.1)	202 (35.1)	-	-	-	45 (30.0)	295 (69.9)	62 (51.7)	78 (21.8)	-	83 (28.0)	-
HY>=3.0, n (%)	0 (0.0)	4 (1.3)	71 (12.4)	12 (3.0)	22 (14.5)	17 (5.1)	11 (7.3)	71 (16.8)	13 (10.8)	1 (0.3)	0 (0.0)	117 (40.8)	57 (23.0)

Continuous variables were summarized in Mean (SD). MMSE, Mini Mental State Examination; Montreal Cognitive Assessment, SEADL, Schwab and England Activities of Daily Living Scale; UPDRS, Unified Parkinson Disease Rating Scale; MDS-UPDRS, Movement Disorder Society revised version of UPDRS.

Table 2. Meta-a-lysis for 13 cohorts and the results of sensitivity analysis												
					Fixed effect model			Leave-one-out a-lysis		Random effect model		
Outcome	rsNo	Known gene or nearest gene	N of cohorts	Scale of the effect	Estimate [95% C.I.]	Р	Test of Homogeneity	I <sup>2</sup> (%)	Estimate [Min, Max]	Max P	Estimate [95% C.I.]	Р
Wearing-off	rs11413876 0	intron_PMVK	9	Multiplicative (HR)	1.66 [1.19, 2.31]	2.62E-03	0.322	12.58	1.66 [1.44, 1.81]	6.22E-02	1.65 [1.14, 2.38]	7.39E-03
Dyskinesia	rs76763715	GBA:N370S	8	Multiplicative (HR)	3.01 [1.81, 5.01]	2.17E-05	0.011	60.53	3.00 [1.98, 4.05]	2.26E-02	2.49 [1.06, 5.86]	3.73E-02
HY>=3.0	rs76763715	GBA:N370S	6	Multiplicative (HR)	4.59 [2.60, 8.10]	1.58E-07	0.654	0.00	4.59 [4.02, 5.41]	2.00E-05	4.59 [2.60, 8.10]	1.58E-07 *
Wearing-off	rs76763715	GBA:N370S	6	Multiplicative (HR)	2.03 [1.28, 3.21]	2.56E-03	0.021	62.70	2.02 [1.61, 2.65]	8.67E-02	1.92 [0.85, 4.33]	1.14E-01
Daytime sleepiness	rs76763715	GBA:N370S	6	Multiplicative (HR)	3.28 [1.69, 6.34]	4.24E-04	0.467	0.00	3.30 [2.85, 4.38]	3.75E-03	3.28 [1.69, 6.34]	4.24E-04 *
HY>=3.0	rs75548401	GBA:T408M	8	Multiplicative (HR)	1.93 [1.34, 2.78]	4.40E-04	0.208	32.43	1.93 [1.70, 2.41]	1.08E-02	1.96 [1.22, 3.14]	5.22E-03
pRBD (Baseline)	rs75548401	GBA:T408M	2	Multiplicative (OR)	6.48 [2.04, 20.60]	1.53E-03	0.118	59.06	-	-	6.25 [1.02, 38.20]	4.72E-02
НҮ	rs2230288	GBA:E365K	12	Continuous	0.10 [0.04, 0.16]	1.53E-03	0.017	48.90	0.10 [0.08, 0.11]	1.02E-02	0.11 [0.02, 0.21]	1.88E-02
Cognitive impairment (Baseline)	rs2230288	GBA:E365K	8	Multiplicative (OR)	2.37 [1.53, 3.66]	1.09E-04	0.794	0.00	2.37 [2.20, 2.59]	8.57E-04	2.37 [1.53, 3.66]	1.09E-04 *
Cognitive impairment	rs2230288	GBA:E365K	9	Multiplicative (HR)	2.78 [1.88, 4.11]	2.97E-07	0.555	0.00	2.78 [2.41, 2.98]	5.08E-05	2.78 [1.88, 4.11]	2.97E-07 *
pRBD	rs2230288	GBA:E365K	2	Multiplicative (HR)	2.57 [1.43, 4.63]	1.69E-03	0.665	0.00	-	-	2.57 [1.43, 4.63]	1.69E-03 *
Age at onset	rs34311866	TMEM175:M393T	13	Continuous	-0.72 [-1.21, -0.23]	3.87E-03	0.515	0.00	-0.72 [-0.83, -0.58]	2.83E-02	-0.72 [-1.21, -0.23]	
Age at onset	rs199347	intron_GPNMB	12	Continuous	0.70 [0.27, 1.14]	1.42E-03	0.824	0.00	0.70 [0.60, 0.77]	1.12E-02	0.70 [0.27, 1.14]	1.42E-03 *
HY>=3.0	rs76904798	5_LRRK2	13	Multiplicative (HR)	1.33 [1.16, 1.52]	5.27E-05	0.049	43.15	1.33 [1.26, 1.43]	1.64E-03	1.34 [1.11, 1.63]	2.80E-03 *
Family History	rs34637584	LRRK2:G2019S	8	Multiplicative (OR)	3.54 [1.72, 7.29]	6.06E-04	0.856	0.00	3.53 [2.78, 3.98]	1.66E-02	3.54 [1.72, 7.29]	6.06E-04 *
Age at onset	rs11060180	intron_CCDC62	13	Continuous	0.62 [0.21, 1.03]	3.32E-03	0.054	42.60	0.62 [0.49, 0.75]	2.74E-02	0.55 [-0.00, 1.11]	5.14E-02
Age at onset	Genetic risk score		13	Continuous	-0.60 [-0.89, -0.31]	5.33E-05	0.749	0.00	-0.60 [-0.65, -0.52]	9.02E-04	-0.60 [-0.89, - 0.31]	5.33E-05 *

Possible REM sleep behavior disorder (pRBD) was only available in 2 cohorts and a leave-one-out a-lysis was not conducted for this outcome

HY, Hoehn and Yahr Scale;

\* Significant after FDR adjustment in a random effect model