

Total Zinc Intake May Modify the Glucose-Raising Effect of a Zinc Transporter (*SLC30A8*) Variant

A 14-Cohort Meta-analysis

Stavroula Kanoni,^{1,2} Jennifer A. Nettleton,³ Marie-France Hivert,⁴ Zheng Ye,⁵ Frank J.A. van Rooij,^{6,7} Dmitry Shungin,^{8,9,10} Emily Sonestedt,⁹ Julius S. Ngwa,¹¹ Mary K. Wojczynski,¹² Rozenn N. Lemaitre,¹³ Stefan Gustafsson,¹⁴ Jennifer S. Anderson,¹⁵ Toshiko Tanaka,¹⁶ George Hindy,⁹ Georgia Saylor,¹⁵ Frida Renstrom,^{8,9,17} Amanda J. Bennett,¹⁸ Cornelia M. van Duijn,^{6,7} Jose C. Florez,^{19,20,21} Caroline S. Fox,^{22,23} Albert Hofman,^{6,7} Ron C. Hoogeveen,^{24,25} Denise K. Houston,²⁶ Frank B. Hu,¹⁷ Paul F. Jacques,²⁷ Ingegerd Johansson,¹⁰ Lars Lind,²⁸ Yongmei Liu,²⁹ Nicola McKeown,^{27,30} Jose Ordovas,³¹ James S. Pankow,³² Eric J.G. Sijbrands,^{7,33} Ann-Christine Syvänen,²⁸ André G. Uitterlinden,^{6,7,33} Mary Yannakoulia,¹ M. Carola Zillikens,³³ the MAGIC Investigators,* Nick J. Wareham,⁵ Inga Prokopenko,^{18,34} Stefania Bandinelli,³⁵ Nita G. Forouhi,⁵ L. Adrienne Cupples,^{11,22} Ruth J. Loos,⁵ Goran Hallmans,⁸ Josée Dupuis,^{11,22} Claudia Langenberg,⁵ Luigi Ferrucci,¹⁶ Stephen B. Kritchevsky,²⁶ Mark I. McCarthy,^{18,34,36} Erik Ingelsson,¹⁴ Ingrid B. Borecki,¹² Jacqueline C.M. Witteman,^{6,7} Marju Orho-Melander,⁹ David S. Siscovick,¹³ James B. Meigs,³⁷ Paul W. Franks,^{8,9,17} and George V. Dedoussis¹

OBJECTIVE—Many genetic variants have been associated with glucose homeostasis and type 2 diabetes in genome-wide association studies. Zinc is an essential micronutrient that is important for β -cell function and glucose homeostasis. We tested the hypothesis that zinc intake could influence the glucose-raising effect of specific variants.

RESEARCH DESIGN AND METHODS—We conducted a 14-cohort meta-analysis to assess the interaction of 20 genetic variants known to be related to glycemic traits and zinc metabolism with dietary zinc intake (food sources) and a 5-cohort meta-analysis to assess the interaction with total zinc intake (food sources

and supplements) on fasting glucose levels among individuals of European ancestry without diabetes.

RESULTS—We observed a significant association of total zinc intake with lower fasting glucose levels (β -coefficient \pm SE per 1 mg/day of zinc intake: -0.0012 ± 0.0003 mmol/L, summary P value = 0.0003), while the association of dietary zinc intake was not significant. We identified a nominally significant interaction between total zinc intake and the *SLC30A8* rs11558471 variant on fasting glucose levels (β -coefficient \pm SE per A allele for 1 mg/day of greater total zinc intake: -0.0017 ± 0.0006 mmol/L, summary interaction P value = 0.005); this result suggests a stronger

From the ¹Department of Nutrition-Dietetics, Harokopio University, Athens, Greece; the ²Wellcome Trust Sanger Institute, Hinxton, U.K.; the ³Division of Epidemiology, Human Genetics, and Environmental Sciences, University of Texas Health Science Center, Houston, Texas; the ⁴Department of Medicine, Division of Endocrinology, Université de Sherbrooke, Sherbrooke, Canada; the ⁵Medical Research Council Epidemiology Unit, Institute of Metabolic Science, Addenbrooke's Hospital, Cambridge, U.K.; the ⁶Department of Epidemiology, Erasmus Medical Center, Rotterdam, the Netherlands; ⁷The Netherlands Genomics Initiative—Sponsored Netherlands Consortium for Healthy Aging (NGL-NCHA), Leiden, the Netherlands; the ⁸Genetic Epidemiology and Clinical Research Group, Department of Public Health and Clinical Medicine, Section of Medicine, Umeå University Hospital, Umeå, Sweden; the ⁹Department of Clinical Sciences, Lund University, Malmö, Sweden; the ¹⁰Department of Odontology, Umeå University, Umeå, Sweden; the ¹¹Department of Biostatistics, Boston University School of Public Health, Boston, Massachusetts; the ¹²Division of Statistical Genomics, Department of Genetics, Washington University School of Medicine, St. Louis, Missouri; the ¹³Cardiovascular Health Research Unit, Department of Medicine and Epidemiology, University of Washington, Seattle, Washington; the ¹⁴Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, Stockholm, Sweden; the ¹⁵Baptist Medical Center, Wake Forest University, Winston-Salem, North Carolina; the ¹⁶Clinical Research Branch, National Institute on Aging, Baltimore, Maryland; the ¹⁷Department of Nutrition, Harvard School of Public Health, Boston, Massachusetts; the ¹⁸Oxford Centre for Diabetes, Endocrinology, and Metabolism, University of Oxford, Churchill Hospital, Oxford, U.K.; the ¹⁹Diabetes Unit, Center for Human Genetic Research and Diabetes Research Center, Massachusetts General Hospital, Boston, Massachusetts; the ²⁰Program in Medical and Population Genetics, Broad Institute, Cambridge, Massachusetts; the ²¹Department of Medicine, Harvard Medical School, Boston, Massachusetts; the ²²Framingham Heart Study, National Heart, Lung, and Blood Institute, Framingham, Massachusetts; the ²³Division of Endocrinology, Diabetes, and Hypertension, Brigham and Women's Hospital, Boston, Massachusetts; the ²⁴Section of Atherosclerosis and Vascular Medicine, Department of Medicine, Baylor College of Medicine, Houston, Texas;

the ²⁵Center for Cardiovascular Disease Prevention, Methodist DeBakey Heart Center, Houston, Texas; the ²⁶Sticht Center on Aging, Department of Internal Medicine, Section on Gerontology and Geriatric Medicine, Wake Forest University School of Medicine, Winston-Salem, North Carolina; the ²⁷Nutrition Epidemiology Program, U.S. Department of Agriculture Human Nutrition Research Center on Aging (USDA HNRCA) at Tufts University, Boston, Massachusetts; the ²⁸Department of Medical Sciences, Uppsala University, Uppsala, Sweden; the ²⁹Department of Epidemiology and Prevention, Division of Public Health Sciences, Wake Forest University Health Sciences, Winston-Salem, North Carolina; the ³⁰Friedman School of Nutrition Science and Policy, Tufts University, Boston, Massachusetts; the ³¹Nutrition and Genomics Laboratory, Jean Mayer USDA HNRCA at Tufts University, Boston, Massachusetts; the ³²Department of Epidemiology, University of Minnesota, Minneapolis, Minnesota; the ³³Department of Internal Medicine, Erasmus Medical Center, Rotterdam, the Netherlands; the ³⁴Wellcome Trust Centre for Human Genetics, University of Oxford, Oxford, U.K.; the ³⁵Geriatric Unit, Azienda Sanitaria Firenze, Florence, Italy; the ³⁶Oxford National Institute for Health Research Biomedical Research Centre, Churchill Hospital, Oxford, U.K.; and the ³⁷General Medicine Division, Clinical Epidemiology Unit and Diabetes Research Unit, Massachusetts General Hospital, Boston, Massachusetts.

Corresponding authors: Stavroula Kanoni, stavroula.kanoni@sanger.ac.uk; Paul W. Franks, paul.franks@med.lu.se; and George V. Dedoussis, dedoussi@hua.gr.

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P.W.F. and G.V.D. contributed equally to this work.

*A complete list of the MAGIC Investigators and institutions is provided in the Supplementary Data.

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inverse association between total zinc intake and fasting glucose in individuals carrying the glucose-raising A allele compared with individuals who do not carry it. None of the other interaction tests were statistically significant.

CONCLUSIONS—Our results suggest that higher total zinc intake may attenuate the glucose-raising effect of the rs11558471 *SLC30A8* (zinc transporter) variant. Our findings also support evidence for the association of higher total zinc intake with lower fasting glucose levels. *Diabetes* 60:2407–2416, 2011

Chronic elevations in fasting or postprandial glucose levels are the cardinal features of type 2 diabetes (T2D), a common complex disease caused by the interplay of genetic and lifestyle factors. Genome-wide association studies (GWAS) have identified genetic loci reproducibly associated with glycaemic traits or T2D (1,2). These studies improved our understanding of the mechanisms underlying impaired glucose homeostasis and T2D, potentially aiding the development of novel and individualized medical therapies (3,4).

Modifiable lifestyle factors, such as diet and physical activity, influence glucose homeostasis and thus represent important targets for T2D prevention and management (5). Zinc is an essential trace element found in most foods that facilitates catalytic, structural, and transcriptional actions within cells (6). Zinc is a critical component of the catalytic site of >300 enzymes, including pancreatic carboxypeptidases and RNA polymerases; coordinates with protein domains; facilitates protein folding; produces structures such as zinc fingers; and regulates the expression of metallothioneins (7–9). Zinc is necessary in β -cells for insulin crystallization in hexamers (10). Moreover, it is cosecreted with insulin and exerts insulinomimetic and antioxidant actions and participates in the regulation of β -cell mass (11,12). Zinc homeostasis is impaired in animal models of diabetes and in humans with diabetes (13,14). Indeed, zinc supplementation studies in animal models support a protective effect of zinc against T2D, and in humans plasma levels of zinc are inversely associated with the risk of T2D (13). Nevertheless, a causal link between zinc and T2D in humans is not well established. Data from population-based studies indicate that dietary and total zinc intake (food sources and supplements) may reduce T2D risk (15–17), but the few intervention studies investigating the effect of zinc supplementation on glucose metabolism, insulin homeostasis, or T2D risk have yielded inconsistent results (13,18). Therefore, although zinc supplementation is a potential therapeutic and preventive target for T2D, additional studies are needed before population-wide recommendations regarding zinc intake can be advocated.

Knowledge of gene-environment interactions may enhance our understanding of disease etiology and pathogenesis, as well as help personalize interventions. We recently have demonstrated a favorable association of whole-grain intake with fasting glucose and insulin and observed a potential interaction between the rs780094 *GCKR* variant and whole-grain intake on fasting insulin concentrations (19). Zinc intake previously has been shown to interact with genetic variants in relation to chronic diseases and inflammatory biomarkers (8,20–22) but not in the context of glycaemic traits.

To address these gaps in the literature, we conducted a meta-analysis that included 14 cohorts, totaling up to 45,821

participants, to test the hypothesis that zinc intake modifies the cross-sectional association between fasting glucose levels and genetic variants known to be related to glycemia and zinc metabolism (2) in individuals of European descent without T2D.

RESEARCH DESIGN AND METHODS

The study sample for the current cross-sectional meta-analysis was combined from 14 cohort studies (Supplementary Table 1), the majority of which are included in the CHARGE (Cohorts for Heart and Aging Research in Genome Epidemiology) Consortium (23) and/or the Meta-Analyses of Glucose and Insulin-Related Traits Consortium (MAGIC). Ethical approval was obtained by local bioethical committees, and all participants provided signed informed consent. Participants included in the current analyses did not have diabetes (defined as diagnosed self-reported diabetes and/or fasting glucose levels ≥ 7 mmol/L and/or use of antidiabetes medications) and had dietary data that met quality-control criteria (Supplementary Table 1).

Anthropometric and blood glucose measurements. We calculated BMI as measured weight divided by the square of measured height (kg/m^2) in all cohorts. We determined blood glucose levels in fasting plasma samples, with the exceptions of participants in the Family Heart Study (FamHS), the Cardiovascular Health Study (CHS), the Rotterdam Study, and the Atherosclerosis Risk in Communities (ARIC) study, for whom we used fasting serum samples, and the Malmö Diet and Cancer (MDC) study, in which we used fasting whole-blood samples (Supplementary Table 1).

Dietary assessment and zinc intake estimation. Dietary data were collected (Supplementary Table 1) and processed to estimate the mean daily dietary zinc intake using an appropriate nutrient/food composition database for each cohort (Supplementary Table 1). Supplemental zinc intake (singly or in a multiple-nutrient supplement) was recorded in five cohorts and was quantified in milligrams per day. Total zinc intake was calculated as the sum of dietary and supplemental intake.

Genotyping and imputation. Twenty single nucleotide polymorphisms (SNPs) were included in the current meta-analysis (Tables 2 and 3 and Supplementary Data). Among these, 18 SNPs were recently associated with fasting glucose and/or fasting insulin levels in a large GWAS and meta-analysis study reported by the MAGIC (2). Two SNPs (rs10493846 and rs11167682) were selected for their potential role in zinc metabolism (24,25). A detailed description of the genotyping methods used for each study is provided in Supplementary Table 1.

Statistical analysis. All cohort-specific statistical analyses followed a uniform analytical plan. Linear regression models were applied to estimate the magnitude of the cross-sectional association of dietary and total zinc intake with fasting glucose levels, as well as the magnitude of the first-order SNP interactions (assuming an additive model) with dietary and total zinc intake on fasting glucose levels. All models were adjusted for age, sex, field center (in the ARIC study, the CHS, the FamHS, the Health Aging and Body Composition [Health ABC] study, and the Aging In the CHIANTI area [InCHIANTI] study), and family structure by principal components (in the Framingham Heart Study [FHS] and the FamHS). Associations were further adjusted for BMI levels to limit potentially confounding effects of adiposity.

Meta-analysis. The sample size for the association analysis of dietary zinc intake (food sources) with fasting glucose was 46,021 participants. The sample size for the interaction analysis between dietary zinc intake and SNPs ranged from 27,010 to 45,821 participants. Corresponding association and interaction analyses of total zinc intake (food sources and supplements) included 34,533 and 18,158 to 34,333 participants, respectively. We performed power calculations using Quanto version 1.2 (<http://hydra.usc.edu/gxe>). Accordingly, at 80% power, $P < 0.05$ detected a difference of 0.0013 mmol/L on fasting glucose levels per 1 mg/day higher dietary zinc intake and 0.0008 mmol/L in fasting glucose levels per 1 mg/day higher total zinc intake. At the same power and critical α , our study was large enough to detect a minimal interaction effect of 0.0031 mmol/L between dietary zinc intake (per 1 mg/day) and SNPs (per effect allele) on fasting blood glucose levels. Our study had 80% power to detect a minimal interaction effect of 0.0012 mmol/L between total zinc intake (per 1 mg/day) and SNPs (per effect allele) on fasting glucose levels.

We conducted an inverse variance-weighted, fixed-effect, meta-analyses on summary statistics provided by each cohort (METAL software, www.sph.umich.edu/csg/abecasis/metal/, for SNPs and zinc intake-interaction effects; SPSS 18.0, SPSS, Chicago, IL, for zinc-intake effects). Heterogeneity was estimated by Cochran's Q statistic and quantified by the I^2 statistic (26). Statistical significance was defined as a P value ≤ 0.0025 (0.05 per 20 tests), after Bonferroni correction for multiple testing.

TABLE 1
Descriptive characteristics of the 14 participating cohorts

| Cohorts | N | Age (years) | Women (%) | Fasting glucose (mmol/L)* | Dietary zinc intake (mg/day) | Total zinc intake (mg/day) | Zinc supplement users (%) | BMI (kg/m ²) |
|-----------------|--------|-------------|-----------|---------------------------|------------------------------|----------------------------|---------------------------|--------------------------|
| FHS | 5,835 | 46.1 (11.5) | 54.7 | 5.2 (0.5) | 12.4 (5.2) | 15.9 (11.2) | 21.8 | 26.7 (5.0) |
| MDC | 4,867 | 57.5 (5.9) | 60.0 | 5.5 (0.5) | 11.4 (3.4) | 13.4 (6.1) | 18.1 | 25.4 (3.8) |
| FamHS | 2,094 | 50.1 (13.1) | 55.5 | 5.2 (0.5) | 11.4 (5.6) | — | — | 27.3 (5.1) |
| CHS | 1,755 | 71.2 (4.5) | 63.8 | 5.1 (0.6) | 11.2 (6.3) | 17.3 (16.5) | 26.0 | 26.4 (4.3) |
| InCHIANTI | 1,071 | 67.7 (15.8) | 56.3 | 4.8 (0.6) | 10.9 (3.4) | — | 0 | 27.0 (4.1) |
| Rotterdam Study | 2,345 | 65.4 (6.6) | 58.3 | 5.5 (0.5) | 10.7 (2.7) | — | 0 | 26.5 (4.0) |
| ARIC | 6,088 | 60.2 (5.6) | 54.4 | 5.5 (0.6) | 10.5 (4.5) | 13.9 (11.0) | 18.5 | 27.6 (5.0) |
| PIVUS† | 770 | 70.2 (0.2) | 51.0 | 5.0 (0.6) | 10.4 (2.6) | — | — | 26.8 (4.1) |
| Health ABC | 1,263 | 74.8 (2.9) | 50.6 | 5.1 (0.6) | 10.3 (5.1) | — | — | 26.2 (4.0) |
| Fenland Study | 1,071 | 45.0 (7.3) | 56.1 | 4.9 (0.5) | 9.4 (3.0) | — | — | 27.0 (4.9) |
| ULSAM‡ | 931 | 71.0 (0.6) | 0.0 | 5.4 (0.6) | 9.3 (2.4) | — | — | 26.0 (3.2) |
| GHRAS§ | 856 | 71.8 (7.5) | 71.2 | 5.8 (1.6) | 8.9 (2.7) | — | 0 | 29.7 (4.8) |
| GENDAI | 1,087 | 11.2 (0.7) | 53.2 | 4.8 (0.5) | 8.7 (3.4) | — | 0 | 20.0 (3.4) |
| GLACIER¶ | 15,988 | 52.3 (8.8) | 60.2 | 5.4 (0.6) | 8.3 (3.0) | 8.7 (3.3) | 8.3 | 25.9 (4.1) |

Data are means (SD). —, not available. *Fasting glucose was measured at the baseline examination (when dietary assessment was conducted) in all cohorts, except for the MDC study and Rotterdam Study, in which fasting glucose was measured ~7 months and 6 years, respectively, after the dietary assessment and for the ARIC study, in which examination “three” dietary/supplement data and glucose measurements were used. †Prospective Investigation of the Vasculature in Uppsala Seniors; ‡Uppsala Longitudinal Study of Adult Men; §Greek Health Randomized Aging Study; ||Gene-Diet Investigation on Childhood Obesity; ¶Gene-Lifestyle Interactions and Complex Traits Involved in Elevated Disease Risk.

RESULTS

The descriptive characteristics of the 14 participating cohorts are summarized in Table 1. Mean dietary zinc intake (food sources) was comparable across cohorts (Fig. 1), ranging

from 8.3 to 12.4 mg/day, with an average dietary zinc intake of 10.3 mg/day. Mean total zinc intake (food sources and supplements) ranged from 8.7 to 17.3 mg/day, resulting in an average total zinc intake of 13.8 mg/day.

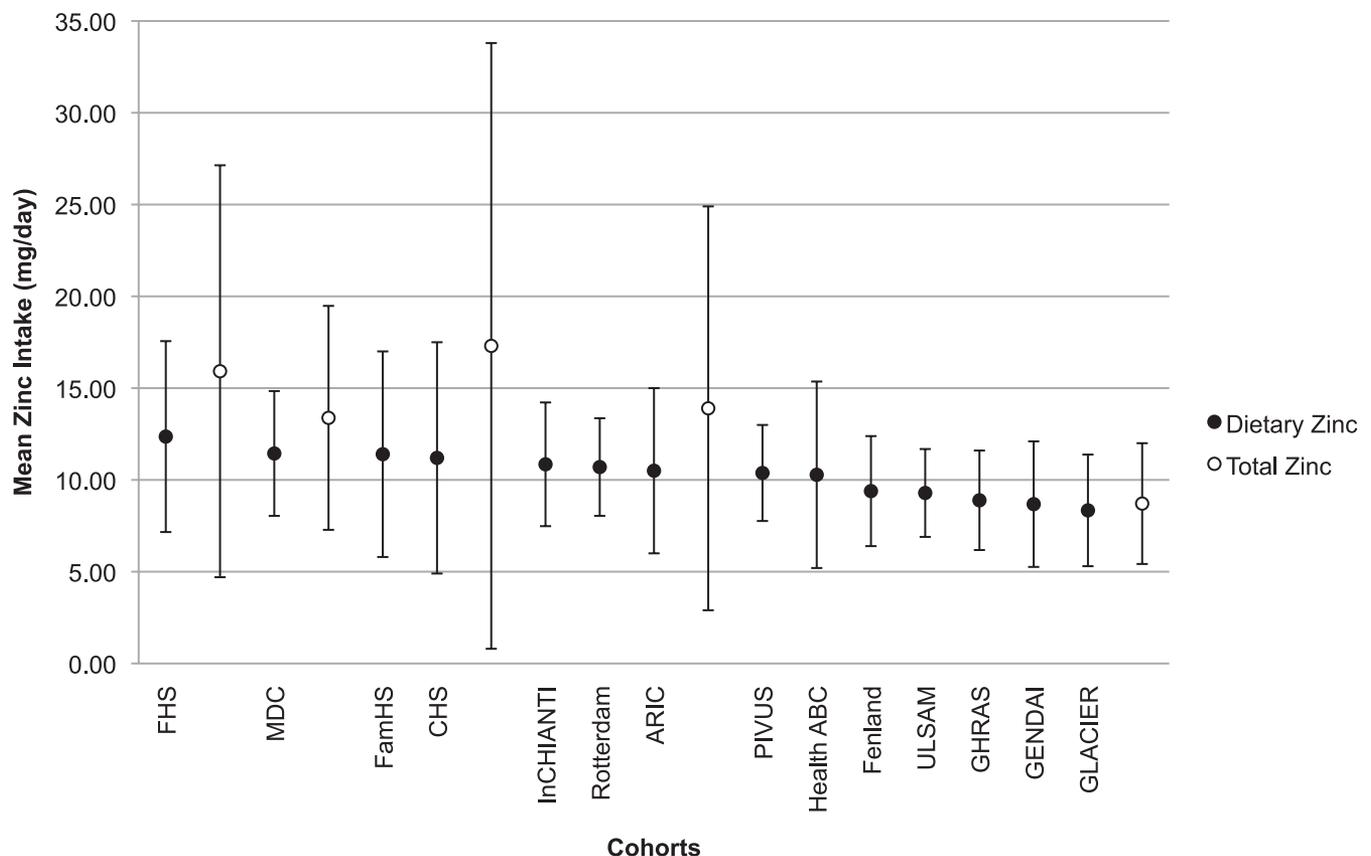


FIG. 1. Mean dietary and total zinc intake across cohorts. Values are presented as means (SD) and expressed as milligrams per day of dietary (food sources) and total (food sources and supplements) zinc intake. ●, dietary zinc; ○, total zinc.

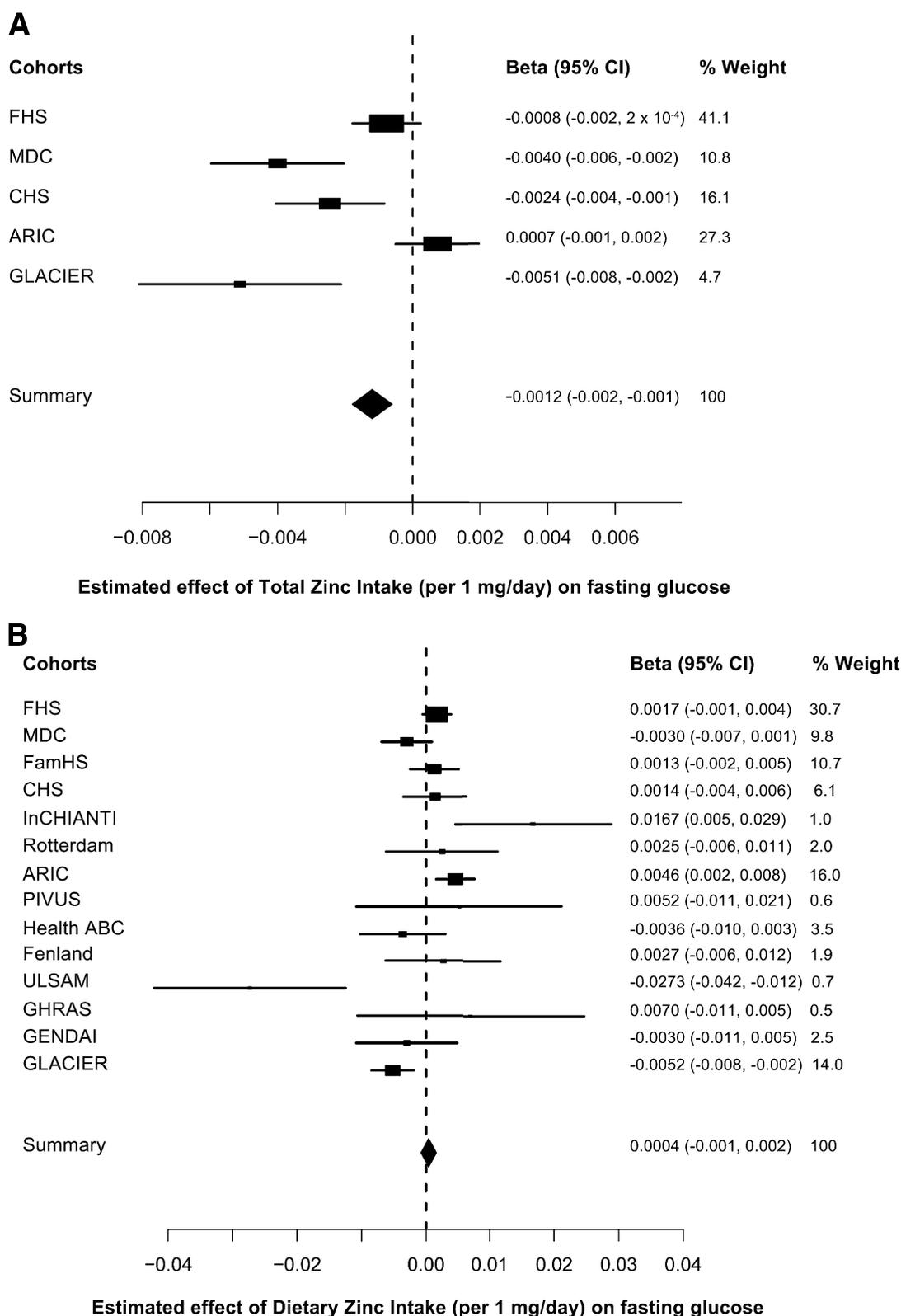


FIG. 2. Forest plots of the effect associations of zinc intake with fasting glucose. **A:** The effect association of total zinc (food sources and supplements) intake (per 1 mg/day) (summary P value = 0.0003, n = 34,533). **B:** Effect association of dietary zinc (food sources) intake (per 1 mg/day) (summary P value = 0.52, N = 46,021). Fasting blood glucose is expressed in micromoles per liter, and all associations are adjusted for age and sex, field center (in the CHS, the InCHIANTI study, the ARIC study, and the Health ABC study), and family structure by principal components (in the FHS and the FamHS).

TABLE 2
Total zinc intake and the genetic variants meta-analyzed interactions on fasting glucose levels*

| SNP | Chromosome | Nearest gene | Effect/other allele | Effect allele frequency | n | Cohorts | Interaction | | Interaction <i>P</i> | <i>F</i> ² (%) (95% CI) | <i>Q</i> test <i>P</i> |
|------------------------|------------|---------------------|---------------------|-------------------------|--------|---------|---------------------------------|--------------|----------------------|---------------------------------------|------------------------|
| | | | | | | | (total zinc × SNPs) [β (SE)] | [β (SE)] | | | |
| Glucose-related | | | | | | | | | | | |
| rs340874 | 1 | <i>PROXI</i> | C/T | 0.52 | 34,037 | 5 | -0.0012 (0.0005) | 0.03 | 0 (0-79) | 0.99 | |
| rs780094 | 2 | <i>GCKR</i> | C/T | 0.58 | 34,307 | 5 | 0.0004 (0.0005) | 0.35 | 0 (0-79) | 0.97 | |
| rs560887 | 2 | <i>G6PC2</i> | C/T | 0.71 | 34,061 | 5 | 3 × 10 ⁻⁵ (0.0005) | 0.95 | 0 (0-79) | 0.84 | |
| rs11708067 | 3 | <i>ADCY5</i> | A/G | 0.76 | 34,111 | 5 | -0.0011 (0.0006) | 0.05 | 25 (0-70) | 0.25 | |
| rs11920090 | 3 | <i>SLC2A2</i> | T/A | 0.83 | 34,033 | 5 | 0.0003 (0.0007) | 0.70 | 42 (0-79) | 0.14 | |
| rs2191349 | 7 | <i>DGKB/TMEM195</i> | T/G | 0.54 | 34,108 | 5 | -0.0002 (0.0005) | 0.71 | 0 (0-79) | 0.47 | |
| rs4607517 | 7 | <i>GCK</i> | A/G | 0.21 | 34,333 | 5 | -0.0003 (0.0006) | 0.58 | 34 (0-75) | 0.19 | |
| rs11558471 | 8 | <i>SLC30A8</i> | A/G | 0.70 | 34,150 | 5 | -0.0017 (0.0006) | 0.005 | 0 (0-79) | 0.78 | |
| rs7034200 | 9 | <i>GLIS3</i> | A/C | 0.49 | 33,874 | 5 | 2 × 10 ⁻⁵ (0.0005) | 0.97 | 0 (0-79) | 0.76 | |
| rs10885122 | 10 | <i>ADRA2A</i> | G/T | 0.85 | 33,933 | 5 | 0.0002 (0.0008) | 0.75 | 25 (0-70) | 0.25 | |
| rs4506565 | 10 | <i>TCF7L2</i> | T/A | 0.32 | 18,235 | 4 | 0.0002 (0.0005) | 0.71 | 57 (0-86) | 0.07 | |
| rs11605924 | 11 | <i>CRY2</i> | A/C | 0.47 | 34,137 | 5 | 0.0007 (0.0005) | 0.16 | 30 (0-73) | 0.22 | |
| rs7944584 | 11 | <i>MADD</i> | A/T | 0.69 | 33,921 | 5 | -0.0004 (0.0006) | 0.44 | 8 (0-81) | 0.36 | |
| rs174450 | 11 | <i>FADS1</i> | T/C | 0.66 | 34,172 | 5 | -0.0005 (0.0005) | 0.33 | 0 (0-79) | 0.82 | |
| rs10830963 | 11 | <i>MTNR1B</i> | G/C | 0.28 | 34,009 | 5 | 0.0001 (0.0006) | 0.91 | 48 (0-81) | 0.11 | |
| rs11071657 | 15 | <i>C2CD4B</i> | A/G | 0.61 | 33,962 | 5 | 0.0003 (0.0005) | 0.55 | 9 (0-81) | 0.36 | |
| Insulin-related | | | | | | | | | | | |
| rs4675095 | 2 | <i>IRS1</i> | T/A | 0.06 | 18,241 | 4 | 0.0003 (0.0008) | 0.74 | 0 (0-85) | 0.76 | |
| rs35767 | 12 | <i>IGF1</i> | G/A | 0.81 | 33,888 | 4 | 0.0001 (0.0007) | 0.92 | 0 (0-79) | 0.82 | |
| Zinc-related | | | | | | | | | | | |
| rs10493846 | 1 | <i>SEC63D1</i> | G/T | 0.74 | 18,158 | 4 | -0.0005 (0.0006) | 0.37 | 0 (0-79) | 0.61 | |
| rs11167682 | 5 | <i>SAP90 L</i> | G/T | 0.76 | 18,249 | 5 | 0.0001 (0.0006) | 0.91 | 0 (0-85) | 0.91 | |

*I*², Higgins heterogeneity index; *Q* test, Cochran heterogeneity test. *Estimates for the interaction between total zinc intake (per 1 mg/day from food sources and supplements) and SNPs (per effect allele) on fasting glucose levels (millimoles per liter), adjusted for age and sex, field center (in the CHS and the ARIC study), and family structure by principal components (in the FHS). Boldface values represent *P* values significant at the conventional level of 0.05.

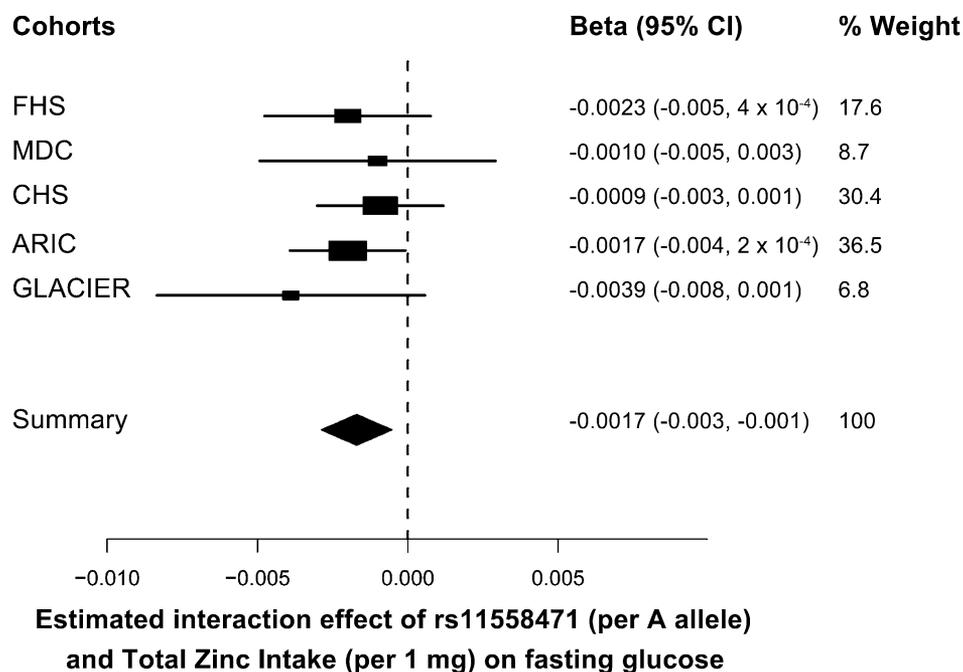


FIG. 3. Forest plot of the interaction between *SLC30A8* rs11558471 and total zinc intake on fasting glucose. The rs11558471 effect is expressed per A allele, total zinc (food sources and supplements) intake per 1 mg/day, and fasting blood glucose in micromoles per liter. Associations are adjusted for age and sex, field center (in the CHS and the ARIC study), and family structure by principal components (in the FHS) (summary P value = 0.0051, N = 34,150).

Association of zinc intake and genetic variants with fasting glucose levels. We observed a significant association of total zinc intake with lower fasting glucose levels (summary P value = 0.0003), with an estimated 0.0012 mmol/L lower fasting glucose concentration per 1 mg greater daily total zinc intake (Fig. 2A). When we additionally adjusted for BMI levels, the magnitude of the association was slightly attenuated (β -coefficient \pm SE: -0.0009 ± 0.0003 mmol/L fasting glucose per 1 mg/day greater total zinc intake, summary P value = 0.0037). In contrast, we did not observe a significant association of dietary zinc intake with fasting glucose levels (Fig. 2B); however, we observed that the magnitude and direction of cohort-specific associations varied notably. After additional adjustments for BMI levels, the association between dietary zinc intake and fasting glucose levels remained nonsignificant (β -coefficient \pm SE: -0.0005 ± 0.0006 , summary P value = 0.43).

In the meta-analysis, 16 of 20 SNPs were significantly associated with fasting glucose levels (Supplementary Table 2). This was consistent with previous published data from the MAGIC (2) that included, as a subsample, roughly one-half of all participants included in the current meta-analysis (Supplementary Table 1).

Interaction of zinc intake and genetic variants on associations with fasting glucose levels. We investigated the interactions between the genetic variants and total zinc intake (available in five cohorts; Supplementary Table 3), and the strongest interaction we observed was for rs11558471 in *SLC30A8* (β -coefficient \pm SE per A allele of 1 mg/day greater total zinc intake: -0.0017 ± 0.0006 ; summary interaction P value = 0.005) (Table 2 and Fig. 3). This interaction coefficient indicates that the glucose-raising effect of the risk allele (A) of rs11558471 diminishes as total zinc intake increases (per 1 mg/day), with the strongest inverse association between glucose and total zinc intake seen in individuals carrying both copies

of the risk allele. When we applied additional adjustment for BMI, the interaction between rs11558471 and total zinc intake with fasting glucose was slightly attenuated (β -coefficient \pm SE: -0.0014 ± 0.0006 ; summary interaction P value = 0.01).

Because the interaction effects reported above may be influenced by the total amount and source of zinc (i.e., dietary or supplemental), we proceeded by stratifying the interaction analyses by zinc supplement use (i.e., zinc supplement users and nonusers) (Fig. 4). As hypothesized, the magnitude of the interaction between total zinc intake and *SLC30A8* variant was weaker among zinc-supplement nonusers (β -coefficient \pm SE: -0.0006 ± 0.0014 ; summary P value = 0.65) compared with zinc-supplement users (β -coefficient \pm SE: -0.0024 ± 0.0009 ; summary P value = 0.008).

Few of the additional tests of gene-times-zinc interactions yielded noteworthy results (Tables 2 and 3). Of these, the most statistically significant tests of interaction were observed for the rs340874 *PROX1* (for total zinc intake) (Table 2), rs2191349 *DGKB/TMEM195*, and rs174550 *FADS1* variants (for dietary zinc intake) (Table 3 and Supplementary Fig. 2; cohort-specific estimates shown in Supplementary Table 4), although none of these P values withstood correction for multiple testing.

DISCUSSION

Given the important role of zinc in β -cell function and insulin homeostasis (12), we hypothesized that zinc intake might modify the effects of previously discovered glucose-raising genetic loci (2), many of which are thought to influence β -cell function. To test this hypothesis, we conducted a large-scale meta-analysis of up to 14 cohorts in which we assessed the interaction between zinc intake and glucose-associated genetic variants on fasting glucose levels among individuals without T2D.

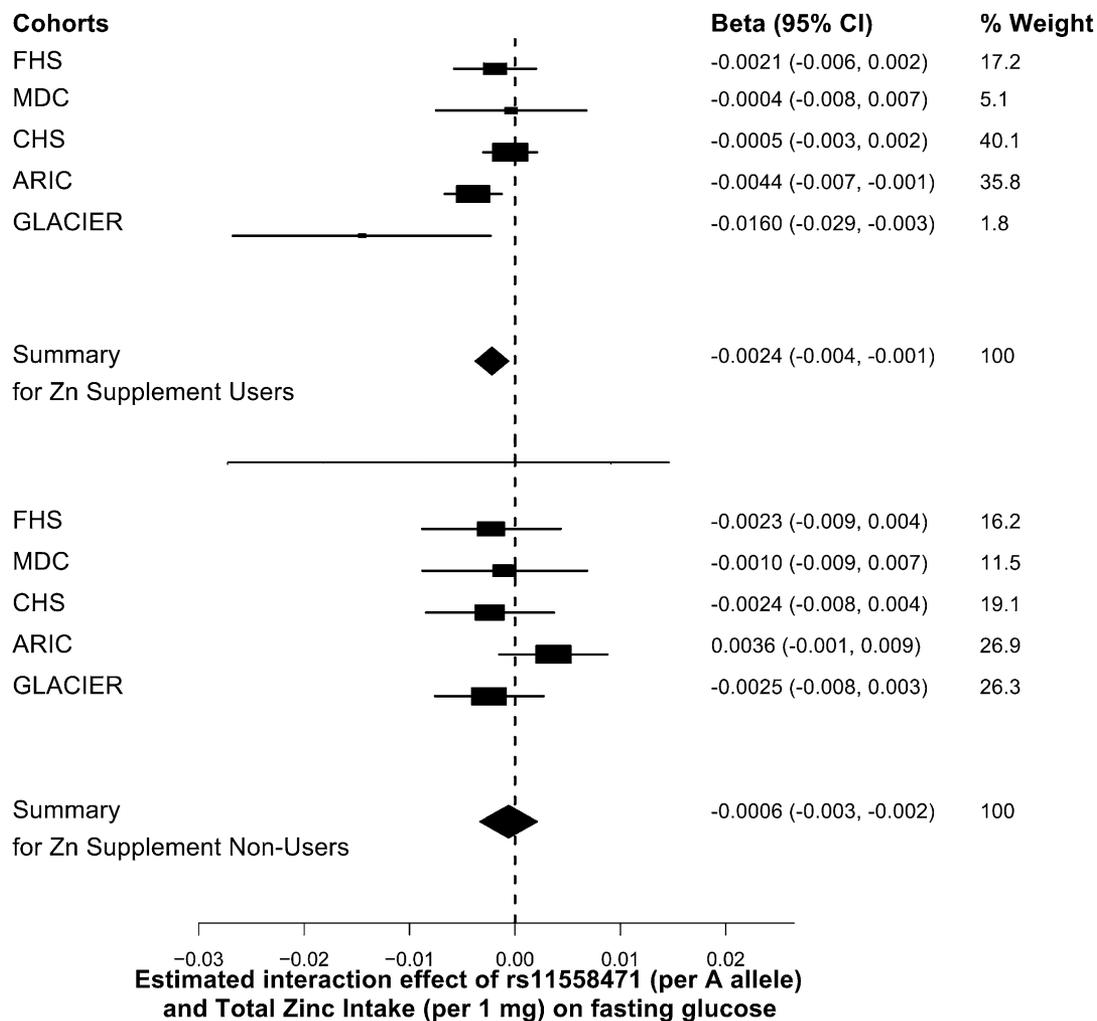


FIG. 4. Forest plots of the *SLC30A8* rs11558471 interaction with zinc intake on glucose, by zinc supplement use. The rs11558471 effect is expressed per A allele, zinc intake per 1 mg/day, and fasting blood glucose in micromoles per liter. Associations are adjusted for age and sex, field center (in the CHS and the ARIC study), and family structure by principal components (in the FHS) (summary P value = 0.0077, N = 4,986 for zinc supplement users and summary P value = 0.65, n = 29,164 for zinc supplement nonusers).

The strongest interaction effect was detected for the *SLC30A8* rs11558471 variant and total zinc intake (food sources and supplements) on fasting glucose levels. We estimated that, for individuals with the A/G genotype (i.e., one glucose-raising allele), an average daily total zinc intake of 14 mg (observed in our study) is associated with a 0.024 mmol/L lower glucose concentration than seen in individuals with the G/G genotype (i.e., two non-glucose-raising alleles). Concordantly, in individuals with both copies of the risk allele (i.e., A/A genotype), the magnitude of this association was doubled (i.e., 0.048 mmol/L). An average daily total zinc intake of 14 mg could be achieved by a nutrient supplement (~10 mg zinc) plus an average serving of red meat or fish/seafood (~3 ounces) or three average servings of dairy products (~2 cups of milk per yogurt and 3 ounces of cheese). Further investigation of this interaction effect showed that it was more evident among zinc supplement users than among nonusers. We also observed a significant association of total zinc intake, but not dietary zinc intake alone, with lower fasting glucose levels. None of the other investigated interactions between zinc intake (dietary or total) and variants were statistically significant after multiple-testing adjustment.

Current findings in the context of gene-environment interaction investigations. The *SLC30A8* gene encodes the newly characterized ZnT8 zinc transporter (27). Zinc homeostasis depends on two families of transporters: the ZIP (*SLC39*) family regulating cellular zinc influx and the ZnT (*SLC30*) family regulating efflux (28). It has been shown that the ZnT8 transporter is primarily expressed in β -cells and colocalizes with insulin-containing secretory granules (29). Alterations in ZnT8 expression strongly modulate insulin secretion (29,30). Moreover, it has recently been shown that ZnT8 β -cell-specific knockout (Znt8KO) mice are glucose intolerant have reduced β -cell zinc accumulation, have atypical insulin granules, have reduced first-phase glucose-stimulated insulin secretion, have reduced insulin-processing enzyme transcripts, and have increased proinsulin levels (14). Furthermore, a genetic variant (rs13266634) in *SLC30A8* has been reliably associated with fasting glucose levels and T2D risk in several GWAS (31–34). Of interest, observations suggest that this variant impairs islet ZnT8 expression, insulin secretion, or glucose homeostasis (35–37). In addition, this variant is associated with the production of a less active zinc transporter protein, suggesting less efficiency of zinc

TABLE 3
Dietary zinc intake and the genetic variants meta-analyzed interactions on fasting glucose levels*

| SNP | Chromosome | Nearest gene | Effect/other allele | Effect allele frequency | n | Cohorts | Interaction (total zinc × SNPs) [β (SE)] | Interaction P | I ² (%) (95% CI) | Q test P |
|------------------------|------------|---------------------|---------------------|-------------------------|--------|---------|--|---------------|-----------------------------|----------|
| Glucose-related | | | | | | | | | | |
| rs340874 | 1 | <i>PROX1</i> | C/T | 0.52 | 45,525 | 14 | -0.0007 (0.0009) | 0.42 | 0 (0-55) | 0.93 |
| rs780094 | 2 | <i>GCKR</i> | C/T | 0.60 | 45,795 | 14 | -0.0006 (0.0009) | 0.52 | 19 (0-56) | 0.25 |
| rs560887 | 2 | <i>G6PC2</i> | C/T | 0.70 | 45,549 | 14 | 0.0001 (0.001) | 0.90 | 0 (0-55) | 0.79 |
| rs11708067 | 3 | <i>ADCY5</i> | A/G | 0.70 | 45,599 | 14 | -0.0005 (0.0011) | 0.65 | 0 (0-55) | 0.45 |
| rs11920090 | 3 | <i>SLC2A2</i> | T/A | 0.76 | 45,521 | 14 | -0.0004 (0.0013) | 0.78 | 30 (0-63) | 0.13 |
| rs2191349 | 7 | <i>DGKB/TMEM195</i> | T/G | 0.54 | 45,596 | 14 | -0.0021 (0.0008) | 0.01 | 17 (0-55) | 0.27 |
| rs4607517 | 7 | <i>GCK</i> | A/G | 0.25 | 45,821 | 14 | 0.0005 (0.0011) | 0.68 | 23 (0-59) | 0.21 |
| rs11558471 | 8 | <i>SLC30A8</i> | A/G | 0.70 | 41,994 | 10 | -0.0009 (0.0011) | 0.39 | 0 (0-62) | 0.58 |
| rs7034200 | 9 | <i>GLIS3</i> | A/C | 0.48 | 45,362 | 14 | 0.0003 (0.0009) | 0.68 | 0 (0-55) | 0.59 |
| rs10885122 | 10 | <i>ADRA2A</i> | G/T | 0.81 | 45,421 | 14 | -0.0016 (0.0012) | 0.17 | 14 (0-52) | 0.31 |
| rs4506565 | 10 | <i>TCF7L2</i> | T/A | 0.32 | 27,010 | 10 | -0.0004 (0.0010) | 0.68 | 0 (0-62) | 0.73 |
| rs11605924 | 11 | <i>CRY2</i> | A/C | 0.47 | 45,625 | 14 | 0.0002 (0.0009) | 0.78 | 49 (6-73) | 0.019 |
| rs7944584 | 11 | <i>MADD</i> | A/T | 0.65 | 45,409 | 14 | 0.0008 (0.0010) | 0.44 | 24 (0-60) | 0.20 |
| rs174550 | 11 | <i>FADS1</i> | T/C | 0.63 | 45,660 | 14 | -0.0019 (0.0009) | 0.04 | 26 (0-61) | 0.18 |
| rs10830963 | 11 | <i>MTNR1B</i> | G/C | 0.28 | 45,497 | 14 | 0.0008 (0.0010) | 0.47 | 34 (0-65) | 0.10 |
| rs11071657 | 15 | <i>C2CD4B</i> | A/G | 0.59 | 45,450 | 14 | -0.0002 (0.0009) | 0.86 | 0 (0-55) | 0.83 |
| Insulin-related | | | | | | | | | | |
| rs4675095 | 2 | <i>IRS1</i> | T/A | 0.07 | 29,729 | 13 | -0.0006 (0.0019) | 0.76 | 42 (0-70) | 0.06 |
| rs35767 | 12 | <i>IGF1</i> | G/A | 0.74 | 45,376 | 14 | -0.0003 (0.0013) | 0.83 | 0 (0-55) | 0.47 |
| Zinc-related | | | | | | | | | | |
| rs10493846 | 1 | <i>SEC63D1</i> | G/T | 0.74 | 29,646 | 13 | -0.0008 (0.0011) | 0.49 | 21 (0-59) | 0.23 |
| rs11167682 | 5 | <i>SAP30 L</i> | G/T | 0.76 | 29,737 | 13 | -0.0003 (0.0011) | 0.81 | 0 (0-57) | 0.50 |

I², Higgins heterogeneity index; Q test, Cochran heterogeneity test. *Estimates for the interaction between dietary zinc intake (per 1 mg/day from food sources) and SNPs (per effect allele) on fasting glucose levels (millimoles per liter), adjusted for age and sex, field center (in the CHS, the InCHIANTI study, the ARIC study, and the Health ABC study), and family structure by principal components (in the FHS and the FamHS). Boldface values represent P values significant at the conventional level of 0.05.

accumulation and insulin crystallization (38). It recently has been demonstrated that the same variant does not affect insulin secretion from human islets as well as islet expressions of *SLC30A8* (39). It is noteworthy that rs13266634 is in strong linkage disequilibrium with rs11558471 ($r^2 = 0.96$), included in the current study. Additional studies on these topics are important to provide evidence that variation at *SLC30A8* influences the regulation of zinc transporter activity or the modulation of islet zinc content; this would support the biological plausibility of the statistical interaction we report here.

Current findings in the context of dietary epidemiological investigations. Our meta-analysis revealed a significant inverse association between total zinc intake and fasting glucose levels. Of interest, we did not observe this association when only zinc derived from food sources was considered. This observation may be attributable to a number of factors, including differences in the bioavailability of dietary zinc compared with zinc supplements (40), a threshold effect of zinc intake on fasting glucose levels, or because when dietary and supplemental zinc are combined the wider trait variance affords greater statistical power to detect associations and interactions. The latter also is based on the power calculations we performed (RESEARCH DESIGN AND METHODS). Moreover, the different dietary assessment tools used to assess dietary zinc intake across the participating cohorts may have varied in precision to a greater extent than when used to assess supplemental zinc intake. However, potential confounding lifestyle characteristics associated with supplement use were not considered in the current study.

Our findings are in line with results from a large 24-year prospective study among women, demonstrating a significant association between total zinc intake and lower risk of T2D (17). In the same study, but in contrast to our results, dietary zinc intake also was associated with a lower risk of T2D after multiple adjustments for dietary and nondietary factors (17). In a cross-sectional study of an Asian-Indian population (16), dietary zinc intake was associated with a lower prevalence of T2D and the metabolic syndrome. However, another study of Chinese adults did not observe a relationship between dietary zinc intake and risk of hyperglycemia (15). However, it is worth mentioning that in the current study we focused on fasting glucose levels in subjects without T2D, as opposed to the prevalence of T2D in the referenced studies.

Strengths and limitations. The strengths of the current work include the large samples, the availability of standardized exposure and outcome data, and the use of a uniform analytical plan prior to meta-analysis. The advantages of meta-analyses like ours over individual studies or literature-based meta-analysis include improved power to detect interaction effects and the minimization of recall bias and publication bias (41,42).

We selected the majority of the genetic variants included in our meta-analysis on the basis of the results of GWAS for genetic main effects on insulin or glucose traits (2). This approach is convenient, and one can be confident that the genetic variants are reliably associated with the traits of interest, which is rarely the case with conventional biologic candidate-gene studies. Nevertheless, it may be that the variants that reach genome-wide significance in main-effect GWAS are less likely to interact with environmental factors because such interactions tend to weaken the

statistical significance of the main effects (43). Although the effect estimates observed here are indeed small in magnitude, it is important to bear in mind that these effects are likely to be underestimated, given that the observed genetic loci may not be causal and because the methods of determining usual zinc intake are imprecise. Our approach does not preclude that other genetic variants not previously known to be associated with glycemic traits might interact with zinc intake.

As with all epidemiological studies assessing dietary intake, systematic measurement error in the diet exposure could have biased our results. However, there currently is no satisfactory biomarker for the assessment of dietary zinc intake, and even though plasma zinc concentration seems to be a reliable marker of zinc status, it has limited sensitivity in the normal range of zinc intake (44,45). The issues of dietary assessment validity versus cost-effectiveness in population-based studies have been widely discussed elsewhere (46,47), but it is unlikely that a glucose-associated genotype is associated with a tendency to misreport dietary intake. Furthermore, we did not assess interactions of zinc with other nutrients; such interactions might influence the bioavailability of zinc (40). Systematic measurement error in the estimation of dietary exposures in gene-environment interaction studies could bias the estimation of the dietary main effect and lead to underestimates of the interaction between dietary factors and genetic variants, potentially raising type II error rates (48).

We conducted a large-scale, gene-diet interaction meta-analysis in which we investigated the association between zinc intake (dietary and total) and fasting glucose levels and the interactions between zinc intake and glucose-, insulin-, and zinc-related genetic variants on fasting glucose levels. We showed that total zinc intake (food sources and supplements), but not zinc from foods alone, is associated with lower fasting glucose levels in individuals without T2D. Our findings suggest that total zinc intake has a stronger inverse association with fasting glucose levels in individuals carrying the glucose-raising A allele of rs11558471 *SLC30A8* (a β -cell zinc transporter), compared with individuals carrying the G allele. The current study indicates that gene-environment interaction analyses can help elucidate our understanding of the biological pathways involved in micronutrient influences on systemic glucose homeostasis.

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